

Scoring Reading Parameters: An Inter-Rater Reliability Study Using The MNREAD Chart

Karthikeyan Baskaran, Antonio Filipe Macedo, Yingchen He, Laura

Hernandez-Moreno, Tatiana Queirós, J Stephen Mansfield, Aurélie Calabrèse

▶ To cite this version:

Karthikeyan Baskaran, Antonio Filipe Macedo, Yingchen He, Laura Hernandez-Moreno, Tatiana Queirós, et al.. Scoring Reading Parameters: An Inter-Rater Reliability Study Using The MNREAD Chart. 2019. hal-02079375

HAL Id: hal-02079375 https://hal.science/hal-02079375

Preprint submitted on 26 Mar 2019

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PLOS ONE Scoring Reading Parameters: An Inter-Rater Reliability Study Using The MNREAD Chart --Manuscript Draft--

Manuscript Number:	
Article Type:	Research Article
Full Title:	Scoring Reading Parameters: An Inter-Rater Reliability Study Using The MNREAD Chart
Short Title:	Inter-Rater Reliability Of The MNREAD Acuity Chart
Corresponding Author:	Aurélie Calabrèse Aix-Marseille Universite Marseille, FRANCE
Keywords:	low vision; reading performance; reading test; MNREAD acuity chart; inter-rater reliability; computer-based scoring algorithms
Abstract:	Purpose: First, to evaluate inter-rater reliability when human raters estimate the reading performance of visually impaired individuals using the MNREAD acuity chart. Second, to evaluate the agreement between computer-based scoring algorithms and compare them with human rating. Methods: Reading performance was measured for 101 individuals with low vision, using the Portuguese version of MNREAD. Seven raters estimated the maximum reading speed (MRS) and critical print size (CPS) of each individual MNREAD curve. MRS and CPS were also calculated automatically for each MNREAD curve using two different algorithms: the original standard deviation method (SDev) and a non-linear mixed effects (NLME) modeling. Intra-class correlation coefficients (ICC) were used to estimate absolute agreement between raters and/or algorithms. Results: Absolute agreement between raters was excellent for MRS (ICC = 0.97; 95%CI [0.96, 0.98]) and good for CPS (ICC = 0.77; 95%CI [0.69, 0.83]). For CPS interrater reliability was poorer among less experienced raters (ICC = 0.70; 95%CI [0.57, 0.80]) compared to experienced ones (ICC = 0.82; 95%CI [0.57, 0.80]). Absolute agreement between the two algorithms was excellent for MRS (ICC = 0.96; 95%CI [0.91, 0.98]). For CPS, the best possible agreement was good and for CPS defined as the print size sustaining 80% of MRS (ICC = 0.77; 95%CI [0.68, 0.84]). Conclusion: For MRS, inter-rater reliability is excellent, even considering the possibility of noisy and/or incomplete data collected in low-vision individuals. For CPS, inter-rater reliability is lower, which may be problematic, for instance in the context of multicenter studies or follow-up examinations. Setting up consensual guidelines to deal with ambiguous datasets may help improve reliability. While the exact definition of CPS should be chosen on a case-by-case basis depending on the clinician or researcher's motivations, evidence suggests that estimating CPS as the smallest print size sustaining about 80% of MRS would increase inter-rater rel
Order of Authors:	Karthikeyan Baskaran
	Antonio Filipe Macedo
	Yingchen He
	Laura Hernandez-Moreno
	Tatiana Queirós
	J. Stephen Mansfield
	Aurélie Calabrèse
Opposed Reviewers:	
Opposed Reviewers: Additional Information:	
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PTDC/DPT-EPI/0412/2012 in the context of the Prevalence and Costs of Visual Impairment in Portugal: a hospital based study (PCVIP-study). The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript.

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the study involved:	protection authority with the reference 9936/2013 and received approval number 5982/2014.
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Aurelie Calabrese, PhD

Post-doctoral Associate

Aix-Marseille University - Centre St Charles - Pôle 3C 3 place Victor Hugo - case D - 13331 Marseille cedex 3

aurelie.calabrese@univ-amu.fr

Date: January 31st 2019

Dear members of the Editorial Board,

Please find enclosed our manuscript entitled: "Scoring Reading Parameters: An Inter-Rater Reliability Study Using The MNREAD Chart", by Karthikeyan Baskaran, Antonio Filipe Macedo, Yingchen He, Laura Hernandez-Moreno, Tatiana Queirós, J. Stephen Mansfield, and Aurélie Calabrèse, which we would like to submit for publication as an research article in PLoS ONE.

The primary goal of this work is to evaluate inter-rater reliability when human raters estimate reading performance using the MNREAD acuity chart. Our motivation for this study was the lack of evidence that different extraction methods used by different raters would lead to comparable estimates of reading performance, which is especially relevant in the context of multicenter studies, or when looking at follow-up data. Our results demonstrate excellent interrater reliability for the Maximum Reading Speed (i.e. the fastest that one can read when print size is not limiting) and good inter-rater reliability for the Critical Print Size (i.e. the print size for which reading speed is maximum). Our work also provides further tips and instructions on how to score noisy and/or incomplete MNREAD data. These tips may serve as a starting point to help clinicians and researchers reduce variability.

We confirm that this manuscript has not been published elsewhere and is not under consideration by another journal. All Authors have approved the manuscript and agree with submission to PLoS ONE.

Authors report no conflict of interest, except for JSM who receives royalties from the sales of MNREAD Acuity Charts.

We appreciate your consideration of publication of this paper.

Sincerely, Aurélie Calabrèse, PhD







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1	Full title: Scoring Reading Parameters: An Inter-Rater Reliability Study
2	Using The MNREAD Chart.
3	
4	Short title: Inter-Rater Reliability Of The MNREAD Acuity Chart.
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6	Authors: Karthikeyan Baskaran ¹ , Antonio Filipe Macedo ^{1, 2} , Yingchen He ³ , Laura Hernandez-
7	Moreno ² , Tatiana Queirós ⁴ , J. Stephen Mansfield ⁵ , Aurélie Calabrèse ^{6,7}
8	
9	Affiliations:
10	¹ Department of Medicine and Optometry, Linnaeus University, Kalmar, Sweden
11	² Low Vision and Visual Rehabilitation Lab, Department and Center of Physics—Optometry and
12	Vision Science, University of Minho Braga, Braga, Portugal
13	³ Department of Ophthalmology & Visual Neurosciences, University of Minnesota, Twin Cities,
14	United States
15	⁴ Serviço de Oftalmologia, Hospital de Braga, Braga, Portugal.
16	⁵ Department of Psychology, SUNY College at Plattsburgh, Plattsburgh, New York, United
17	States
18	⁶ Aix-Marseille University, Marseille, France
19	⁷ Laboratoire de Psychologie Cognitive, CNRS, Marseille, France

21	Corresponding author:
22	Aurélie Calabrèse (AC)
23	aurelie.calabrese@univ-amu.fr
24	3 place Victor Hugo - 13003 Marseille, France
25	
26	Word count:
27	Abstract = 300 words
28	Text = 3740 words
29	Intro = 1025; Methods = 1010; Results = 580; Discussion = 1011; Conclusion = 114
30	References = 30; Figures = 5; Tables = 4
31	
32	Funding:
33	This work was partially supported by FCT (COMPETE/QREN).
34	
35	Declarations of interest:
36	Author JSM receives royalties from the sales of MNREAD Acuity Charts.
37	

38 Abstract

Purpose: First, to evaluate inter-rater reliability when human raters estimate the reading
performance of visually impaired individuals using the MNREAD acuity chart. Second, to
evaluate the agreement between computer-based scoring algorithms and compare them with
human rating.

Methods: Reading performance was measured for 101 individuals with low vision, using the Portuguese version of MNREAD. Seven raters estimated the maximum reading speed (MRS) and critical print size (CPS) of each individual MNREAD curve. MRS and CPS were also calculated automatically for each MNREAD curve using two different algorithms: the original standard deviation method (SDev) and a non-linear mixed effects (NLME) modeling. Intra-class correlation coefficients (ICC) were used to estimate absolute agreement between raters and/or algorithms.

Results: Absolute agreement between raters was excellent for MRS (ICC = 0.97; 95%CI [0.96, 0.98]) and good for CPS (ICC = 0.77; 95%CI [0.69, 0.83]). For CPS inter-rater reliability was poorer among less experienced raters (ICC = 0.70; 95%CI [0.57, 0.80]) compared to experienced ones (ICC = 0.82; 95%CI [0.57, 0.80]). Absolute agreement between the two algorithms was excellent for MRS (ICC = 0.96; 95%CI [0.91, 0.98]). For CPS, the best possible agreement was good and for CPS defined as the print size sustaining 80% of MRS (ICC = 0.77; 95%CI [0.68, 0.84]).

57 Conclusion: For MRS, inter-rater reliability is excellent, even considering the possibility of noisy
58 and/or incomplete data collected in low-vision individuals. For CPS, inter-rater reliability is
59 lower, which may be problematic, for instance in the context of multicenter studies or follow-up

60	examinations. Setting up consensual guidelines to deal with ambiguous datasets may help
61	improve reliability. While the exact definition of CPS should be chosen on a case-by-case basis
62	depending on the clinician or researcher's motivations, evidence suggests that estimating CPS as
63	the smallest print size sustaining about 80% of MRS would increase inter-rater reliability.

64 Introduction

65 Reading difficulty is a major concern for patients referred to low-vision centers [1]. Therefore, 66 most Quality-of-Life questionnaires assessing the severity of vision disability contain one or 67 more items on subjective reading difficulty [2-5]. However, substantial discrepancy has been 68 observed between self-reported reading difficulty and measured reading speed [6]. For this 69 reason, reading performance should be evaluated objectively to serve as a reliable outcome 70 measure in clinical trials, multisite investigations or longitudinal studies. To assess, for instance, 71 the success of vision rehabilitation techniques, surgical procedures or ophthalmic treatments, 72 measures of reading ability should be obtained using standardized tests with demonstrated high 73 repeatability.

74 Among the standardized tests available, the MNREAD acuity chart can be used to evaluate 75 reading performance for people with normal vision or low vision in clinical and research 76 environments [7]. In brief, the MNREAD chart measures four parameters that characterize how 77 reading performance changes when print size decreases: the maximum reading speed (MRS), the 78 critical print size (CPS), the reading acuity (RA) and the reading accessibility index (ACC) [8]. 79 The reading acuity and reading accessibility index are clearly defined by the number of reading 80 errors made at small print sizes and the reading speeds for a range of larger sizes. In the original 81 MNREAD manual, provided with the chart, MRS and CPS are defined as follows: "The critical 82 print size is the smallest print size at which patients can read with their maximum reading speed. 83 [...] Typically, reading time remains fairly constant for large print sizes. But as the acuity limit is 84 approached there comes a print size where reading starts to slow down. This is the critical print 85 size. The maximum reading speed with print larger than the critical print size is the maximum 86 reading speed (MRS)." In short, values for MRS and CPS depend on the location of the flexion 87 point in the curve of reading speed versus print size (Fig 1). In normally sighted individuals, for 88 whom the MNREAD curve usually exhibits a standard shape (Fig 1-A), the above definitions 89 may be sufficient to extract MRS and CPS confidently by inspecting the curve. However, they 90 can be difficult to determine, especially for readers with visual impairments, who may experience 91 visual field defects (e.g. ring scotoma; Fig 1-B) or the use of multiple fixation sites (i.e. PRL; Fig 92 1-C) [9]. In such cases, the noisy and/or incomplete dataset resulting from atypical visual 93 function may be inconsistent with the assumption that people will read at a fairly constant speed 94 until font size compromises their ability to identify words and MNREAD curves may take an 95 unusual shape (Fig 1-D). If so, subjective decisions (e.g. ignoring outliers) must be made by the individual analysing the data (referred to as the "rater" in the present work, as opposed to the 96 97 "experimenter" who recorded the data). For this reason, MRS and CPS estimates may be 98 considered highly sensitive to inter-rater variability.

99

100 Fig 1: MNREAD curve examples.

101

102 In an attempt to reduce variability and unify the process of curve information extraction, 103 alternative scoring methods have been proposed. According to these "simpler" scoring rules, 104 MRS equals either the single largest reading speed [10] or the mean of the three largest reading 105 speeds [11]. Nonetheless, a criterion must be chosen for the CPS (smallest print size supporting 106 reading speed at either: 90% of MRS, 85%, 80%, etc.) but there is no general agreement on the 107 appropriate criterion to use. Overall, open discussions on how to score MNREAD parameters 108 optimally still persist in the literature [12]and the choice of scoring method constitutes an 109 additional factor contributing to inter-rater variability.

110 Another approach to reduce variability is to fit the MNREAD curve and estimate its parameters 111 using automated algorithms [13]. In the present work, we will focus on two of these methods. 112 The first one has been described by the MNREAD creators [14,15] and is used in the MNREAD 113 iPad app [16]. It is also the most widely used in the literature [11,17,18]. In short, it determines 114 the CPS as the smallest print size that supports reading speeds that are not significantly different 115 from the reader's maximum reading speed; we will refer to it as the standard deviation method 116 (SDev). The second method, especially recommended with large but incomplete datasets, 117 estimates the critical print size from smooth curve-fit to the MNREAD data using non-linear 118 mixed effects (NLME) modeling [19]; we will refer to it as the NLME method. Both methods are 119 described in the Methods section. Despite the advantage of these algorithms in operationalizing 120 the estimation of the MNREAD parameters, they present two major drawbacks: (1) they may not 121 be easily accessible in clinical environments, (2) they may fail to provide satisfactory measures 122 with noisy or small and incomplete datasets, necessitating further human inspection of the curves 123 for validation.

124 The Repeatability of the MNREAD chart measures has been assessed before in low vision 125 populations. Overall, studies have reported good intra and inter-session reliability [11,17,18,20], 126 as well as good repeatability across multiple testing sites and experimenters [21]. But to our 127 knowledge, variability of the MNREAD estimates scored by different raters from the same 128 dataset has not been evaluated. This question of inter-rater variability is especially relevant (1) in 129 the context of multicenter studies, where data are scored by different raters with different levels 130 of expertise, (2) when comparing results from different studies performed by different groups, or 131 (3) when looking at follow-up data involving different raters.

We have investigated the reliability of CPS and MRS estimates for MNREAD data collected
from participants with visual impairments. First, we evaluated the inter-rater reliability among
raters (Analysis 1). Second, we evaluate agreement between the NLME and SDev algorithms
(Analysis 2). Third, we evaluated agreement between raters and the two algorithms (Analysis 3).

136

137 Methods

138 **Participants**

139 Data from 101 participants with visual impairment were selected from a larger dataset, originally 140 collected to study the prevalence and costs of visual impairment in Portugal (PCVIP-study) 141 [22,23]. Only participants whose visual acuity in the better eye was 0.5 decimal (0.3 logMAR) or 142 worse and/or whose visual field was less than 20 degrees were selected for the present study. 143 Among them, only the participants who read at least five sentences on the MNREAD chart with 144 their "presenting reading glasses" were included. The study protocol was reviewed by the ethics 145 committee for Life Sciences and Health of the University of Minho (REF: SECVS-084/2013) and 146 was conducted in accordance with the principles of the Declaration of Helsinki. Written informed 147 consent was obtained from all participants. The study was registered with the Portuguese data 148 protection authority with the reference 9936/2013 and received approval number 5982/2014.

149 MNREAD Data

150 Reading performance was measured for each participant using the Portuguese version of the

- 151 MNREAD acuity chart [24]. Reading distance was adjusted for each participant and chosen
- according to his/her near visual acuity. Participants were asked to read the chart aloud as fast and

153 accurately as possible, one sentence at a time, starting from the largest print size. For each 154 sentence, reading time and number of misread words were recorded and reported on a score sheet 155 by the experimenter. Data were then transferred into a digital file and further processed in R [25]. 156 For each individual test, a corresponding MNREAD curve was plotted using the mnreadR 157 package [26] to display log reading speed as a function of print size (see S1 Appendix for all 101 158 curves). Because the shape of the curve can influence visual estimation of the reading parameters, 159 reading speed was plotted using a logarithmic scale so that reading speed variability (which is 160 proportional to the overall measure of reading speed) was constant at all speeds [14].

161 Raters' visual scoring

162 Seven raters were recruited to estimate the MRS and CPS of each individual MNREAD curve. 163 Since inter-rater reliability may be influenced by raters' prior experience with the MNREAD 164 chart, we included raters with different levels of expertise in MNREAD parameters estimation. 165 Each rater gave a self-rated score of expertise (on a 5 point scale from 0 = 'no previous to 166 experience' to 4 = 'top expertise'), both before and after rating all the MNREAD curves, to 167 account for the amount of practice gained during the study. Each rater was provided with S1 168 Appendix, containing the 101 MNREAD curves to score. Raters were instructed to follow the 169 standard guidelines provided with the MNREAD chart instructions (see Introduction). However, 170 coming from patients with impaired vision, many of the curves had noisy or incomplete data, 171 which potentially made it difficult to estimate the MRS and CPS. In such cases, we provided 172 more detailed instructions to the raters. These detailed instructions are available in S2 Appendix.

173 Algorithms' automated scoring

MRS and CPS were also calculated automatically for each 101 datasets using two algorithmbased estimations: the 'standard deviation' method and non-linear mixed effects modeling. The

176 standard deviation method (SDev) uses the original algorithm described in [14] and [15] to 177 estimate the MNREAD parameters. This algorithm iterates over the data searching for an optimal 178 reading speed plateau, from which MRS and CPS will be derived. To be considered optimal, a 179 plateau must encompass a range of print sizes that supports reading speed at a significantly faster 180 rate (1.96 \times standard deviation) than the print sizes smaller or larger than the plateau range (Fig 181 2). MRS is estimated as the mean reading speed for print sizes included in the plateau and CPS is 182 defined as the smallest print size on the plateau. In most cases, several print-size ranges can 183 qualify as an optimal plateau and the algorithm chooses the one with the fastest average reading 184 speed. In the present work, the standard deviation method estimation was performed using the 185 curveParam RT () function from the mnreadR R package.

186

187 Fig 2: Example of the standard deviation algorithm calculation on a typical dataset.

188 On iteration 1 (dark blue), the algorithm selects the first two sentences as plateau 1 (1.3 and 1.2 logMAR) and 189 calculates a selection criterion for this plateau. Criterion plateau 1 = mean (reading speed plateau 1) - 1.96 x 190 standard deviation (reading speed plateau 1) = $60.5 - 1.96 \times 2.1 = 56.3$ wpm. The point adjacent to plateau 1 (1.1) 191 logMAR) was read at 60 wpm, which is faster than criterion plateau 1, indicating that this point belongs to the 192 optimal plateau. A second iteration is then launched (light blue) with plateau 2 now encompassing the first 193 three sentences and a new criterion calculation. Criterion $_{plateau 2} = 60.3 - 1.96 \times 1.5 = 57.3$ wpm. Among the 194 points adjacent to plateau 2, there is still a value higher than this criterion (59 wpm at 0.9 logMAR), so the 195 algorithm continues to iterate one sentence at a time, including 1.0 logMAR in plateau 3 and 0.9 logMAR in 196 plateau 4. The calculations stop with plateau 4, for which selection criterion is higher than any remaining 197 points (criterion plateau 4 = 44.7 wpm). MRS is estimated as 57.2 wpm and CPS as 0.9 logMAR.

199 The non-linear mixed effects (NLME) modeling method is particularly suited for incomplete 200 datasets from individuals with reading or visual impairment [19]. The NLME model uses 201 parameter estimates from a larger group (101 datasets here) to allow suitable curve fits for 202 individual datasets that contain few data points. In the present work, we used an NLME model 203 with a negative exponential decay function, as described in details in [19], where a single 204 estimate of MRS can yield several measures of CPS depending on the definition chosen (e.g. 205 print size required to achieve 90% of MRS, 80% of MRS, etc.). Therefore, five values of CPS 206 were estimated, i.e. 95%, 90%, 85%, 80% and 75% of MRS. NLME modeling and parameters 207 estimation were performed using the nlmeModel () and nlmeParam () functions from mnreadR.

208

209 Statistical Analysis

210 In all three analyses, intra-class correlation coefficient (ICC) was used to assess absolute 211 agreement between raters and/or algorithms [27]. This reliability index (ranging from 0 to 1; 1 212 meaning perfect agreement) is widely used in the literature in test-retest, intra-rater, and inter-213 rater reliability analyses [28]. In the present work, ICC values estimate the variation between two 214 or more methods (whether raters or algorithms) in scoring the same data by calculating the 215 absolute agreement between them. For each analysis, the appropriate ICC form (dependent on 216 research design and assumptions) was chosen by selecting the correct combination of "model", 217 "type" and "definition", as detailed in Table 1 [29]. ICC values were calculated using SPSS 218 statistical package and limits of agreement were visualized with Bland-Altman plots. Following 219 guidelines from [28], ICC values and their 95% confidence intervals (95% CI) were interpreted 220 as showing: "poor agreement" if less than 0.5; "moderate agreement" if comprised between 0.5

and 0.75; "good agreement" if comprised between 0.75 and 0.9 and "excellent agreement" if

- greater than 0.9.
- 223

224 <u>Table 1: Details of the ICC form chosen for Analyses 1, 2 and 3</u>

	Intra-class correlation coefficient (ICC) form					
	Model Type Definition					
Analysis 1	2-way random effects	Single rater	Absolute			
Agreement among	Both raters & curves are	Each rater is	agreement			
the 7 raters	considered as selected randomly	compared against all				
	from a larger population	others				
Analysis 2	2-way mixed-effects	Single measurement	Absolute			
Agreement between	Raters are fixed & curves are		agreement			
the 2 automated	considered as selected randomly					
algorithms	from a larger population					
Analysis 3	2-way mixed effects	Mean of 7 raters	Absolute			
Agreement between			agreement			
raters and automated						
algorithms						

225 **Results**

226 Analysis 1: Agreement between raters (221 words)

For MRS, ICC value was 0.97 (95% CI [0.96, 0.98]), indicating excellent agreement between

raters (Fig 3). For CPS, ICC value was 0.77 (95% CI [0.69, 0.83]), suggesting good agreement

229	between raters. We hypothesized that the weaker agreement for CPS could be attributed to the
230	difference in raters' expertise level. These scores, both before and after evaluating the 101
231	MNREAD curves, are reported in Table 2. Prior to rating, one rater had no previous experience in
232	rating MNREAD curves (TQ), three raters considered themselves intermediate raters (LM, AM
233	and KB), two raters scored themselves as advanced raters (SM and YH) and one rater reported to
234	be an expert rater (AC). Among the less experienced raters (score 0-2), CPS estimation reliability
235	was only moderate (ICC = 0.70 ; 95% CI [0.57 , 0.80]). Among the most experienced raters (score
236	3-4), it was good (ICC = 0.82 ; 95% CI [0.57 , 0.80]). Interestingly, three raters (43%) considered
237	that their expertise improved (TQ, LM and AM), whereas the remaining four (57%) did not
238	report any change in their expertise level (KB, SM, YH and AC).

240 <u>Table 2: Self-reported score of expertise for our 7 raters</u>

Raters		TQ	LM	AM	KB	SM	YH	AC
Self-reported score of expertise	Prior rating	0	2	2	2	3	3	4
	After rating	1	3	3	2	3	3	4

241

242 <u>Score of expertise in rating low-vision MNREAD data before and after rating the 101 curves (0 – no prior</u>

- 243 <u>experience, 1 novice, 2 intermediate, 3 Advance, 4 Expert).</u>
- 244 Fig 3: Box and whisker plots of estimated MRS (left) and CPS (right), grouped by raters and sorted in
- 245 ascending order of expertise level (from 0 to 4). Boxes represent the 25th to 75th percentiles and whiskers
- 246 <u>range from min to max values. Medians (lines) and means (cross) are also represented.</u>

248 **Analysis 2: Agreement between automated algorithms (245 words)** 249 For MRS, the ICC value of absolute agreement between SDev and NLME methods was 0.96 250 (95% CI [0.91, 0.98]), showing excellent agreement. Contrary to the SDdev method, for which a 251 single MNREAD test yields only one estimate for MRS and one estimate for CPS, the NLME 252 method can generate several measures of CPS depending on the reading-speed criterion chosen to 253 define the CPS (e.g. print size required to achieve 90% of MRS, 80% of MRS, etc.). Therefore, 254 for each of the 101 MNREAD datasets, we estimated five values of CPS with NLME 255 (corresponding to: 95%, 90%, 85%, 80% and 75% of MRS) and measured agreement between 256 SDev and NLME for each of them. The results are reported in Table 3. The strongest agreement 257 between the two automated methods was found for the 80% criterion, and was good, with an ICC 258 value of 0.77 (95% CI [0.68, 0.84]). Additionally, limits of agreement between the two 259 algorithms were estimated using Bland – Altman plots for both MRS and CPS (Fig 4). For MRS, 260 the average difference (i.e. bias) between the SDev method and the NLME model was 5.8 wpm 261 (i.e. 4.5%), with 95% limits of agreement of 11.4 wpm (i.e. 10%). For CPS (defined as 80% of 262 MRS, which showed the best agreement between methods), bias was 0.031 logMAR with 95% 263 limits of agreement of 0.06 logMAR (1 step unit being 0.1 logMAR). Overall, we concluded that 264 no significant difference could be observed between the two automated algorithms.

265

266

267 <u>Table 3: Absolute agreement (ICC values and their 95 % confidence intervals) between CPS values estimated</u>
 268 <u>with the SDev method and the NLME model for five different definitions of CPS.</u>

	ICC value	95% CI	Absolute agreement
95% CPS	0.56	[0.10, 0.77]	
90% CPS	0.70	[0.53, 0.81]	Moderate
85% CPS	0.76	[0.66, 0.83]	
80% CPS	0.77	[0.68, 0.84]	Good
75% CPS	0.76	[0.62, 0.84]	
		[

270 <u>Best agreement is highlighted in grey.</u>

271

272 Fig 4: Bland – Altman plots showing agreement between SDev method and NLME model for both MRS (left)

273 and CPS (right). x-axes represent the mean estimate for both methods; y-axes represent the estimate

274 difference between SDev method and NLME model. Dashed lines show the mean difference (i.e. bias) and the

275 dotted lines represent the 95% CI of limits of agreement (i.e. confidence limits of the bias, defined as the mean

276 <u>difference ± 1.96 times the standard deviation of the difference).</u>

277

278 Analysis 3: Agreement between raters and automated algorithms (139 words)

For MRS, absolute agreement between raters (k = 7) and automated algorithms was found to be

excellent for both the SDev method (ICC = 0.96; 95% CI [0.88, 0.98] and the NLME model (ICC

281 = 0.97; 95% CI [0.95, 0.98]). For CPS, agreement between raters and the SDev method was only

- moderate (ICC = 0.66; 95% CI [0.3, 0.80]), whereas agreement between raters and the NLME
- 283 model was 'good' for CPS defined as 90% of MRS (ICC = 0.83; 95% CI [0.76, 0.88] Table 4
- shows the ICC values for each of the five CPS definitions). Overall, the NLME model showed

285 better agreement with the raters than the SDev method for both reading parameters. Fig 5 shows

the MRS and CPS obtained by the automated algorithms and the 7 raters.

287

- 288 Fig 5: Box and whisker plots showing the median and average MRS (left panel) and CPS (right panel) from
- 289 the two algorithms and the mean of raters. The box represents 25th to 75th percentile with median line and
- 290 <u>the + sign represents the mean and the whiskers represent minimum to maximum.</u>
- 291
- 292 <u>Table 4: Absolute agreement (ICC values and their 95 % confidence intervals) between CPS values estimated</u>
- 293 by the raters and with the NLME model for five different definitions of CPS.

				294
	ICC value	95% CI	Absolute agree	ement
				295
95% CPS	0.78	[0.61, 0.87]		
				296
90% CPS	0.83	[0.76, 0.88]	Good	270
				207
85% CPS	0.79	[0.55, 0.71]		297
80% CPS	0.72	[0.18, 0.88]		298
			Moderate	
75% CPS	0.66	[0.02, 0.87]		299

300

301 **Best agreement is highlighted in grey.**

302

303

304

306 **Discussion (1001 words)**

307

308	In this project we investigated i) the agreement between raters for MNREAD parameters
309	extracted from reading curves (Analysis 1), <i>ii</i>) the agreement between SDev and NLME
310	automated methods extracting reading parameters from raw data (Analysis 2) and <i>iii</i>) the
311	agreement between raters and automated methods (Analysis 3).

312

313 Our first main result was that inter-rater reliability can be classified as excellent for MRS (ICC of 314 0.97) and good for CPS (ICC of 0.77). Because they are lower than 1, these agreement indexes 315 reveal the existence of discrepancies when extracting MNREAD parameters visually from 316 reading curves. Whilst the variability for MRS can be considered residual, the CPS estimation 317 may be questionable. On average, the range of difference in CPS estimates was 0.19 logMAR 318 (i.e. almost 2 lines on a logMAR chart), implying that the variability among raters can be 319 considered clinically significant and potentially problematic, for example when CPS is used to 320 prescribe optimal magnifying power. To identify the underlying factors of the discrepancies 321 observed in CPS rating, we considered whether the data itself could be involved, hypothesizing 322 that the modest ICC value that we found (0.77) was largely due to the presence of highly noisy 323 data. To confirm this hypothesis, we identified extreme outliers for which CPS values were three 324 times larger than the standard deviation of the mean. A total of five curves (5%) were identified 325 as extreme outliers (#2, #31, #58, #70 and #89 in S1 Appendix). What these curves have in 326 common is: the lack of a clear plateau and/or the lack of a clear drop point. After removing these 327 five outliers, the resulting ICC value for CPS improved to 0.82 (95%CI [0.76, 0.87]. This

328 increased value suggests that, to increase inter-rater reliability, ambiguous cases of noisy data 329 should be discussed before final estimates of CPS are reached. Therefore, the advice for our 330 fellow researchers is to inspect our 5 ambiguous samples and define how to deal with such cases 331 on an individual basis whilst maintaining consistency in data extraction. The tips provided in S2 332 Appendix on how to score ambiguous data can serve as a starting point. When possible, 333 measurements should be repeated to help interpret problematic data. 334 335 We also found that for CPS inter-rater reliability was poorer among less experienced raters 336 compared to experienced ones. We speculate that this tendency may be related to both the lack of 337 experience in administrating and rating the test that would lead more naïve raters to follow 338 strictly the definitions of CPS and MRS. Taking the example of curve #2 (see S1 Appendix), 339 raters SM and AC (self-reported expertise scores of 3 and 4) estimated CPS to be 0.7 logMAR 340 (MRS = 68 wpm, both) whilst TQ and KB (self-reported expertise score of 0 and 2) estimated 341 CPS to be 1.3 and 1.1 logMAR (MRS = 85 and 75 wpm, respectively). In this case, the more 342 experienced raters (SM and AC) may have decided to ignore the outlier initial data point, 343 assuming that this measure resulted from experimental noise.

344

Our second main result is the excellent agreement between the two automated methods for MRS.
Regarding CPS estimation, the NLME method provides more flexibility over the SDev method,
since it allows to determine CPS for different levels of MRS. For instance a higher, more
conservative criterion, can be chosen for fluent reading while a lower criterion would be
preferred for spot reading. However, there is no rule yet on how to set this criterion optimally to

350 increase reliability. Our results show that the reading speed cut-off to determine CPS yielding the 351 best reliability between methods is 80% MRS. This result resonates with conclusions from [19], 352 who showed that agreement between NLME models using a two-limb function and an 353 exponential decay function was greater if CPS was set at 80% MRS. On the question of test-retest 354 reliability, [11] also reported that using a criterion of 80% yield improved repeatability of the 355 CPS (when compared to 90%). While an optimal criterion should be chosen on a case-by-case 356 basis depending on the clinician or researcher's motivations, all these evidence suggest that a 357 criterion close to 80 % would increase both inter-rater and test-retest variability.

358

Our third result is that raters and automated methods show excellent agreement for MRS values (ICC of 0.96 and 0.97 for the SDev and NLME respectively). The agreement for CPS was more variable. It was found to be poor for the SDev (ICC of 0.66) and good for the NLME (ICC of 0.83 with a CPS criterion set to 90% MRS). It is worth noting that ICC values were almost identical when measuring agreement between raters and agreement between algorithms for both MRS and CPS. This observation is quite interesting and somehow indicates the robustness and efficacy of human visual inspection of MNREAD curves.

366

367 The represent work presents some limitations. First, despite the relatively large sample of 368 MNREAD data considered in the present work, it is hard to predict to what extent the different 369 shaped curves are representative of the curves found in typical clinical practice. Second, it is 370 likely that the new instructions helped reduce inter-rater variability, but there are no data to 371 support this assumption. While all raters used these extended instructions, the ICC value for CPS

372 was still low, suggesting that additional fixes should be considered to help increase reliability. It 373 is possible to run participants through the test more than once, at least with the English version 374 [16,30]. Repeated measures would make it easier for the rater to determine whether a measure 375 should be considered as noise or not. Another possibility might be to pool estimates from 376 multiple raters or in combination with curve fits. Third, the finding that 80% MRS yields the 377 most reliable CPS using the NLME method is convenient to parameterize the curve in research 378 studies using curve fitting. But for low vision rehabilitation the goal ought to be to enlarge text so 379 that it can be read at the reader's MRS, not at the 80% of the reader's MRS.

380

381 **Conclusions**

382 In summary, our study shows that extraction of the maximum reading speed from MNREAD data 383 is highly consistent across methods and researchers. It also reveals that for low-vision data, it is 384 difficult to obtain excellent inter-rater reliability for CPS estimates. Future studies, such as 385 rehabilitation interventions aiming at improving reading ability in people with low vision, can 386 now follow the advices and instructions resulting from our investigation. Using a standard set of 387 instructions and criteria to analyze reading curves may help increase the reliability of the results. 388 Additional ways to improve inter-rater reliability should also be considered, e.g. use the curve 389 fits, collect multiple runs per participant or combine the estimates of multiple raters.

390 Acknowledgments

391

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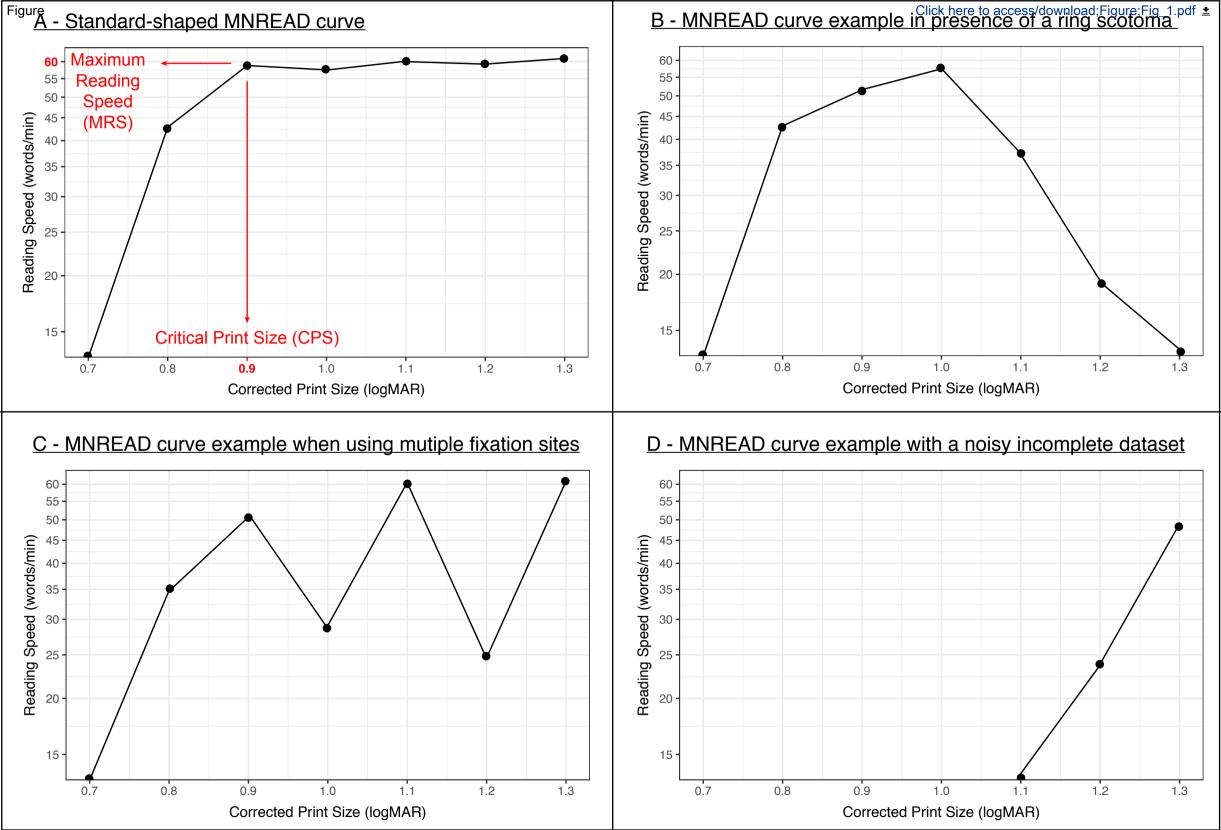
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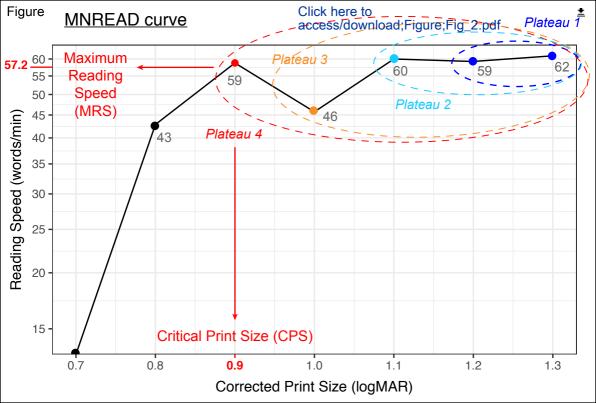
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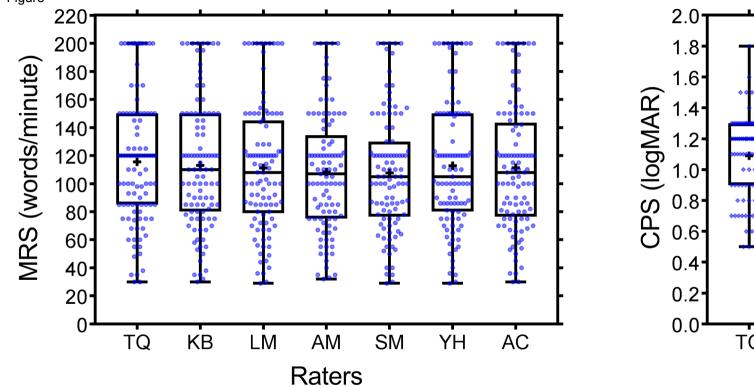
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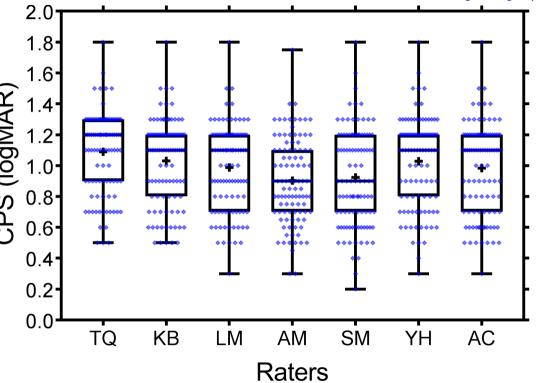
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498	Supporting information captions
499	S1 Appendix. Individual MNREAD curves from the 101 MNREAD measurements.
500	S2 Appendix. Detailed scoring instructions provided to the raters.



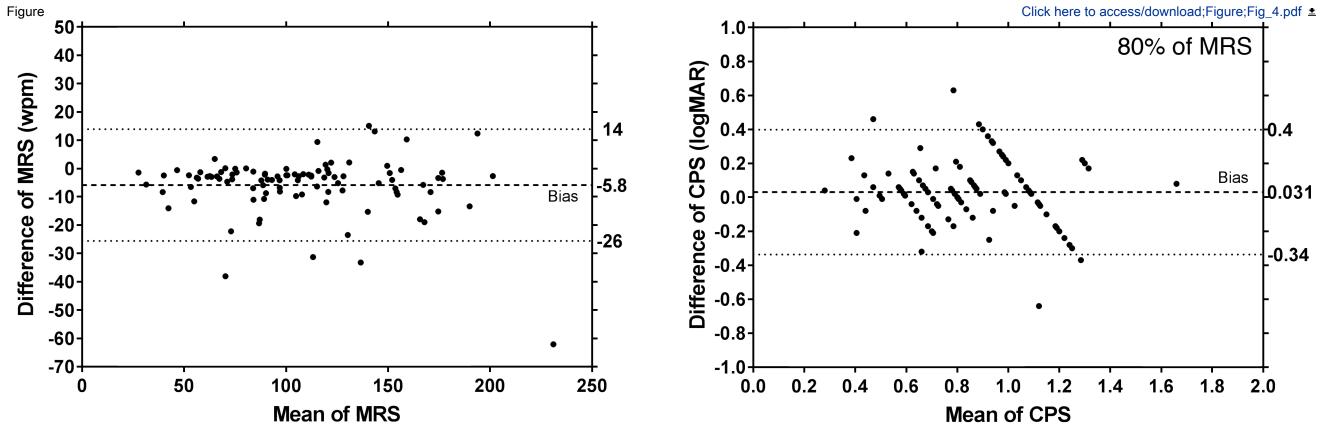






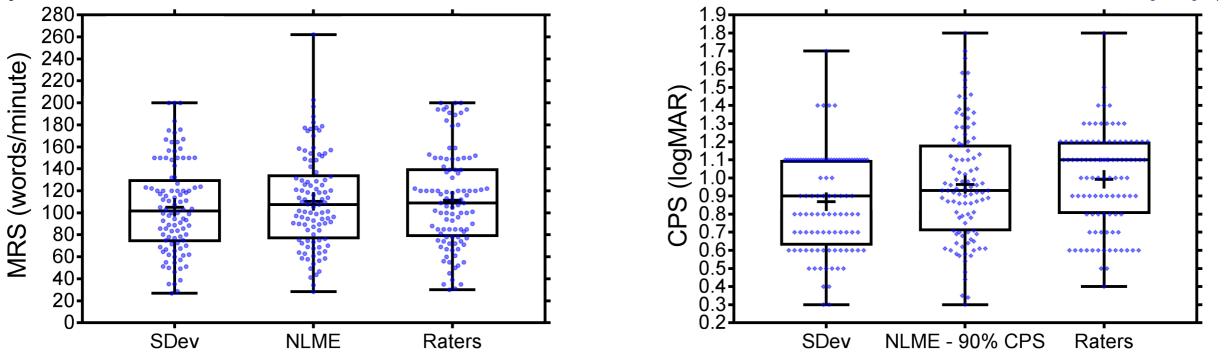


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Figure

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