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Simultaneous Full-Field Multi-Experiment Identification

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Abstract

Identification of constitutive parameters relies mainly on their sensitivity to the measurands. In particular, the specific static and kinematic responses controlled by each parameter of interest has to be captured by full-field measurements. The development of modern constitutive models has led to many new and interesting sample geometries and loading histories, aiming at maximizing the sensitivity to their delicate material parameters, especially through their kinematics of interest. However, it is often impossible to design an experiment that activates all material parameters of interest, and thus multiple experiments are needed. This paper discusses a methodology for combining the data from such multi-experiments into a single identification process to calibrate a complete set of parameters at once. Many different ways of merging experimental data exist, leading to unbiased identifications of the parameters of interest. However, only one *optimal* procedure leads to minimal uncertainty, taking into account the noise of each acquisition source. The proposed identification method is a natural extension of inverse methods such as Finite Element Method Updating (FEMU) with appropriate weights or Integrated Digital Image Correlation (I-DIC). This procedure is illustrated by the identification of the planar parameters of the so-called Hill48 anisotropic yield surface together with an exponential isotropic hardening law of AA2219.

Keywords: Anisotropic plasticity, Digital image correlation, Full-Field

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1. Introduction

Engineering applications are increasingly demanding fidelity of nonlinear constitutive models. The mechanics of materials community has been addressing this demand for many years. The field has significantly progressed in modeling
5 nonlinear mechanisms such as plasticity, viscosity, damage, fracture, delamination. Most of these modern constitutive models have in common that they require more material parameters often difficult to calibrate for a given material. This identification challenge hampers the usability of some of these models, thereby restricting their academic value and their industrial relevance [1, 2].

10 Fortunately, the experimental mechanics community is developing new methods to address these challenges. Currently, full-field measurement methods such as Digital Image Correlation (DIC [3, 4]) and the grid-method [5] have reached a data density level rich enough to identify multi-parameter nonlinear models. The increase in accuracy, efficiency and versatility of these full-field identification
15 methods have greatly expanded the applicability of full-field identification methods [6, 7].

The most common identification method is referred to as Finite Element Model Updating (FEMU [6, 8–13]). In FEMU, the gap between the experiment and simulation of the same experiment is minimized by optimizing (*i.e.*, updating)
20 the unknown model parameters. Within this paper a similar method is applied, which is referred to as Integrated-DIC [14–20]. Integrated-DIC optimizes the gap between simulation and experiment directly on the captured images by integrating the identification step within the DIC algorithm. However, the differences between the two methods are not essential for the discussion
25 in this paper. The interested reader is referred to Refs. [18, 21] for more details on this last point.

The goal of this paper is to propose a method that aggregates measurement results from difference sources into a single identification method while account-

ing for the uncertainty of each measurement. The critical part is the proper
30 weighting of each contribution. Some measurement sources consist of numerous
data with poor quality (*e.g.*, images) while others are limited in number but of
higher quality (*e.g.*, force measurements, strain gauges). The proposed method
is defined such that, the weighting is not an arbitrary choice, but instead nat-
urally follows from a Bayesian formulation given the number of measurements
35 and their respective uncertainty of each source. The obtained weight is opti-
mal, *i.e.*, leading to the minimum uncertainty in the parameters for the case of
Gaussian noise and insignificant model error.

Besides FEMU and Integrated-DIC there are other inverse identification
methods (see *e.g.*, [6, 7] for overview). The present paper focuses on these two
40 methods because both utilize the sensitivity fields in the same way. However, the
concepts discussed herein may be equally valid for other identification methods
such as the equilibrium gap method [22] or the virtual fields method [23, 24].

The proposed simultaneous multi-experiment identification method is pre-
sented using a proof of concept experiment. For the experiment, three uniaxial
45 tensile tests were performed on Aluminum Alloy 2219 at the typical 0° , 90°
and 45° angles with respect to the rolling direction. Both classical identifica-
tion and Integrated-DIC are performed and compared. The identified law is
an elastoplastic material model with Ludwik (power law) isotropic hardening
postulate [25], and the plane-stress Hill anisotropic yield surface [26]. It is well
50 known that Hill’s model is not optimal for describing the “anomalous” behavior
often observed in aluminum alloys [2]. However, a perfect identification of the
model is not the present goal. Validating that the multi-experiment implemen-
tation is able to calibrate an 8-parameter constitutive model using three simple
uniaxial experiments is considered more important.

55 2. Methodology

For the application at hand, the experimental data will be the force mea-
sured by the load cell of the testing machine and the images of the sample

surface recorded for each of the three experiments. It is important to propagate the measurement uncertainties along the entire identification chain. Therefore, the applied global DIC algorithm is briefly recalled, followed by a concise summary of single experiment identification using I-DIC to ultimately combine all ingredients into the multi-experiment identification method.

2.1. Measurement Aggregation

In the following, it will be shown that the weighting factor to be used is not a parameter to be chosen by the user but instead follows from the propagation of uncertainty. Consider N observables, x_i with $i = 1, \dots, N$, for which a model \mathcal{G} can generate corresponding estimates from a set of m parameters $\{p\} = \{p_1, p_2, \dots, p_m\}$ ¹

$$x_i = \mathcal{G}_i(\{p\}). \quad (1)$$

Note that the observables x_i can be of different types with different units (*e.g.*, a set of measured forces and displacements). The measurement \hat{x}_i of the observable x_i is corrupted by Gaussian noise ζ_i

$$\hat{x}_i = x_i + \zeta_i. \quad (2)$$

The statistical distribution of ζ_i is the normal law $\mathcal{N}(0, \gamma_i^2)$ of zero mean and variance γ_i^2 , thus the probability of the estimated observable to be equal to x_i for each measurement \hat{x}_i reads

$$P_i = \frac{1}{(2\pi)^{1/2} \gamma_i} \exp\left(-\frac{(\hat{x}_i - x_i)^2}{2\gamma_i^2}\right), \quad (3)$$

and hence for the entire set of N measurements, assuming they are statistically independent, the probability reads

$$P = \frac{1}{(2\pi)^{N/2} \prod_{i=1}^N \gamma_i} \exp\left(-\sum_{i=1}^N \frac{(\hat{x}_i - x_i)^2}{2\gamma_i^2}\right). \quad (4)$$

¹Throughout this article, the notations $\{\square\}$ and $[\square]$ are used to express column vectors and matrices, reserving the $\vec{\square}$ notation for vector quantities in \mathbb{R}^2 or \mathbb{R}^3 .

Inference of the most likely set of parameters that corresponds to a given set of measurements is equivalent to finding the maximum of Equation (4), or the minimum of the log-likelihood that is (up to irrelevant constants)

$$\eta^2(\{p\}) = \sum_{i=1}^N \frac{(\hat{x}_i - x_i)^2}{\gamma_i^2}. \quad (5)$$

The choice of the prefactor is such that the expectation of η^2 at convergence is N , assuming that the difference $(\hat{x}_i - x_i)$ is only due to noise. It is also noteworthy that the quadratic difference is not any arbitrary choice among many convex functions that are minimum at the origin, but the consequence of the Gaussian probability density function assumed for noise. At this point, typically nonlinear optimization methods are applied to find the optimal set of parameters that minimize Equation (5).

The most important conclusion from Equation (5) is that an identification using different aggregated measurement sets (in the presence of white noise) translates into minimizing the weighted sum of the squared differences. Taking the example of two sets of measurement data, with N_1 and N_2 measurements and variances γ_1^2 and γ_2^2 respectively, and labeled sequentially from 1 to $N = N_1 + N_2$, the cost function is decomposed as

$$\eta_{\text{total}}^2(\{p\}) = \frac{1}{\gamma_1^2} \sum_{i=1}^{N_1} (\hat{x}_i - x_i(\{p\}))^2 + \frac{1}{\gamma_2^2} \sum_{j=N_1+1}^N (\hat{x}_j - x_j(\{p\}))^2. \quad (6)$$

This property endows the above cost function with a very convenient extensivity property that will be used farther down. This equation is equivalent to what is obtained from a variational approach or nonlinear least squares method [18, 19]. However, the important difference is that the weights that are required to combine multiple cost functions follow from the derivation instead of being chosen arbitrarily. Choosing a different weight will still result in an unbiased identification, but it will overemphasize the importance of one measurement source as compared to the others, and hence lead to larger uncertainties. Let us also note that an intensive version can be defined, $\hat{\eta}^2 \equiv \eta^2/N$, in which case the total intensive cost function is the average of that of subgroups, weighted by

the proportion of measurement number, a mathematically equivalent property convenient to assess model error when $\hat{\eta} \gg 1$, but requiring N to be known for assembling different measurements.

The above case was written in the simple case of *independent* measurements. Correlated measurements can be treated in the very same framework and still result in a quadratic form to be minimized but now with a metric that is not Euclidean, but based on a kernel that is the inverse of the covariance matrix of the measurements [27]. If C_{ij} denotes the covariance of the noise affecting measurements of x_i and x_j , then the general expression of η becomes

$$\eta^2(\{p\}) = \sum_{i,j=1}^N (\hat{x}_i - x_i) C_{ij}^{-1} (\hat{x}_j - x_j). \quad (7)$$

2.2. Global DIC

Digital image correlation aims at minimizing the difference in the observed intensity of a reference image f and an image taken in the deformed state g within some region of interest. The cost function that is minimized is defined as

$$\eta_{\text{dic}}^2 = \frac{1}{2\gamma_I^2} \sum_{k=1}^{N_k} (f_k - \tilde{g}_k)^2, \quad (8)$$

85 where N_k is the number of pixels considered in the region of interest, f_k is the gray level of the k^{th} pixel at the location \vec{x}_k , γ_I^2 the variance of the noise on each pixel, and \tilde{g}_k the gray level at the location $\vec{x}_k + \vec{u}_k$ in the image g . A factor of 1/2 has been included in the prefactor since noise affects both reference and
 90 deformed images, and hence with this factor the expectation of η_{dic}^2 is equal to the number of pixels. Note that \tilde{g} typically requires gray level interpolation at non-integer pixel locations.

This DIC cost function is to be optimized with respect to the displacement field parameterized by a finite set of degrees of freedom. Since the proposed identification method already relies on finite elements, it is natural to adopt the very same finite element mesh for this purpose, for which the displacement field

reads

$$\vec{u}_k = \sum_{i=1}^{N_a} \vec{\varphi}_{ki} a_i, \quad (9)$$

where a_i are the nodal degrees of freedom and $\vec{\varphi}_{ki}$ the corresponding finite element shape functions. In this paper, a single camera system is adopted resulting in 2D measurements. Ergo, the number of nodal degrees of freedom 95 N_a is two times the number of mesh nodes.

Newton's method is then applied to iteratively minimize Equation (8) starting from an initial guess $\{a_0\}$. The solution at iteration $n + 1$ is defined as

$$\{a\}^{(n+1)} = \{a\}^n + \{\delta a\}, \quad (10)$$

where $\{\delta a\}$ is the iterative update. It is found by solving the linearized cost function

$$[M^{\text{dic}}]\{\delta a\} = \{b^{\text{dic}}\}, \quad (11)$$

where

$$M_{ij}^{\text{dic}} = \sum_{k=1}^{N_k} (\nabla f_k \cdot \vec{\varphi}_{ki}) (\vec{\varphi}_{kj} \cdot \nabla f_k), \quad (12)$$

$$b_i^{\text{dic}} = \sum_{k=1}^{N_k} \nabla f_k \cdot \vec{\varphi}_{ki} \left(f_k - \tilde{g}_k(\{a\}^n) \right). \quad (13)$$

The above global DIC formulation is common and discussed in the literature (see *e.g.*, Refs. [28, 29]).

Besides operating as the tangent matrix in the optimization algorithm, the Hessian $[M_{\text{dic}}]$ serves a second purpose as its inverse represents the covariance matrix $[C^{\text{dic}}]$ if converged

$$[C^{\text{dic}}] = [M^{\text{dic}}]^{-1}, \quad (14)$$

providing the means to relate gray level uncertainty to nodal displacement uncertainty while accounting for element size and pattern contrast variations.

100 *2.3. Single-Experiment Identification*

The cost function for an Integrated-DIC routine is similar to the DIC cost function with the major change being that the unknowns now represent N_p material parameters. A dimensionless version of the parameters is applied to allow for relative comparisons in sensitivity and uncertainty later on in the paper

$$\{p\} = \left\{ \frac{p_1}{p_1^*}, \frac{p_2}{p_2^*}, \dots, \frac{p_{N_p}}{p_{N_p}^*} \right\}, \quad (15)$$

where for the scaling values p_i^* the initial parameters are used (*i.e.*, $\{p_0\}$). The Integrated-DIC cost function then becomes

$$\eta_I^2(\{p\}) = \frac{1}{2\gamma_I^2} \sum_{t=1}^{N_t} \sum_{k=1}^{N_k} \left(f_k - \tilde{g}_{kt}(\{p\}) \right)^2, \quad (16)$$

where N_t is the number of time steps considered for identification and γ_I^2 the estimated gray level variance for each acquisition. The additional summation over time reflects the fact that the constitutive parameters are equally valid for all time steps and according to Section 2.1, the different log-likelihood function-
 105 als are to be simply summed. Hence, the gray level \tilde{g}_{kt} is taken at the location $\vec{x}_k + \vec{u}_{kt}$ of the image acquired at time step t . The displacement field $\vec{u}_{kt}(\{p\})$ is provided by an FE simulation using the current estimate of the parameters.

Similarly additional cost functions can be defined for other observables that will be used for identification purposes. In the test case discussed hereafter, besides images, the reaction force at one of the grips is also utilized

$$\eta_F^2(\{p\}) = \frac{1}{\gamma_F^2} \sum_{t=1}^{N_t} (F_t^{\text{exp}} - F_t(\{p\}))^2, \quad (17)$$

where F_t^{exp} and F_t are respectively the measured and simulated resultant force at the sample boundary at time step t . Similarly as with Equation (16), the cost
 110 function is normalized using the estimated variance for each measurement γ_F^2 .

Exactly as in Equation (6) in Section 2.1, the cost functions for each observable is combined by summation in the normalized forms as expressed in Equations (16) and (17),

$$\eta_e^2 = \eta_I^2 + \eta_F^2. \quad (18)$$

The solution to this inverse problem is obtained by iteratively minimizing this cost function using Newton's method starting with an initial guess for the parameters $\{p_0\}$

$$\{p\}^{(n+1)} = \{p\}^n + \{\delta p\}, \quad (19)$$

where the incremental update to the parameters $\{\delta p\}$ is found by solving

$$([M_I] + [M_F]) \{\delta p\} = \{b_I\} + \{b_F\}, \quad (20)$$

with

$$[M_I] = \frac{1}{2\gamma_I^2} [S_I]^\top [M^{\text{dic}}] [S_I], \quad (21)$$

$$\{b_I\} = \frac{1}{2\gamma_I^2} [S_I]^\top \{b^{\text{dic}}\}, \quad (22)$$

and

$$[M_F] = \frac{1}{\gamma_F^2} [S_F]^\top [S_I], \quad (23)$$

$$\{b_F\} = \frac{1}{\gamma_F^2} [S_F]^\top \{F^{\text{exp}} - F\}, \quad (24)$$

where $[S_I]$ and $[S_F]$ represent the displacement and the force sensitivities. They are rectangular matrices projecting from the node-time space to the parameter space and follow from the derivation using Newton's method

$$\{S_{Ii}\} = \frac{\partial \{a\}}{\partial p_i}, \quad (25)$$

$$\{S_{Fi}\} = \frac{\partial \{F\}}{\partial p_i}. \quad (26)$$

In some special cases, analytical derivatives are obtainable (see *e.g.*, Refs. [30–33]). However, for most constitutive models, approximative derivatives have to be evaluated, as for example via finite differences. In the above equations, the sensitivity matrices are computed by performing one simulation using the
115 current estimate of the parameters to obtain the reference nodal values $\{a\}$ and $\{F\}$ and a second one per each unknown p_i where the parameter is increased by 1 %.

Another interpretation of Equation (21) is that the displacement field (instead of the images) is considered the measurand and the $[C^{\text{dic}}]$ matrix is applied as the metric with which to transport covariance from a previous measurement (DIC) to the identification method. This observation highlights that the build-up of covariance is transparent to the identification method if it is consistently transported along the measurement chain (see also Equation 7). In other words, performing FEMU (with the displacement field as measurand), results in exactly the same cost function definition as I-DIC (using the image as measurand), provided the covariance accumulated during the DIC phase is taken into account (see also Ref. [18]).

2.4. Multi-experiment identification

The premise that drives multi-experiment identification is that the constitutive parameters are valid for a range of experiments of the studied material. Consequently, different experiments may be performed to activate different material parameters. While some experiments may be relevant to trigger sensitivity for specific parameters, other parameters will be active in all experiments. The proposed method of measurement aggregation (see Section 2.1) and its adoption in the previous section to aggregate multiple data sources into a single experiment is extended to aggregate data from multiple experiments

$$\eta^2 = \sum_{e=1}^{N_e} \eta_e^2, \quad (27)$$

where the total cost function η^2 is simply the sum of the individual cost functions for each experiment, as defined in the previous section, N_e being the total number of experiments.

The chosen solution strategy remains unchanged from Newton's method as defined in Equation (19), where the update to the degrees of freedom obtained by solving the linearized system of equations reads

$$[M]\{\delta p\} = \{b\}, \quad (28)$$

where the Hessian $[M]$ and the right hand member $\{b\}$ are the sums of the “per experiment” respective contributions

$$[M] = \sum_{e=1}^{N_e} ([M_{Ie}] + [M_{Fe}]), \quad \{b\} = \sum_{e=1}^{N_e} (\{b_{Ie}\} + \{b_{Fe}\}), \quad (29)$$

thereby resulting in one global Hessian and right hand member seamlessly combining the data from multiple experiments. Iteratively solving this system of equations will result in finding a single set of parameters that optimally minimizes the discrepancy between multiple experimental sets of data and multiple
135 simulations as gauged by their uncertainty.

2.5. Identification framework

The previous sections have introduced the general identification procedure. However, details still remain to be discussed for the method to become practical.

140 *Boundary Conditions.* The previously discussed approach relies on simulating the experiment using FE models. The quality of the parameter identification relies on the simulations to mimic the experiments as closely as possible and thus they need to be driven with accurate boundary conditions. The chosen method is to use a regular (*i.e.*, non-integrated) DIC routine to measure the
145 displacements near the edges of the field of view and use these displacements as boundary conditions for the simulation. These displacements are measured by DIC and thus are affected by noise creating an alternative way for measurement noise to influence the identification. The influence of this noise is an ongoing research subject and is not considered in the analyses and results [34]. However,
150 effort is made to minimize this influence, which will be discussed in Section 2.9. On the other hand, the method of using the DIC measured boundary conditions has a few advantages, namely, it limits the simulation domain to the region of interest, it provides time synchronization, it naturally deals with any sample misalignment and rigid body motion [35].

Conditioning. Although the proposed method condenses a large amount of data to a relatively smaller system of equations to be solved (Equation (28)), the

system can still be poorly conditioned. Reasons for ill-conditioning can be, for example, insufficient sensitivity to a given parameter (or linear combinations thereof). To limit the impact of these issues, Tikhonov regularization is applied [36]

$$\left([M] + \alpha[I]\right)\{\delta p\} = \{b\} + \alpha(\{p_{\text{ref}}\} - \{p\}), \quad (30)$$

155 where $[I]$ is the identity matrix, α the regularization strength set to $\alpha = 10^{-5}\lambda$, where λ is the largest eigenvalue of $[M]$ and $\{p_{\text{ref}}\}$ is a set of parameters obtained from external sources such as other experiments of the literature (here chosen as $\{p_0\}$ the initial guess of the minimization procedure). The regularization strength is set sufficiently low such that parameters with sufficient sensitivity will
 160 be calibrated by the multi-experiment identification while suppressing spurious behavior of the ill-conditioned parameters. Note that when $\alpha = 0$ the original system is obtained.

Secant method. For the proposed method, most of the cost is in the finite element simulations. Consequently, updating the sensitivity matrices (Equation (25) and 26) is costly, since they require $N_p + 1$ simulations. Therefore,
 165 within this paper the sensitivity matrices are not updated for every iteration. This effectively modifies Newton’s scheme to a secant method. The update in the sensitivity matrices is only performed when the cumulative update in the parameters since the last sensitivity update is larger than 5 %.

170 *Algorithm Summary.* The implemented framework consists of the following steps:

1. Measure the boundary conditions using DIC
2. Choose the initial parameter set $\{p_0\}$
3. Compute the reference displacement field and force for each experiment
- 175 4. If large changes in parameters, compute the sensitivity matrices for each experiment (Equations (25) and (26))
5. Aggregate the data from all experiments (Equation (29))

6. Solve for the correction in the parameters (Equation (28))
7. Update the parameters (Equation (19))
- 180 8. If not converged return to step 3

Most of the computational cost lies in steps 3 and 4 since they involve a number of often heavy finite element simulations. However, these simulations can be performed concurrently while each of them can be parallelized. This can greatly reduce computation time if the computational resources are available. Typically step 5 can be performed within a few minutes for a moderate data set on a normal desktop computer and thus its cost will only be significant for cases with comparable finite element simulation costs.

2.6. Experiment

The three typical tensile experiments treated in this paper are those that are commonly performed for identifying planar anisotropic yielding [37, 38]. The samples were cut such that their respective loading axes aligned with three important rolling direction, namely the Longitudinal (L), Transverse (T) and Diagonal (D) directions. The longitudinal direction was aligned with the rolling direction of the sheet at 0° . The transverse sample was cut at 90° while the diagonal sample was cut at 45° with respect to the rolling direction, as shown in Figure 1. These sample directions will be denoted as L, T and D henceforth.

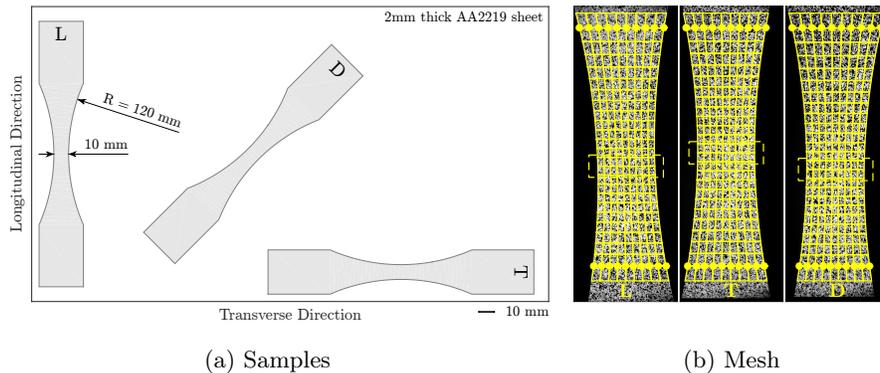


Figure 1: (a) Orientation of the samples as cut from the aluminum alloy sheet. (b) Picture of the three samples with their speckle pattern and conforming mesh.

Each sample was 2 mm thick and 180 mm long, the width at the narrowest section was 10 mm, with the dog-bone radius of each side of the sample equal to 120 mm, see Figure 1. The samples were loaded in a servohydraulic
 200 tension/compression testing machine equipped with a 50 kN load-cell. The top grip of the tensile machine was stationary while the bottom grip was driven at constant velocity of 0.01 mm/s to specific load levels (*i.e.*, [1.8, 6.0, 8.5, 9.1, 9.3, 9.5, 9.6, 9.65] kN). When a load level was reached the sample was unloaded to 0.1 kN upon which the next loading cycle starts.

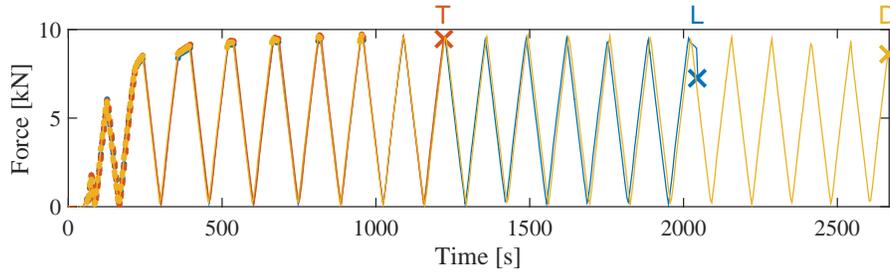


Figure 2: Force response over time for all three experiments. This figure shows all experimental cycles the three samples, indicating where each sample ultimately failed (marked X) and the images that were considered herein (marked •). Note that the first two cycles were used with unloading, while upon plasticity only the parts where the force was monotonically increasing are used up to the 8th cycle.

205 Only the first 8 cycles are considered as the experimental dataset in this paper. Moreover, for the “plastic” cycles, only the images where the force was monotonically increasing are kept in the dataset (Figure 2). The later cycles were discarded because the applied constitutive model (to be introduced below) was not rich enough to allow for parameter calibration after the onset of necking.
 210 Each experiment was started holding the load at 0 N while capturing 20 images. These images are only used to estimate the measurement uncertainty. The identification dataset is based on [114, 110, 114] images per experiment respectively, 1 reference and [113, 109, 113] deformed increments containing two elastic load-unload cycles and the monotonic loading part up to 9.65 kN of
 215 load.

The images were captured with a single Manta G-223 camera equipped with a telecentric lens of magnification $\times 0.125$, and captured one side of the speckle-painted sample twice at 5 s intervals, once at 1 s and once at 5 s in each interval. The recorded images are of definition 1120×2160 px, where each pixel captures
220 the intensity, digitized with 16 bits, of $26 \times 26 \mu\text{m}^2$ of the sample area. The three experiments were performed consecutively without any changes to the camera and lighting settings. Consequently, the dynamic range of the images, which is defined as the difference between the brightest and the darkest pixel, is 63,600 gray levels over all images and experiments.

225 Besides using the images for the full-field identification procedure the images are first processed using a relatively standard Finite Element (FE) based global DIC algorithm [29]. The adopted FE mesh (used for DIC) for each experiment is shown in Figure 1(b) using 207 4-noded bilinear elements with 240 nodes in total for each experiment. The meshes are adjusted to the geometry as
230 captured in the reference image of each experiment. Although the meshes are equal in topology, they are slightly different in shape due to differences in sample alignment. Consequently, the number of pixels supporting each mesh is [691, 686, 700] $\times 10^3$ pixels for the L-, T-, D-experiments respectively.

All full-field data discussed in this paper are 3D, namely 2D space and 1D
235 time. Furthermore some are vector or tensor fields. Consequently, a condensed plotting style is adopted throughout the paper that merits a few words of explanation (see *e.g.*, Figure 3). The 3D data volume is represented by two 2D cross-sections, the first is in the time- y plane taken at $x = [200, 145, 170]$ px respectively for the L-, T- and D-samples, the second is in the x - y plane taken
240 at increment [108, 104, 108] respectively. Both cross-section planes are indicated in the figures with dashed lines in the other cross-section. Furthermore, various vector components of the fields are plotted as rows of sub-figures. Single point data (*e.g.*, force) are plotted when relevant using the same time axes below the final vector/tensor component time- y figure. In this visualization style all
245 data considering a single experiment can be compactly summarized in a vertical arrangement, allowing the data of the three experiments (L, T and D) to be

plotted side by side for easy comparison.

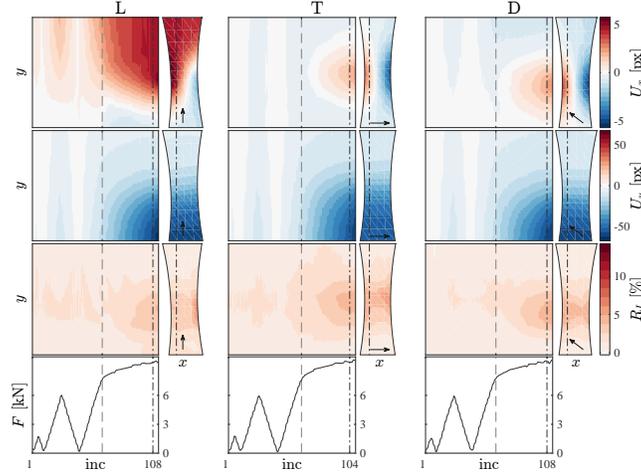


Figure 3: Displacement field (in space and time) as measured by non-integrated DIC for all three experiments (rows 1 and 2), and the respective image residuals (row 3) and the measured force at the load-cell (row 4).

Figure 3 summarizes the full-field space-time displacement data as obtained by DIC for all three experiments. In the experiments, the top grip of the tensile machine is fixed and the bottom grip is driven. However, since the imaged field of view is smaller than the clamped sample length (*i.e.*, ~ 48 mm) both top and bottom mesh boundaries are moving vertically. The L-sample was mounted with a visible misalignment that resulted in a significant horizontal motion of the bottom mesh boundary due to the lateral freedom of the bottom grip. Due to the limited number of samples, the experiment was not discarded but the data were used as-is. Fortunately, the adopted full-field identification method is naturally tolerant to these real-world experimental issues.

DIC is also applied to the first 20 images, captured without load on the sample. The root-mean-square (RMS) of the image residuals, standard deviation of the nodal displacements and the standard deviation of the measured force over these 19 time steps are used as estimates of the respective measurement

uncertainties

$$\begin{aligned}\gamma_I &\approx 1911 \text{ GV} \approx 2.9\%, \\ \gamma_U &\approx 0.017 \text{ px} \approx 0.44 \text{ }\mu\text{m}, \\ \gamma_F &\approx 5 \text{ N.}\end{aligned}\tag{31}$$

where the gray level uncertainty is computed using

$$\gamma_I = \frac{1}{6}\sqrt{6}G\langle\gamma_U L\rangle,\tag{32}$$

where G is the mean field average of the image gradient and L the element length which is approximated as the square root of the average element area [29, 39].

260 The images are recorded using 16-bit gray levels, resulting in the gray level uncertainty being approximately 2.9% of the dynamic range. This method of obtaining the gray level uncertainty is used to be consistent with a previous study that used the same images [20]. In the latter reference, this estimate was introduced as an effective gray level uncertainty that also accounts for gray level
265 variations from other sources such as cross-pixel correlation, speckle degradation and sub-pixel interpolation errors. The impact of this choice is currently under research but is estimated to be small.

2.7. Constitutive Model

The constitutive model chosen to be identified using the three experiments is an elastoplastic law with isotropic exponential hardening and an anisotropic yield surface. The chosen hardening law is that proposed by Ludwik [25]. In a 1D setting, under increasing plastic strain ε_p the stress (σ) strain (ε) relationship reads

$$\sigma = E\varepsilon, \quad \sigma \leq \sigma_0, \tag{33}$$

$$\sigma = \sigma_0 + h\varepsilon_p^m, \quad \sigma > \sigma_0, \tag{34}$$

where E is the Young's modulus, σ_0 the yield stress, h the hardening modulus, and m the hardening power. Together with Poisson's ratio ν , this model results in five unknown material parameters. The anisotropic yield surface is described

with Hill's model [26], which is readily available in Abaqus [40]. The anisotropic plasticity criterion is defined as

$$H_1(\sigma_{22} - \sigma_{33})^2 + H_2(\sigma_{33} - \sigma_{11})^2 + H_3(\sigma_{11} - \sigma_{22})^2 + 2H_4\sigma_{23}^2 + 2H_5\sigma_{31}^2 + 2H_6\sigma_{12}^2 = 1, \quad (35)$$

with

$$H_1 = \frac{1}{2} \left(\frac{1}{R_{22}^2} + \frac{1}{R_{33}^2} - \frac{1}{R_{11}^2} \right), \quad H_4 = \frac{3}{2} \frac{1}{R_{23}^2}, \quad (36)$$

$$H_2 = \frac{1}{2} \left(\frac{1}{R_{33}^2} + \frac{1}{R_{11}^2} - \frac{1}{R_{22}^2} \right), \quad H_5 = \frac{3}{2} \frac{1}{R_{13}^2}, \quad (37)$$

$$H_3 = \frac{1}{2} \left(\frac{1}{R_{11}^2} + \frac{1}{R_{22}^2} - \frac{1}{R_{33}^2} \right), \quad H_6 = \frac{3}{2} \frac{1}{R_{12}^2}, \quad (38)$$

where R_{11} , R_{22} , R_{33} , R_{12} , R_{13} and R_{23} are yield stress ratios. For a near plane-
270 stress state, these ratios are interrelated such that four parameters suffice to fully describe the yield surface [2]. Therefore, the parameters R_{22} , R_{33} , R_{12} and σ_0 are used adding 3 parameters to the identification routine. A large transformation framework is used in the FE simulations.

2.8. Classical Identification Methods

A more classical identification approach is first applied before using the proposed full-field multi-experiment identification. This analysis serves two purposes, (i) provide a set of parameters $\{p_0\}$ to initialize the identification, (ii) provide a reference to compare the identified parameters with. This classical identification is not considered as a golden standard, yet for the test case at
280 hand, both methods are within their application range and should provide comparable results.

The non-integrated DIC displacement measurements are used to compute the logarithmic strain fields that are shown in Figure 4. A virtual strain gauge is applied by averaging the strains in the center region of the samples indicated
285 by the dashed box in Figure 1(b).

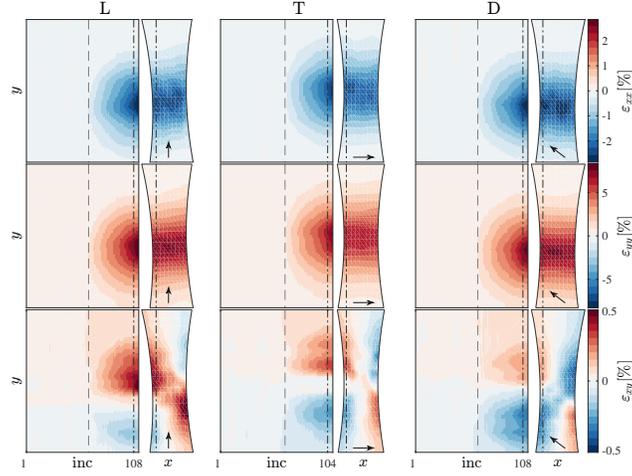


Figure 4: Strain fields (in space and time) as measured by non-integrated DIC for all three experiments.

The stress in the tensile direction of the sample (*i.e.*, y -direction) is estimated using the measured force F , the cross-sectional area in the reference state A_0 and the stretch in the same direction

$$\sigma_{iyy} \approx \frac{F_i \exp(\varepsilon_{iyy})}{A_0}, \quad i = [L, T, D], \quad (39)$$

Equation (34) is fitted to the L-data using Newton's method to obtain the parameters E , σ_0 , h and m . The ratio between the xx - and yy -components of strain for the sample experiment are used to identify ν . The result of the stress-strain curve obtained by the fit is shown with a dashed line in Figure 5(a).

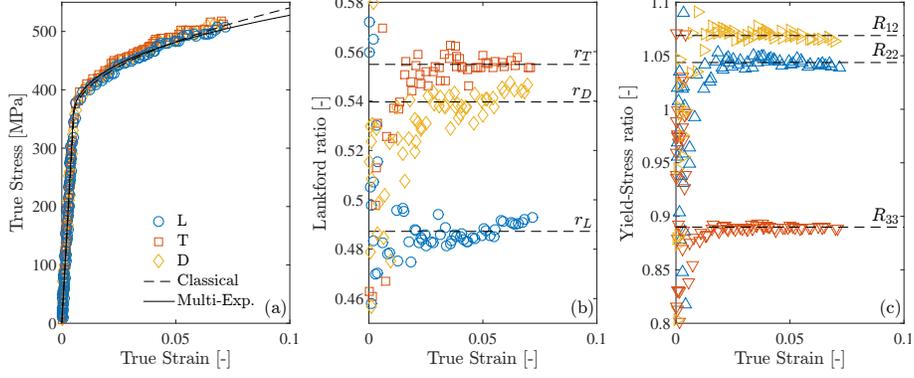


Figure 5: Plots of (left) stress-strain data for all three experiments, (center) Lankford ratio vs. strain and (right) Hill's yield stress ratio vs. strain

From each experiment Lankford's ratios are extracted using the plastic strains ε_p from the two directions in the cross-sectional plane (*i.e.*, the xx - and zz -directions)

$$r_i = \frac{\varepsilon_{p_{i_{xx}}}}{\varepsilon_{p_{i_{zz}}}}, \quad i = [L, T, D], \quad (40)$$

where the plastic strains in those directions are estimated using the total strain the xx - and yy -directions and the parameters identified in the previous section

$$\varepsilon_{p_{xx}} = \varepsilon_{xx} + \nu \frac{F}{AE}, \quad (41)$$

$$\varepsilon_{p_{yy}} = \varepsilon_{yy} - \frac{F}{AE}, \quad (42)$$

$$\varepsilon_{p_{zz}} = \log \left(\frac{1}{\exp(\varepsilon_{p_{xx}} + \varepsilon_{p_{yy}})} \right). \quad (43)$$

290 The Lankford ratios for each time increment are shown in Figure 5(b).

Last, the three yield stress ratios are obtained from the three Lankford ratios

$$R_{22} = \frac{\sigma_{22}}{\sigma_0} = \sqrt{\frac{r_T(r_L + 1)}{r_L(r_T + 1)}} \quad (44)$$

$$R_{33} = \frac{\sigma_{33}}{\sigma_0} = \sqrt{\frac{r_T(r_L + 1)}{r_L + r_T}} \quad (45)$$

$$R_{12} = \frac{\sigma_{12}}{\tau_0} = \sqrt{\frac{3r_T(r_L + 1)}{(2r_D + 1)(r_L + r_T)}} \quad (46)$$

where 1, 2 and 3 indicate directions corresponding to the material orientation, 1 and 2 are in the plane of the metal sheet while 3 is perpendicular to the plane of the sheet. The yield stress ratios for each time increment are shown in Figure 5(c). The values above 3 % strain are averaged to identify the yield stress ratios.

2.9. FE Simulations

The simultaneous multi-experiment identification algorithm detailed in the previous section is implemented within the Correli 3.0 framework, which is in continuous development at LMT [41]. The framework contains the tools to perform non-integrated DIC, manage the launch and data retrieval of all the FE simulations. For this paper Abaqus² is used to perform the FE simulations. In general, Integrated-DIC treats the simulation software as a “black box” where each simulation is started with a different set of parameters and the only exploited output data are the displacement fields and nodal forces for each time step.

The FE meshes used for the simulations are based on those shown in Figure 1(b) with two modifications. The first is the removal of one row of elements from each side of the sample (see Figure 6). This is done so that the applied boundary conditions are those measured by DIC on internal nodes, which are less sensitive to acquisition noise compared to boundary nodes [29]. To further reduce the impact of measurement noise the measured displacements are filtered using 4th order polynomials [42]. The second modification to each mesh is that it is extruded in the z -direction to create a 3D model (see Figure 6). Even though the applied measurement is 2D, there is no limitation on performing the simulations in 3D. It was shown that it is desirable when dealing with elastoplastic laws [35].

²Abaqus Standard: Dassault Systèmes Simulia [40]

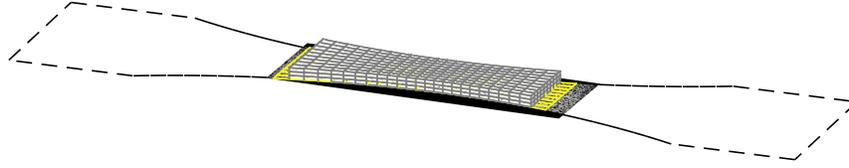


Figure 6: Complete sample shape (dashed line) with the image positioned on top with the 2D DIC mesh (yellow) and the 3D FE mesh.

3. Identification Results

The main objective of identification is the set of parameters. However, the proposed method also allows for the analysis of the obtained result providing
 320 understanding about the sensitivity and uncertainty for each parameter. The analysis is mainly based on the sensitivity matrices (Equations (25) and (26)), the Hessians (Equations (21), (23) and (29)) and the full-field residuals.

3.1. Sensitivity Matrices

Figure 7 shows the sensitivity fields for the five isotropic parameters for the
 325 L-experiment. Similar figures can be created for the other two experiments. However, these will not show significantly different results. The figure is composed of five columns of five sub-figures each. Each column corresponds to one parameter. The displacement component of the sensitivity matrix for one parameter has the same data structure as a displacement field, which is in this
 330 case 3-dimensional with two dimensions in space and one in time. Consequently, the same representation can be used as shown in Figure 3. The full field data in the displayed figures (first and second rows) are scaled per parameter with the amplitude that is given above the column \bar{p}_i . This normalization is mainly performed to aid visualization since it brings all the drawn fields within the same
 335 $[-1, 1]$ range allowing a single color bar. Since the parameters $\{p\}$ are dimensionless, the units of the sensitivity matrices are equal to the respective measured quantities, namely, px for the displacements and N for the force sensitivity.

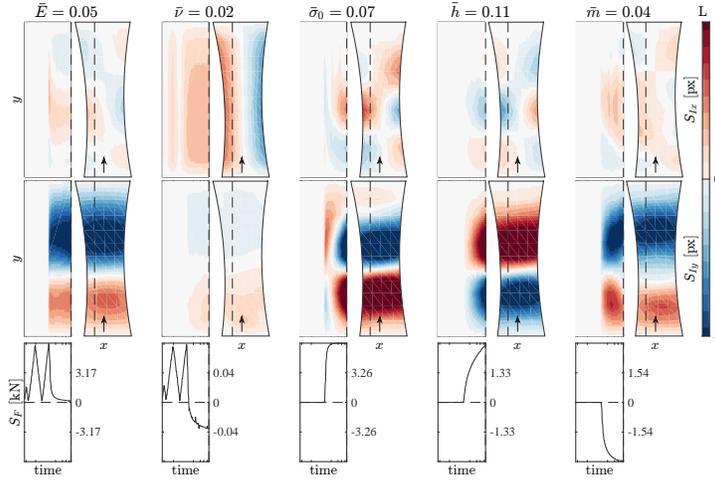


Figure 7: Sensitivity fields in terms of the x - and y -displacement and force sensitivity for the five isotropic parameters for the L-sample each scaled with the amplitude \bar{p} given on top.

Figure 7 quantitatively shows where in space and time each parameter is sensitive. For instance, it illustrates that the second parameter ν has the weakest sensitivity, indicated by its normalization factor $\bar{\nu} = 0.02$ px and its force amplitude being approximately 0.05 N. Both are within the expected measurable range and thus this parameter should be tunable. However, it will be the most prone to measurement noise. It should be noted that it is possible to compute these figures before performing the experiment. Consequently, they are a useful tool for optimizing the experiment to enhance the sensitivity of certain parameters [34]. For example, as expected, most of the sensitivity for the plastic parameters comes from the end of the experiment.

Figure 8 shows the sensitivity fields for the first anisotropic parameter (*i.e.*, R_{22}) for each of the performed tests (*i.e.*, L, T, and D). As expected, the contribution of each experiment to the sensitivity is distinct. Interestingly, the L and D experiments show the strongest displacement sensitivity while the T experiment shows the strongest force sensitivity. The difference in sensitivity in the force part is much more significant showing a 50 times increase in sensitivity for the T experiment compared to the L experiment.

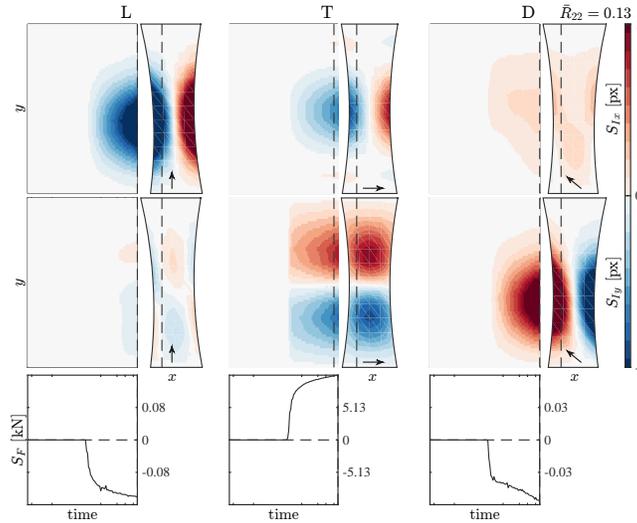


Figure 8: Sensitivity fields for the anisotropic stress ratio R_{22} for each of the three tensile directions (L, T and D).

355 Figure 9 shows similar results for the other two anisotropic stress ratios R_{33} and R_{12} . The R_{33} stress ratio shows equivalent kinematic sensitivity fields for all three experiments while the static sensitivity is approximately 10 times stronger for the D-sample. As expected, the shear stress ratio R_{12} requires the 45° experiment to activate significant sensitivity. This sensitivity field analyses
 360 show that all three anisotropic parameters are significantly activated in one or more experiments both from kinematic and static points of view. Sensitivity for the R_{22} parameter depends mostly on the T-sample while R_{33} and R_{12} mostly on the D-sample. Interestingly, it seems that the L-sample is not significantly carrying sensitivity not already present in the other two experiments. The data
 365 from the L-sample to the identification will however increase the sensitivity for these three and the five isotropic parameters. Since uncertainty is inversely related to sensitivity it is advised to use all three experiments. However, if some uncertainty can be tolerated, the experimental load can be reduced by omitting the L-experiment.

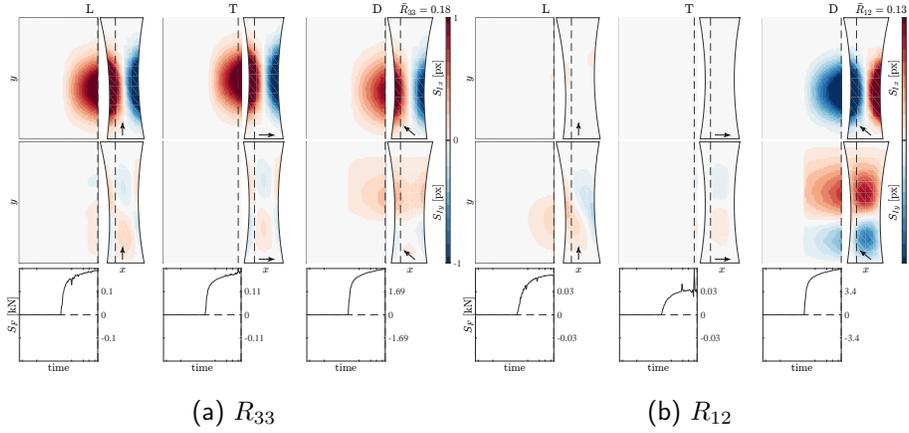


Figure 9: Sensitivity fields for the anisotropic stress ratios R_{33} (a) and R_{12} (b) for each of the three tensile direction (L, T and D).

3.2. Uncertainty quantification

The full-field sensitivity analysis is effective for analyzing where in space and time each parameter is active. However, the final system of equations that is solved during the iterative process is a condensed version of the sensitivity fields (see Equation (29)). The Hessian is the sum of many sub-Hessian matrices, namely, one displacement and one force per experiment. Each experiment generates two sub-Hessians, only the two for the L-experiment are visually represented in Figure 10. This figure shows that the image (or kinematic) part of the measurement carries a significant portion of the sensitivity. It also shows that some groups of parameters are dependent on each other. This is visible from a block diagonal aspect of the Hessian. However, to further interpret the structure of this Hessian, one should consider the eigenvectors and eigenvalues of $[M]$, as the presence of off-diagonal terms means that the actual uncertainty is more intricate. This is considered out of the scope of this paper.

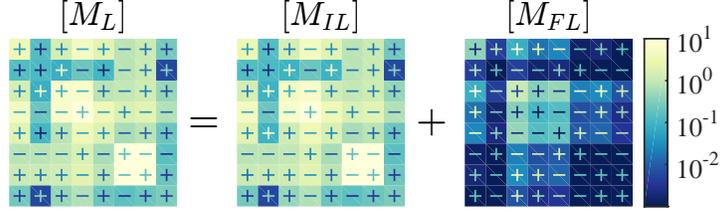


Figure 10: Hessian matrix for the L-sample $[M_L]$, which is the sum of its kinematic part $[M_{IL}]$ and static part $[M_{FL}]$. The matrix rows and columns correspond to the constitutive parameters in order $E, \nu, \sigma_0, h, m, R_{22}, R_{33}, R_{12}$.

Figure 11(a) shows the aggregation of the per experiment Hessian matrices into a single Hessian. Note how the sensitivity accumulates from the three experiments. The parameter sensitivity of the three experiments is perhaps more clearly illustrated in the covariance matrix, which is the inverse of sensitivity, namely, uncertainty (see Figure 11(b)). For Gauss-Newton algorithms, the covariance matrix can be approximated by the inverse of the Hessian in the converged state [43]. Equivalently as the Hessian matrix expresses sensitivity, so is uncertainty expressed by the covariance matrix. Combining the three experiments into one identification results in a significantly lower variance and covariance for the anisotropic parameters. To show the correlation between parameters, the covariance matrix is normalized to obtain the correlation matrix (no index summation)

$$Q_{ij} = \frac{M_{ij}^{-1}}{\sqrt{M_{ii}^{-1}M_{jj}^{-1}}}, \quad (47)$$

which is shown in Figure 11(c). The two elastic parameters E and ν are relatively uncorrelated, the three plastic parameters σ_0 , h and m are strongly correlated as are the three anisotropy parameters. In general, correlation of parameters does not lead to identification issues as long as there are no double eigenvalues or near zero eigenvalues in the set of calibrated parameters. However, correlation does complicate the discussion revolving about uncertainty, since the uncertainty of one parameter is influenced by other parameters.

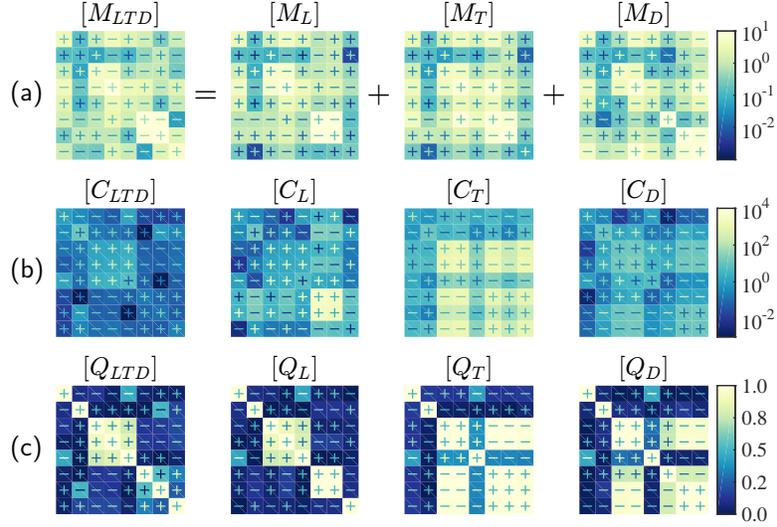


Figure 11: (a) Total Hessian $[M_{LTD}]$, which is the sum of the per-experiment Hessians ($[M_L]$, $[M_T]$ and $[M_D]$). (b) Inverse of each matrix shown above, which are approximations of the covariance matrices. (c) Correlation matrices, which are obtained by normalizing covariance matrices. The colors represent the absolute values of the components, while the signs are indicated with “+” and “-”. The constitutive parameter order is $E, \nu, \sigma_0, h, m, R_{22}, R_{33}, R_{12}$.

3.3. Identified Parameters

The proposed multi-experiment identification method iteratively minimizes the gap between the simulated and measured data starting from an initial guess for the unknown parameters. The case at hand is chosen since it can be well
 395 identified using more classical methods as discussed in Section 2.8. Consequently, the proposed method should return parameters that are similar to the classical parameter set. Table 1 shows the two sets of parameters. Most parameters are nearly identical between the two sets; most notably the elastic parameters and the anisotropy parameters. The plastic parameters σ_0 , h and
 400 m have changed significantly. However, even with seemingly large changes in these parameters, the stress-strain curves obtained from both parameter sets are remarkably similar (see Figure 5).

Table 1: Identified parameters using the classical method and the discussed multi-experiment framework

	E	ν	σ_0	h	m	R_{22}	R_{33}	R_{12}
	GPa	-	MPa	MPa	-	-	-	-
Classical	71.7	0.3	356	602	0.5	1.04	0.89	1.07
Multi-Exp.	70.6	0.31	330	446	0.38	1.03	0.9	1.07

To test the dependence on the initial guess, the multi-experiment identification routine was initialized with the ‘‘Classical’’ set and the same set multiplied
405 with various factors ranging from 0.5 to 1.5. The obtained parameters for each initial guess were within 10^{-4} of the parameters shown in Table 1.

3.4. Residuals

The gap between the simulated and measured observables is an integral part of the discussed identification method. The classical identification methods
410 typically depend on indirect residuals, such as the gap between the modeled and measured stress-strain curves. However, the parameters obtained with any identification method can be used to simulate the experiment to construct the full-field residuals as available in full-field identification methods.

The full-field residuals are part of the cost functions discussed in Section 2.3 and are the difference between the simulated and measured quantities

$$R_{Ikte} = f_{ke} - \tilde{g}_{kte}, \quad \hat{R}_I = \text{rms}(R_I), \quad (48)$$

$$R_{Fte} = F_{te}^{\text{EXP}} - F_{te}, \quad \hat{R}_F = \text{rms}(R_F), \quad (49)$$

$$\vec{R}_{Ukte} = \sum_i^{N_a} \vec{\varphi}_{ike} (a_{ite}^{\text{DIC}} - a_{ite}), \quad \hat{R}_U = \text{rms}(|\vec{R}_U|). \quad (50)$$

In this paper three residuals are discussed, where the image and force residuals
415 (*i.e.*, R_I and R_F) are part of the identification algorithm. The third (displacement) residual R_U is optional and available due to the application of global DIC with the same mesh as is used for the identification. Note that the indices k , t , and e indicate pixels, time steps and experiments respectively. The number of

pixels N_{ke} and time steps N_{te} depend on the experiment, the number of experi-
 420 ments N_e is three. Consequently R_I and \bar{R}_U are 3D fields (space and time) per
 experiment, and the former and latter are respectively scalar and vector fields.
 The force residual R_F is 1D in time per experiment.

The cost functions (Equations (21), (23) and (29)) are defined as extensive
 quantities so that they can be aggregated by summation. Consequently, their
 expectation is equal to the number of observables in the cost function. To
 facilitate the discussion on the quality of the identification intensive versions of
 three cost functions are defined here as

$$\begin{aligned}\hat{\eta}_I^2 &= \frac{\sum_e^{N_e} \eta_{Ie}^2}{\sum_e N_{te} N_{ke}} \\ \hat{\eta}_F^2 &= \frac{\sum_e^{N_e} \eta_{Fe}^2}{\sum_e N_{te}} \\ \hat{\eta}^2 &= \frac{\sum_e^{N_e} \eta_e^2}{\sum_e N_{te} N_{ke} + N_{te}}\end{aligned}\tag{51}$$

The first two are not used for identification purposes, but they are useful as
 indicators of the agreement of the individual data sets with the parameter set
 425 obtained with the identification set. All three have expectations of unity under
 the condition that the assumptions of measurement noise and model error are
 valid.

Table 2 shows the displacement \bar{R}_U , image \bar{R}_I and force \bar{R}_F residuals for the
 two parameter sets given in Table 1 as well as the image residual obtained using
 430 non-integrated DIC. The displacement and image residuals are only marginally
 larger than their respective uncertainties for both the ‘‘Classical’’ and the ‘‘Multi-
 Exp.’’ parameter sets. However, the force residual is considerably larger than its
 expected uncertainty for both identification cases. The same is expressed in the
 intensive cost function values $\hat{\eta}$, which are expected to be close to unity for the
 435 image part of the cost function while being far above unity for the force contri-
 bution. However, both identification methods return similar residuals with the
 multi-experiment identification method residuals being lower for all quantities.

Table 2: Residuals remaining after convergence for both identification cases, $\hat{\eta}$ are the intensive cost function residuals (Equation (51))

	\hat{R}_U [px]	\hat{R}_I [%]	\hat{R}_F [N]	$\hat{\eta}_I$	$\hat{\eta}_F$	$\hat{\eta}$
Classical	0.24	3.13	380	0.76	76	0.76
Multi-Exp.	0.22	2.78	336	0.67	67	0.68
DIC	-	1.80	-	-	-	-

An average gap between the simulated and measured forces of 336 N compared to a peak force of 10 kN may be considered both reasonable or unreasonable depending on the understanding of the experimental setup. The fact is that there are significant forces present in the measured data that are not accounted for by the model. These differences can have a number of sources, not limited to constitutive model error, geometrical model error, synchronization errors and measurement artifacts. The model error is present in both identification methods. However, the advantage of the proposed method is that the model error is now clearly visualized to allow the user of the method to judge whether this error is tolerable or not. Additionally, the signature of the model error can be analyzed to propose model enrichments that can account for the missing kinematics [20].

The $\hat{\eta}$ values reported in Table 2 are unexpectedly smaller than unity. First let us note that the image intensive cost function $\hat{\eta}_I$ is dominant in the global cost function, essentially because of the very large number of pixels as compared to the number of force measurements. Even, if $\hat{\eta}_F$ is very large as a result of a model error, this does not increase much the (intensive) global residual. The most likely explanation for $\hat{\eta}_I$ being smaller than 1 is that the estimate for the image uncertainty γ_I was high. The 20 images that were used to estimate image uncertainty may have been contaminated by a small but non-zero load. A deeper analysis of the uncertainty estimate may resolve this. However, it is considered out of the scope of this paper. A possible consequence is that the current identification method slightly overemphasized the force data.

Figure 12 shows the residuals for the parameter set obtained using the clas-

sical identification method. It presents the residuals as three columns, one per experiment, with seven sub-figures each. The first two rows are the displacement residual between the simulated displacements and those measured with non-integrated DIC (see Equation (48)) for one time and one space cross-section. The third row reports the image residual (Equation (50)) and the fourth and final row shows the residual force (Equation (49)) as a function of time.

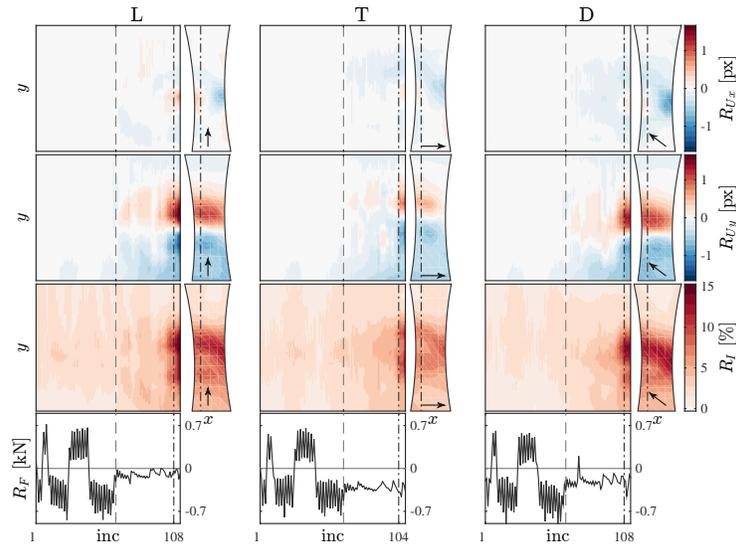


Figure 12: Full-field residuals in terms of displacement (first two rows), gray levels (3rd row) and force (4th row) for all three experiments when using the “Classical” parameter set as shown in Table 1.

The vertical dashed lines in the figure indicate the transition from the elastic part of the experiment to the plastic part roughly in the center of the time axis (see also Figure 3). The displacement and image residuals show a relatively low residual before this instant of time, and then gradually increase as plasticity is mobilized in the sample. This behavior is similar for all three experiments. However, the residuals for the T-sample are generally lower. For the force residual (4th row) the inverse is shown, with high residuals in the elastic part and lower residuals after the onset of plasticity. Additionally, the force residual is lowest for the L-sample, which is expected because the elastoplastic parameters

(*i.e.*, E , ν , σ_0 , h and m) were all obtained from the L-experiment data only. The high residuals in force in the elastic part correlate with the load-unloading cycles in the elastic regime and clearly indicate a phenomenon present in the
 480 experiment that is not modeled. This type of model error is common in identification, and the consequence of using real data instead of synthetic data. No further investigation of the source of this model error will be carried herein as it lies outside of the scope of the paper.

Similar to Figure 12, Figure 13 shows the full-field residuals, now for the
 485 parameter set calibrated with the discussed identification method. As was already known from the global residuals (Table 2), the multi-experiment residuals are generally lower for the multi-experiment identification as compared to the classical method. This is expected, since the multi-experiment framework is targeting these residuals directly in its cost function. Additionally, the residuals
 490 are more homogeneously distributed among the three experiments.

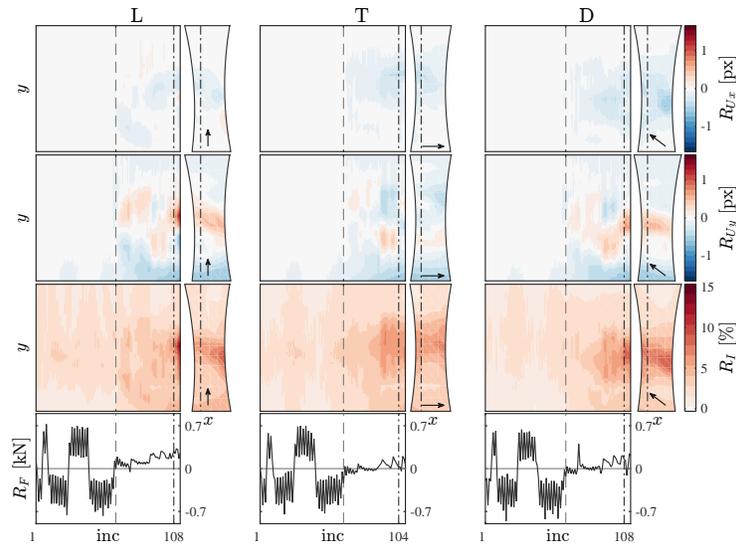


Figure 13: Full-field residuals in terms of displacement (first two rows), gray levels (3rd row) and force (4th row) for all three experiments when using the multi-experiment parameter set as shown in Table 1.

4. Conclusions

An identification method is proposed that aggregates the experimental data from multiple sources of multiple experiments into an identification method to calibrate a single set of constitutive parameters. The method is based on
495 Bayesian principles, thereby automatically accounting for the uncertainty of each data source to find their optimal weighting. An extractor is constructed to identify the set of unknown parameters considering all data that can reliably be represented by the chosen numerical model. The method provides an identification framework that naturally encompasses and extends on previously published
500 global DIC and Integrated-DIC frameworks. However, the proposed method is not limited to full-field data and works equally well with other measurement types, such as, load-cells, strain gages.

As a proof of concept, the method is applied to the identification of an 8 parameter isotropic elastoplastic model with a planar anisotropic yield surface.
505 The parameters calibrated with the proposed method are very similar to those using more classical methods. Moreover, the residual gap between the measured and modeled quantities is significantly reduced using the proposed method. The test case is chosen to illustrate this comparison, but it is emphasized that the proposed method is applicable to a much wider range of constitutive models
510 and inhomogeneous experiments.

One of the challenges in the identification of complex material models is dealing with non-uniqueness. Often there are multiple solutions that minimize the identification cost function. For cases that are not interested in the parameter values directly, picking any of the parameter sets will suffice, provided the
515 application of the calibrated model remains close to the identification regime. However, often the purpose of identification is to study the constitutive model directly. In those cases the parameter values are important. A way to discriminate between the solution sets is to enrich the identification with an additional experiment, designed specifically to trigger mechanisms previously not or not
520 sufficiently present. Within this paper methods are discussed to aid in de-

signing these experiments by means of analyzing the sensitivity of promising experiments.

The proposed method is based on the concept of transporting uncertainty along the experimental chain up to parameter identification. Consequently, it
525 provides uncertainty assessments of the calibrated parameters. It is based on a Gaussian model of the acquisition noise. Although a reasonable assumption for acquisition noise, such an assessment does not include any non-stochastic effects. The major critical point not yet addressed is the impact of model error. Most, if not all, parameter identification methods leave a residual gap between
530 the measured and modeled quantities. Often, the amplitude of this gap is more significant than the uncertainty due to random noise. Comparing the residual gap of the proposed method to more classical and trusted identification methods reveals that this issue is only mildly improved with the proposed method. The key difference is that the proposed method shows the residual gap very visually
535 and thus provides means of assessing whether the gap is acceptable. For the example discussed herein, the most likely cause of the residual gap is a limited model (be it from the constitutive law, experimental control, or measurement acquisition) that is not capable of describing the experiment correctly at the transition from loading to unloading and for higher plastic strains. One way
540 to address this issue is the successive enrichment of the model with more parameters [20]. One should take care of the risk of degeneracies. However, the proposed method provides various tools for progressing along this challenging path.

Acknowledgement

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