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1 2	Urban hydrologic trend analysis based on rainfall and runoff data analysis and conceptual model calibration
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11	
12	Keywords
13	Urban hydrology; urbanization; conceptual rainfall-runoff model; trend analysis; Mann-
14	Kendall test
15	Abstract
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18 19 20 21 22 23 24 25 26	therefore important to assess any hydrologic trends in urban catchments for stormwater management and planning. This study addresses urban hydrological trend analysis by examining trends in variables that characterize hydrological processes. The original and modified Mann-Kendall methods are applied to trend detection in two French catchments, i.e., Chassieu and La Lechere, based on approximately one decade of data from local monitoring programs. In both catchments, no trend is found in the major hydrological process driver (i.e., rainfall variables), whereas increasing trends are detected in runoff flow rates. As a

27 rainfall-runoff model parameters, which are identified via model calibration with an event

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based approach, are examined. Trend detection results indicate that there is no trend in the
time of concentration in Chassieu, whereas a decreasing trend is present in La Lechere, which,
however, needs to be validated with additional data. Sensitivity analysis indicates that the
original Mann-Kendall method is not sensitive to a few noisy values in the data series.

32 **1. Introduction**

41

Urban stormwater is a major cause of urban flooding and water pollution, which can lead to serious economic and social consequences. Currently, it is widely recognized that effects that are closely related to human activities, such as climate change and urbanization can result in significant alteration of stormwater quantity and quality (e.g., Astaraie-Imani et al., 2012). The design and operation of urban storm water projects that aim to reduce the adverse impact of stormwater should take possible changes in urban hydrology patterns into account. It is therefore important to assess any hydrologic trend (if one exists) in urban catchments.

40 Considerable attention has been paid to trend analysis in research areas in climatology,

hydrology and water quality in recent years. Examples of trend detection applications in water

resources studies include trend studies of precipitation (e.g., Xu et al., 2003; Partal and Kahya, 42 43 2006; Gocic and Trajkovic, 2013), streamflow (e.g., Douglas et al., 2000; Zhang et al., 2001; Burn and Elnur, 2002; Yue et al., 2002; Aziz and Burn, 2006), and river and drainage water 44 quality (Hirsch et al., 1982; Awadallah et al., 2011; Sun et al., 2015). A number of trend 45 detection methods including parametric and non-parametric tests have been applied (Hess et 46 al. 2001). A non-parametric test is generally more suitable for non-normally distributed and 47 48 censored data, which are frequently encountered in water resources data (Yue et al., 2002). The Mann-Kendall method (Mann, 1945; Kendall, 1955) is one of the most commonly used 49 non-parametric trend detection tests (e.g., Omar et al., 2006). Because the Mann-Kendall 50 method generally requires serial independence, some previous studies have modified the 51

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52 Mann-Kendall Method to apply it to data presenting seasonality or serial correlations (e.g.,

53 Hirsch and Slack, 1984; Hamed and Rao, 1998; Yue et al., 2002).

A rainfall-runoff model that transforms the meteorological forcing (rainfall) into the 54 hydrological response of a catchment (runoff) is an important tool for theoretical and applied 55 56 research in hydrology. A simple conceptual rainfall-runoff model can sometimes serve as a powerful tool for aiding in understanding local hydrological processes. For instance, the 57 hydrological response can be linked to landscape attributes by deriving the relationship 58 59 between hydrological model parameters and landscape attributes (e.g., Post and Jakeman, 1999). Statistically significant correlations between some parameters of a conceptual daily 60 rainfall-runoff model and the catchment physical and climatic characteristics were found 61 (Chiew et al., 2002). Hence, the interpretability of the rainfall-runoff model parameters can 62 possibly shed some light on the hydrologic behaviour of a catchment (Maneta et al., 2007). 63

Given the importance of assessing hydrologic trend in urban catchments, the main objective 64 of this paper is to present a methodology that identifies and quantifies hydrologic trend. The 65 hydrological trend analysis is addressed using a dual approach, i.e., by analysing both rainfall 66 and runoff data and model parameters identified from the calibration of a conceptual 67 hydrological model. The rationale of examining possible hydrological trends via calibrated 68 hydrological model parameters is that if there is any change in the local hydrologic process, 69 70 non-stationarity will probably be present in model parameters characterizing the temporal 71 hydrological process. In addition to direct analysis of rainfall and runoff data, a conceptual 72 urban hydrological model helps investigate more characteristics of local hydrology (e.g., initial precipitation losses and the time of concentration), which cannot be measured directly. 73 74 However, it should be noted that the evolution of hydrological processes is an aggregate result of changes of many distributed factors and processes on smaller scales (than the catchment 75 scale) in local hydrology (e.g., changes in distributed land uses and hydrological properties), 76

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using a conceptual rainfall-runoff model, this study only addresses the evolution of model 77 78 parameters that are highly lumped on the catchment scale which reveal the general hydrological responses of urban/peri-urban catchments. The evolution of the hydrological 79 trend provides the most direct evidence of changes in local hydrology and the quantification 80 of the trends in relevant parameters is useful in urban stormwater management. The evolution 81 82 of processes, factors and properties on smaller scales than the catchment scale is beyond the 83 discussion of this study. This study also provides guidance for analysing long-term hydrological trends by applying temporal calibration of a conceptual hydrological model. The 84 effectiveness of such an approach is demonstrated via applications of the methodology to two 85 French catchments. 86

87 **2.** Case studies and data

Two urban catchments in the suburbs of Lyon, France, i.e., the Chassieu catchment and the La
Lechere catchment, are studied in this paper. Both catchments are monitoring sites under the
OTHU program (www.othu.org), which has been operating for over a decade to improve our
knowledge on urban water system management by acquiring reliable data of both wet and dry
weather flows and their impacts on the receiving environment.

The Chassieu catchment is located in the east of the Greater Lyon area. It covers an industrial 93 area of 185 ha with an imperviousness coefficient of approximately 0.72. The catchment is 94 drained by a separate stormwater sewer system, which also receives dry weather flows from 95 cooling of industrial processes (that can be assumed clean). The pervious area is not 96 97 connected to the sewer system. Rainfall in Chassieu was measured by a tipping-bucket rain gauge installed in the catchment, and a six-minute rainfall time series is available. The 98 catchment runoff flow rate was computed from water depth measurements in the 1.6 m 99 100 circular concrete pipe at the outlet of the catchment using the Manning equation (calibrated and validated using measured water depth and velocity data in the pipe) with a two-minute 101

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interval from 2004 to 2011. The rainfall and runoff time series suffer from 7.4% and 13.1%missing data, respectively.

La Lechere catchment is located to the west of Lyon. The catchment covers a maximum area 104 of 410ha when all of the combined sewer overflows (CSOs) are activated. The land in the 105 catchment is mainly composed of urban areas (53%), agricultural fields (45%) and forests 106 107 (3%) (Braud et al., 2013). The urban areas are drained by combined sewer networks, with several CSOs connected to the Chaudanne River. The contributing area of stormwater to the 108 combined sewer system is approximately 120 ha (stormwater from other areas are not 109 connected to the sewers). In addition to the CSOs, the Chaudanne River also receives natural 110 flows from rural areas. More details about the catchment can be found in Jankowfsky et al. 111 (2014). Rainfall was measured by a tipping-bucket rain gauge located in the catchment. A 112 one-minute rainfall time series is available with approximately 13.7% missing data. The 113 runoff flow in the Chaudanne River, mainly composed of CSOs from urban areas and natural 114 streamflow from rural areas, was computed from water depth measurements in the calibrated 115 Parshall flume at a gauge station. The runoff flow data are registered with varied time steps, 116 typically from two minutes to one hour. A two-minute time series was created using a linear 117 interpolation method. 2.8% of the runoff flow data is missing. Rainfall and flow data are both 118 available from June 2005 to December 2014. 119

120 **3. Methodology**

121 3.1 Mann-Kendal test and modified Mann-Kendall test

122 The Mann-Kendall method is a rank-based nonparametric trend detection test extensively

- applied in climatology and hydrology. The null hypothesis of the Mann-Kendall test H_0 states
- that the data are a sample of *n* independent and identically distributed random variables,
- whereas the alternative hypothesis H_1 is that x_k and x_j are not from identical distributions (k, j
- 126 $\leq n$ and $k \neq j$). The test statistic *S* is defined as:

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$$S = \sum_{k=1}^{n-1} \sum_{j=k+1}^{n} \operatorname{sgn}(x_j - x_k)$$
(1)

127	where sgn(θ) is the sign function that equals -1, 0 and 1 when θ is below, equal to and above 0,
128	respectively. Under the null hypothesis, S is asymptotically normally distributed with the
129	mean of 0 and a constant variance, which is a function of the number of data in a tested data
130	series (e.g., Hess et al., 2001; Gocic and Trajkovic, 2013). A P-value, which presents the
131	probability of obtaining samples as extreme as the observed ones, can be computed from
132	given S and its variance. The null hypothesis is accepted with a higher P-value than a
133	predefined significance level. Otherwise, the null hypothesis is rejected, suggesting that a
134	trend is detected. Two levels of significance, i.e., 5% and 1%, are used in this study.
135	The Mann-Kendall method usually indicates a higher false positive outcome for data with
136	positive autocorrelation, and it is thus no more effective for auto-correlated time series (e.g.,
137	Yue et al., 2002). A modified Mann-Kendall trend test (Hamed and Rao, 1998) is used when
138	the autocorrelation effect in a data series is significant. In the modified Mann-Kendall method,
139	the variance of <i>S</i> is calculated using an empirical formula with a multiplicative coefficient,
140	which is a function of the autocorrelation coefficient (see Hamed and Rao, 1998 for the
141	formula). As a trend generally leads to positive autocorrelation, a data series is firstly
142	detrended with a linear trend estimated from linear regression. The detrended data series is
143	then tested for its autocorrelation effect (Yue et al., 2002). If the autocorrelation in a data
144	series is insignificant at the 5% significance level, the Mann-Kendall test is applied.
145	Otherwise, the modified Mann-Kendall test is performed.
143	oulerwise, ale mounted mann Rendan test is performed.

146 The slope of a trend (if one exists) can be estimated by a non-parametric index (Sen, 1968)147 based on the assumption of a linear trend:

$$\beta = Median\left(\left(x_j - x_i \right) / (j - i) \right), \quad i < j$$
⁽²⁾

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148 The value of β represents the changing value per event if an event-based data series is 149 considered. An annual slope of a trend is computed by multiplying the average number of 150 events in one year to this value.

151 **3.2 Rainfall and runoff event identification**

Storm events are identified from continuous rainfall time series with a dry period over four 152 hours between two events, which is empirically identified (Métadier and Bertrand-Krajewski, 153 2012; Sun et al., 2015). This study concerns only urban rapid flow. For most events, the rapid 154 response of runoff to rainfall from urban areas ends in four hours in both catchments 155 according to visual inspection. The time series of runoff is thus identified covering the period 156 157 of a corresponding rainfall and four hours more after the rainfall event ends. In La Lechere, the measured runoff flows in the Chaudanne River are partly from upstream rural areas, which 158 respond much slower than urban areas. A baseflow, considered as a constant flow with the 159 flow rate equal to the minimum flow rate measured during an event, is subtracted from the 160 measured runoff time series to identify the urban rapid flow part. The baseflow in Chassieu 161 162 mainly comes from industrial wastewater. For most events, the assumption of a constant baseflow during an event for several hours is reasonable. 163

Only significant events are considered for trend analysis of urban hydrological processes, as parameters characterizing small events are easily affected by influential factors such as initial rainfall loss, baseflow and measurement uncertainty. Significant events in Chassieu are defined with a total rainfall depth over 1 mm and duration over 30 minutes. In La Lechere, a higher threshold with 2 mm rainfall depth and over 30-minute duration is adopted because the CSOs in the combined sewers are often not activated during smaller events. In addition, a significant event in both catchments requires a mean urban runoff flow over 5 L/s.

171 **3.3** Simple parameters characterizing hydrological processes

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Two simple parameters that can be directly calculated from rainfall and runoff data without 172 modelling, i.e., the runoff coefficient and lag time, are derived for each event. The runoff 173 coefficient is the ratio of runoff water volume to rainfall volume, arbitrarily representing the 174 proportion of rainfall entering the drainage system, which is generally a function of land 175 covers and imperviousness of the area. Trend analysis of the runoff coefficient of rapid urban 176 runoff can possibly reveal the evolution in local urbanization. The lag time measures the 177 response time of a catchment to a rainfall event, which is closely related to the topography, 178 geology and land use within a catchment. In this study, the lag time of one event is evaluated 179 as the time difference between the mass centres of the hydrograph and the hydrograph. 180 181 Evolution of the lag time can possibly reveal changes in mechanisms and processes governing the runoff generation and transportation in a catchment. For instance, the lag time in an area 182 drained with sewer pipes is much shorter than that in a naturally drained area of a comparable 183 184 size and slope.

185 **3.4 Conceptual urban rainfall-runoff model**

Noting that the simple parameters defined in the above section are directly calculated from rainfall and runoff data, ignoring the non-linear complex real rainfall-runoff mechanisms may create bias in characterising the local process. Therefore, a conceptual urban rainfall-runoff model is employed in this study to aid in identifying more relevant variables based on a more comprehensive description of the hydrological process.

The conceptual urban rainfall-runoff model consists of a simple rainfall loss model and a routing model of two cascaded linear reservoirs. The evaporation and evapotranspiration are negligible at an event scale, and are thus not considered in the model structure using an eventbased approach. This model has been successfully applied in urban hydrology for small impervious catchments drained by artificial sewer systems (e.g., Sun and Bertrand-Krajewski, 2013; Leonhardt et al., 2014). The rainfall loss model calculates net rainfall *I*_{net} by subtracting

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an initial loss L_{ini} (mm) and a proportional loss P_{cons} (-) during a rainfall event from gross rainfall *I*.

$$I_{\rm net}(t) = \begin{cases} 0 & \text{if } \int_{t=0}^{t} I dt \leq L_{\rm ini} \\ I(t) (1 - P_{\rm cons}) & \text{if } \int_{t=0}^{t} I dt > L_{\rm ini} \end{cases}$$
(1)

199 Net rainfall is shifted with a time shift T_{shift} and is converted to inflow by multiplying it by the 200 effective catchment area *A*:

$$Q_{\rm in}(t) = I_{\rm net}(t - T_{\rm shift}) \times A$$
(2)

In the runoff routing model, Q_{in} is routed through two cascaded linear reservoirs with the same reservoir constants (*K*). A linear reservoir assumes that the outflow Q_{out} is linearly related to the storage volume. The analytical solution of a linear reservoir model over the time interval $[t - \Delta t, t]$ is

$$Q_{\rm out}(t) = \exp(-\frac{\Delta t}{K})Q_{\rm out}(t - \Delta t) + \left[1 - \exp(-\frac{\Delta t}{K})\right]Q_{\rm in}(t)$$
(3)

The outflow from the first linear reservoir is then routed to the second reservoir as the inflow and the outflow from the second reservoir is the output of the conceptual model. The conceptual urban hydrological model contains four parameters, i.e., rainfall initial loss L_{ini} , rainfall constant proportional loss P_{cons} , time shift of inflow T_{shift} and reservoir constant of the two reservoirs *K*.

210 **3.5 Event-based conceptual hydrological model calibration**

An event-based approach for hydrological process modelling is computationally efficient
when runoff is restricted to a short period after a storm event (Maneta *et al.*, 2007). This is

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often the case in urban catchments equipped with storm sewer systems. A conceptual 213 214 hydrological model is often too simple to cover all conditions of the catchment and it has to adjust itself (by adjusting model parameters) to represent different conditions. As a 215 216 consequence, model parameters are usually temporally different for varied rainfall-runoff events. The dynamics of model parameters identified from event-based model calibration 217 218 using measured rainfall and runoff data thus possibly represent temporal catchment 219 characteristics and conditions. For instance, in our urban rainfall-runoff conceptual model, the 220 initial loss roughly indicates the antecedent weather condition of an event; the proportional loss is probably an indicator of imperviousness; the time shift and reservoir constant are 221 222 related to the time of the catchment responding to rainfall. These model parameters can further be used to study the evolution of long-term catchment properties. 223

The model is calibrated using the DREAM algorithm (Vrugt et al., 2008), which searches for optimal parameters based on the Monte Carlo Markov Chain method. The effectiveness of DREAM in calibrating hydrological models has been demonstrated by many studies in the literature (e.g., Schoups *et al.*, 2010).

228 **3.6** Time of concentration estimated from conceptual rainfall-runoff model parameters

The time of concentration, which is usually defined as the time of water flowing from the 229 point with the longest temporal flow path within a catchment to the catchment outlet, also 230 measures the response time of a catchment to a rain event. The time of concentration is an 231 important concept in hydrology because many practical designs and operation strategies rely 232 233 on the prediction of the catchment response time. The most common method to estimate the time of concentration is via the identification of the flow path and the time of concentration is 234 235 the travel time of flows through the flow path. In this study, the time of concentration is 236 computed as the length of a unit hydrograph, which is the response of a watershed (in terms of

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runoff volume and timing) to the input of a unit of rainfall. Once the urban hydrological 237 238 model is calibrated for a specific event, its unit hydrograph can be determined as a function of two calibrated model parameters (i.e., the time shift and reservoir constant). More specifically, 239 240 a unit depth of net rainfall is input into the urban rainfall-runoff model (Eqs. (2) and (3)) with the two calibrated model parameters, and the model output is its corresponding unit 241 hydrograph. Because a unit hydrograph from the linear reservoir model has a very long 242 243 recession limb, the time of concentration is estimated as the duration when 95% water volume reaches the watershed outlet (a higher percentage of volume generally does not change the 244 relative magnitudes of the time of concentration from different hydrographs). 245

246 **3.7 Uncertainty in trend analysis due to noisy events**

Due to various reasons (e.g., measurement uncertainty, data errors and uncertainty in model 247 248 calibration), there might be erroneous values (from a noisy event) in data series. To study the influence of the erroneous values on the trend detection results, uncertainty in trend detection 249 results due to erroneous values in data series is examined. One synthetic noisy value is 250 introduced into a data series at different positions. A noisy value is assumed to be extremely 251 big, small or median. An extremely big (small) value is even bigger (smaller) than the 252 253 maximum (minimum) value in the data series. The sensitivity of the trend detection results to the magnitude and position of noisy values is then investigated by comparing the results of the 254 data series without and with an erroneous value. 255

256 4 Results and discussion

257 4.1 Rainfall-Runoff event characteristics

A total number of 692 significant events in Chassieu and 584 significant events in La Lechere
have been identified with complete one-minute or two-minute rainfall time series. A total
number of 584 and 442 runoff events are identified with complete runoff data with mean flow

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over 5 L/s in Chassieu and La Lechere, respectively. The conceptual urban hydrological 261 model described above is calibrated to fit the rainfall and runoff data for each specific event 262 (see Section 4.3 for more details on the model calibration results). Most events (477 events in 263 Chassieu and 398 events in La Lechere) can be satisfactorily described by the conceptual 264 urban rainfall-runoff model with a Nash-Sutcliffe (NS) model efficiency coefficient over 0.7. 265 However, for a small number of events, the conceptual urban rainfall-runoff model produces 266 outputs with significant discrepancies from the measured data. This is likely due to errors 267 either in rainfall or runoff measurements (e.g., catchment areal rainfall not captured by point 268 measurement, Leonhardt et al., 2014) or the simple assumption of a constant baseflow. Only 269 270 events that can be satisfactorily described by the conceptual urban rainfall-runoff model are considered for the following trend analysis. Fig. 1 shows the rainfall depth and duration of all 271 rainfall events and selected events that can be satisfactorily described by the conceptual urban 272 273 rainfall-runoff model.

274 Table 1 summarizes some statistical characteristics of the rainfall and runoff variables of the selected events in the two catchments. The runoff-based variables are considered after 275 subtracting the baseflow, because this study is only concerned with the fast response of the 276 urban areas to rainfall. Big rainfall events were observed with a maximum rainfall depth of 277 134.6 mm in Chassieu and 91.4 mm in La Lechere. The rainfall and runoff variables were 278 generally distributed over wide ranges, with relative standard deviations typically over 1, 279 indicating positively skewed distributions of the variables, which are consistent with previous 280 findings (Brezonik and Stadelmann, 2002). 281

282 4.2 Trend analysis of simple variables without modelling

Variables of the selected events based on simple data analysis, i.e., baseflow, runoff
coefficient and lag time, are shown in Fig. 2 for the two catchments. Their statistical

characteristics are also summarized in Table 1.

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In Chassieu, the baseflow values are mostly under 25 L/s with a median value of 2.0 L/s. Only 286 a few values in the latter years are over 25 L/s. Values of the runoff coefficient are between 287 0.03 and 0.81, with a median value of 0.30. The wide range of the runoff coefficient can be 288 explained by influential factors such as initial loss and dry weather flows, which are event-289 specific. The lag time also covers a large range, mostly between 20 and 200 minutes with a 290 median of 77 minutes and with two negative values. The negative lag time of one event is 291 292 likely due to errors in rainfall and runoff measurements and the assumption of a constant baseflow. The lag time varies depending on the profiles of the hyetograph and hydrograph, 293 294 due to the non-linear relationship between rainfall and runoff and possibly varied temporal hydrological regimes. 295

In La Lechere, the median baseflow is 7.2 L/s and most values are in the interval of [0, 100].
One event with an extreme baseflow (over 300 L/s) occurred at the end of 2008, which can be
explained by a rainfall event of 73.4 mm ending only 15 hours before it, discharging high
natural flow from rural areas. The runoff coefficient is distributed across a range of [0.01, 0.3],
with a median value of 0.04. The generally low runoff coefficient (in comparison with that of
Chassieu) is found because only CSOs contribute to the considered runoff in La Lechere. The
median lag time is 110 minutes, with the range in [-50, 405] minutes.

Table 2 lists trend detection results for different rainfall and runoff variables. The original or modified Mann-Kendall test is applied according to the significance of the autocorrelation test. For both catchments, no trend is found for most rainfall-based variables, including depth, duration and mean intensity. A relatively low *P*-value of 3.2% is obtained for the rainfall duration in La Lechere, which is significant at the 5% significance level, but is still above the less strict 1% significance level. In addition, the trend test is performed on many other rainfall variables in addition to those listed in Table 2 (e.g., rainfall depth and intensity in a specific

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310	duration). The results indicate no trend in any rainfall-based variables. There is likely no trend
311	in the rainfall variables (the driver of the hydrological process) for the study period.
312	For the urban runoff-based variables, data series of the runoff volume and mean runoff flow
313	are investigated. According to the original or modified Mann-Kendall test, all variables
314	present an increasing trend with P-values on the order of 1% or lower. The mean urban runoff
315	flow rate increases by 3.1 L/s in Chassieu and 0.7 L/s in La Lechere on average per year.
316	As a result of the relatively stable rainfall and increasing runoff, an increasing trend is
317	detected in the runoff coefficient in both catchments with very low P-values (on the order of
318	10^{-11} in Chassieu and 10^{-3} in La Lechere). The overall increasing rates of the runoff
319	coefficient are evaluated as 0.012 and 0.002 per year in Chassieu and La Lechere, respectively.
320	The increasing runoff coefficient is likely due to growing imperviousness caused by
321	urbanization in both catchments. In Chassieu, a comparison of aerial views at the beginning
322	and end of the study period shows more buildings being constructed, which led to increasing
323	imperviousness during the study period (Sun et al., 2015). However, it is worth noting that
324	urbanization can only be regarded indicative to runoff coefficient changes, because
325	urbanization does not always lead to higher runoff volumes, particularly with the
326	implementation of low urban development (LID) techniques such as porous pavement,
327	infiltration trenches and green roofs. In contrast, the lag time seems not to present a trend,
328	indicating no significant change in the travel time of flows in the catchments.
329	An increasing trend in the baseflow is confirmed in Chassieu with a <i>P</i> -value of 2.3%,
330	indicating more industrial wastewater draining into the system, whereas no trend is detected
331	in the baseflow in La Lechere, implying a relatively stable rural flow in this area.
332	4.3 Conceptual urban rainfall-runoff model based analysis

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An event-based calibration is implemented for all of the available rainfall-runoff events in the 333 two catchments based on measured rainfall and runoff data. The optimal model parameters for 334 each event are identified using the DREAM algorithm based on 10^4 model evaluations. The 335 search ranges of the model parameters, which are determined based on catchment properties, 336 are given in Table 3. Only events that are satisfactorily described by the conceptual rainfall-337 runoff model with an NS value over 0.7 are considered. Table 3 also lists the summary 338 statistics of the optimal model parameters. Fig. 3 shows the optimal model parameters. The 339 parameter of initial loss is broadly distributed in the search range of [0-2mm] or [0-3mm] in 340 the two catchments, with several events reaching the limits. However, the limits are not 341 342 extended to reflect the physical reality in the catchments. The proportional loss is also eventdependent and distributed in a wide range with a median of 0.65 in Chassieu and 0.88 in La 343 Lechere. This parameter can be roughly linked to the runoff coefficient, with the absolute 344 345 values of the correlation coefficients of 0.73 in Chassiu and 0.83 in La Lechere. The reservoir constant and time shift reveal the response time of the catchments to rainfall events. These 346 347 two parameters also show great variability.

The temporal variability of the model parameters implies that the lumped conceptual urban 348 rainfall-runoff model is too simple to cover all conditions encountered by all of the events in 349 the urban catchments, which is consistent with previous findings (Maneta et al., 2007). The 350 calibrated parameters reveal the temporal hydrological regime/conditions during one specific 351 event. However, the variability of optimal model parameters may also result from other 352 reasons, leading to bias in some model parameter estimates in the lumped conceptual urban 353 354 rainfall-runoff model. For instance, errors in areal rainfall (represented by point measurements) and runoff measurements lead to biased calibration results, as in calibration, the model 355 parameters are adjusted to make the lumped model outputs match the measurements. 356 357 Additionally, the model parameters are correlated (e.g., Sun and Bertrand-Krajewski, 2013).

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The initial loss and proportional loss are negatively correlated; the reservoir constant and the time shift are also negatively correlated. Simultaneously adjusting correlated parameters gives equivalent model performances.

The time of concentration for each event is evaluated from the calibrated reservoir constant and time shift (Fig. 4). The median time of concentration in Chassieu is 92 minutes and most values are in the range of [10-400]. The time of concentration in La Lechere covers a range of [24, 724] minutes and the median is 132 minutes.

Fig. 5 shows the scatter points of the lag time directly derived from the data and the time of concentration estimated from calibrated urban rainfall-runoff model parameters. These two quantities both indicating the response time of the urban catchments to rainfall events are highly correlated. However, the relation between these two variables also shows some randomness, resulting from the non-linear relation between rainfall and runoff and the different methods with which the two parameters are calculated.

Assuming that other factors listed above (e.g., errors in rainfall and runoff data, uncertainty in 371 calibration and parameter correlation) leading to temporal variation of model parameters are 372 373 random without a trend, trend analysis of model parameters can still provide useful information on hydrological regime evolution in the two catchments. Table 4 lists the trend 374 detection results of optimal model parameters along with the time of concentration. The initial 375 376 loss does not present a trend in both catchments, indicating stable initial conditions of rainfall events, which are linked to antecedent dry periods and precedent events. A decreasing trend is 377 378 found in the constant loss in both catchments with low P-values, which is consistent with the detected increasing trend in the runoff coefficient in the above analysis due to urban 379 development in the catchments. The constant loss is evaluated with a decreasing rate of 380 381 approximately 0.014 per year in Chassieu and 0.005 in La Lechere. These figures are close to

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the values of the increasing rates in the runoff coefficient estimated using the simple dataanalysis method (see Table 2).

The reservoir constant and lag time in Chassieu seem to not contain a trend according to the 384 Mann-Kendall method, suggesting that the routing function governing runoff generation does 385 not have a significant change in Chassieu. In contrast, the reservoir constant possibly presents 386 a decreasing tendency in La Lechere with a P-value under 5%. Consequently, the time of 387 concentration evaluated based on this parameter in La Lechere is also detected with a 388 declining trend of approximately 3.5 minutes per year. This is consistent with the common 389 sense notion that urbanization leads to a shorter time of concentration because an artificial 390 drainage system often transports stormwater much faster than a natural water course. In the 391 392 study period of ten years, a decrease of 35 minutes in the concentration time is expected, which is non-negligible in comparison with the median time of concentration of 132 minutes. 393 However, the trends in both variables are rejected with the stricter 1% significance level. 394 Further data are required in order to confirm the trend in these variables. 395

4.4 Sensitivity of noisy events to trend detection results

397 Due to the various reasons presented above (e.g., the lumped and simplified model representing the complex rainfall-runoff process, errors in rainfall and runoff measurements 398 and the correlation between model parameters in calibration), parameters characterizing local 399 hydrology obtained from data analysis and model calibration for some events may be 400 erroneous. For instance, the negative lag time and the reservoir constant of three events in La 401 402 Lechere close to 240 minutes (which is the search limit) are likely to be erroneous. This section investigates the influence of the presence of noisy events (with erroneous values) in 403 the data series on the trend detection results using the Mann-Kendall test. 404

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Four data series of optimal model parameters are used to study the sensitivity of trend
detection results to noisy values. The original data series are considered as the comparison
benchmarks, assuming that they do not contain any erroneous data. Fig. 6 shows the *P*-values
of different cases with one noisy value together with the *P*-value of the original data series. It
is clear that the influence of a noisy event depends both on the magnitude and the position of
the noisy event.

In Fig. 6 (a), the data series of the initial loss in Chassieu, which does not present a trend is studied. The original Mann-Kendall method is used because the autocorrelation effect in the data series is insignificant. An extreme noisy value introduced at the two sides of the data series leads to the most significant change in the *P*-value. The influence of a noisy value is negligible when it is in a middle position in the data series. A noisy median value does not significantly affect the *P*-value at any position in the data series. The *P*-value changes gradually as a noisy event moves in the data series.

In Fig. 6(b), the data series of the constant loss in Chassieu, which presents a decreasing trend as indicated by the original Mann-Kendall method, is studied. An extremely high noisy value located at the beginning of the data series leads to a lower *P*-Value, as expected. An extreme noisy value in the middle positions of the data series and a noisy median value at any position do not have significant impacts on the *P*-value.

In Fig. 6 (c) and (d), data series of the constant loss and time shift in La Lechere with
significant autocorrelation are studied using the modified Mann-Kendall method. Different
from those in Fig. 6 (a) and (b), the *P*-values do not change monotonously as the position of a
noisy value changes gradually in a data series, due to the influence of the autocorrelation
coefficients incorporated in the modified Mann-Kendall test. The relation between the *P*value and the position of an erroneous value is rather random. The most significant change in

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a *P*-value due to an extreme noisy value does not necessarily occur when it is located at the
two sides of the data series. In both cases, the introduction of one erroneous value does not
change the trend detection results.

Furthermore, the impact of the number of noisy values on the trend detection results using the 432 original Mann-Kendall method is also studied. The trend detection results from the original 433 Mann-Kendall method are most sensitive to extreme erroneous values at the two sides of the 434 data series. Therefore, the impact of the number of noisy values is studied here by examining 435 cases with extreme noisy values introduced in the beginning of data series. Fig. 7 shows the 436 results. In Fig. 7(a), for the data series of the initial loss (containing 477 events) in Chassieu, 437 which does not present a trend, a trend will only be detected with more than 9 extremely small 438 noisy events or 21 extremely large events introduced in the beginning of the data series. In Fig. 439 7(b), for the constant loss data series in Chassieu presenting a significantly decreasing trend 440 with a *P*-value of 2.7×10^{-9} %, the Mann-Kendall test only gives a different indication of trend 441 442 absence when over 32 extremely small noisy values are introduced.

Based on the above analysis, a different trend result generally requires a number of extreme erroneous values at one side of a data series. The Mann-Kendall method, which is a function of the ranks of the observations rather than their actual values, is not sensitive to a few noisy values. Therefore, in this study, the trend detection results using the original Mann-Kendall method are probably reliable regardless of a few possible noisy values in the data series. However, the sensitivity of the trend detection results from the modified Mann-Kendall method seems to be more complicated due to the altered autocorrelation coefficients.

450 **5. Conclusions**

It is important to assess any hydrological trend (if one exists) in urban catchments because thedesign and operation of urban stormwater projects aiming to reduce the adverse impact of

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stormwater should take into account possible changes in urban hydrology patterns. This paper 453 454 addresses hydrological trend analysis in urban catchments. Using rainfall and runoff data and a conceptual rainfall-runoff model, lumped parameters that aggregate distributed spatial 455 456 information in the urban/peri-urban catchments are analysed to reveal the global evolution of the hydrological responses (noting that distributed processes, factors and properties on smaller 457 scales in local hydrology are beyond the discussion). The evolution of aggregate hydrologic 458 459 relevant parameters is useful in indicating changes in local hydrology and needs to be considered in urban stormwater management. The original and modified Mann-Kendall 460 methods are applied for trend detection in data series in two catchments in France, i.e., 461 462 Chassieu and La Lechere, and there seems to be no difficulty of applying the methodology to other urban catchments. Based on the application results for approximately one decade of data, 463 the following conclusions can be drawn: 464 1. A trend is absent in the driving force (precipitation) of the rainfall-runoff processes in both 465 466 catchments in the suburbs of Lyon for the study period. 2. An increasing trend is found in the urban runoff variables. On average, the mean runoff 467 flow rate increases by 3.1 L/s in Chassieu and 0.7 L/s in La Lechere per year. 468 3. As a result of the relatively stable rainfall and increasing urban runoff volumes, the runoff 469 coefficient presents an increasing trend in both catchments, probably due to growing 470 imperviousness caused by urbanization, even in Chassieu, where the catchment is already 471 very densely urbanized, rather than climate change factors. However, it is worth noting that 472 nowadays urbanization does not automatically lead to more imperviousness and increasing 473 catchment outflow. The influence of urbanization development on the local hydrological 474

475 processes is much more complicated than simple imperviousness evolution because it is

476 highly dependent on the detailed changes and processes. For instance, in the LID context,

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when urban areas expand, water on the impervious surface can be connected to vegetated
areas or infiltration devices; some rain can be retained via green roofs. In addition, the
efficiency of an LID device may decline over time (e.g., when an infiltration system is
clogged), leading to a part of stormwater back to catchment outflow. Therefore, the evolution
of stormwater quantities is difficult to be precisely inferred only from urbanization
information. The methodology developed in this study provides the most direct evidence of
the elevated runoff coefficient in urban/peri-urban catchments.

484

485 4. Optimal parameters of a conceptual urban rainfall -runoff model obtained from calibration using an event-based approach are used to represent the temporal hydrological regime. 486 Though temporal variation of optimal model parameters is also affected by other factors such 487 488 as errors in rainfall and runoff measurements, uncertainty in calibration and parameter correlation, the variance due to these factors is assumed to be random. The trend analysis 489 results of these model parameters therefore reveal local hydrology evolution. No trend is 490 present in the initial loss. A decreasing trend is found in the constant proportional loss in both 491 catchments, which is consistent with the increasing runoff coefficient. The time of 492 493 concentration in Chassieu does not exhibit a trend, whereas it seems to decrease by 3.5 minutes per year in La Lechere. The trend in the time of concentration in La Lechere needs to 494 be confirmed with further data. 495

5. The sensitivity analysis indicates that the original Mann-Kendall method is not sensitive to a few noisy values in the data series. Therefore, the trend detection results from the original Mann-Kendall method are reliable, regardless of a few possible erroneous data. However, the relation between trend detection results and noisy values using the modified Mann-Kendall method is rather complicated and difficult to quantify.

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585 Figure captions

- 586 Fig 1. All rainfall events and events that can be satisfactorily described by rainfall-runoff
- 587 model: (a) Chassieu; (b) La Lechere
- 588 Fig 2. Simple variables in chronological order: (1) Chassieu; (2) La Lechere
- 589 Fig 3. Optimal model parameters of conceptual urban rainfall-runoff models using event-
- 590 based calibration: (1) Chassieu; (2) La Lechere
- 591 Fig 4. Time of concentration from calibrated model parameters: (a) Chassieu; (b) La Lechere
- 592 Fig 5. Scatter plots of lag time and time of concentration: (a) Chassieu; (b) La Lechere
- 593 Fig 6. Sensitivity of P-values to one noisy value: (a) the initial loss in Chassieu; (b) the
- constant loss in Chassieu; (c) the constant loss in La Lechere; (d) the time shift in La Lechere
- 595 Fig 7. Sensitivity of *P*-values to a number of extreme events introduced in the beginning of
- data series using the original Mann-Kendall method: (a) the initial loss in Chassieu; (b) the
- 597 constant loss in Chassieu.
- 598

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600 **Table captions**

- 601 Table 1. Statistical characteristics of event-based variables in Chassieu and La Lechere
- Table 2. Trend detection results on simple variables without modeling
- Table 3. Parameters of conceptual rainfall-runoff model in the two catchments
- Table 4. Trend detection results for the parameters in the conceptual urban rainfall-runoff
- 605 model

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Catchment	Variables	Characteristics	Median [min- max]	Mean (standard deviation)
		Depth (mm)	4.8 [1.2-134.6]	9.0 (12.5)
	D : C 11	Duration (h)	4.7 [0.6-42.6]	6.8 (6.3)
	Rainfall	Mean intensity (mm/h)	1.00 [0.21-23.1]	1.70 (2.12)
Chassieu	Runoff	Total volume $(\times 10^3 \text{ m}^3)$	2.00[0.16-60.7]	4.29 (6.68)
(477 events)	deducted with baseflow	Average rate (L/s)	58.2 [0.90-805.7]	98.7 (112.0)
	Baseflow	(L/s)	2.0 [0.0-46.8]	4.5 (5.4)
	Rainfall-runoff process	Runoff coefficient (-)	0.30 [0.03-0.81]	0.31 (0.10)
		Lag time (min)	77.1[-23.0-207.4]	81.1 (33.5)
	Rainfall	Depth (mm)	7.4 [2.0-91.4]	10.6 (10.1)
		Duration (h)	5.9 [0.5-33.8]	7.7 (6.4)
		Mean intensity (mm/h)	1.4 [0.3-16.3]	2.0 (2.0)
La Lechere	runoff	Total volume $(\times 10^3 \text{ m}^3)$	0.93 [0.10-28.0]	2.07 (3.34)
(398 events)	deducted with baseflow	Average rate (L/s)	24.3 [5.0- 410.6]	44.7 (55.3)
	Baseflow	(L/s)	7.2 [0-317.6]	15.8 (25.6)
	Rainfall-runoff	Runoff coefficient (-)	0.04 [0.01- 0.30]	0.05 (0.04)
	process	Lag time (min)	110 [-50-405]	121 (57)

607 Ta	le 1. Statistical	characteristics of	event-based	variables in	Chassieu ar	nd La Lechere
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	Variable	Parameter	Mann-	<i>P</i> -value	Slope
			Kendall	(%)	(-
			test		/year)
	Rainfall	Depth (mm)	Original	10.2	-
		Duration (h)	Original	58.2	-
		Mean intensity (mm/h)	Modified	12.5	-
	Runoff without	Total runoff volume	Original	0.1**	98.7
Chassieu	baseflow	(m^3)	-		
(477 events)		Mean runoff flow rate	Original	$1.2 \times 10^{-2^{**}}$	3.1
		(L/s)	-		
	Rainfall-runoff	Runoff coefficient (-)	Original	1.4×10 ^{-9**}	0.012
	processes	Lag time (min)	Modified	43.6	-
	Other	Baseflow (L/s)	Modified	2.3*	0.32
	Rainfall	Depth (mm)	Original	34.6	-
		Duration (h)	Original	3.2*	0.18
		Mean intensity (mm/h)	Original	6.0	-
	Runoff	Total runoff volume	Original	0.9^{**}	35.8
La Lechere		(m^3)	C		
(398 events)		Mean runoff flow rate	Original	3.7*	0.7
		(L/s)	-		
	Rainfall-runoff	Runoff coefficient (-)	Modified	0.2^{**}	0.002
	processes	Lag time (min)	Modified	68.5	-
	Other	Baseflow (L/s)	Modified	55.6	-

Table 2. Trend detection results on simple variables without modeling

610 The *P*-value is remarked with * when it is significant at 5% level, with ** when it is significant 611 at 1% level.

- 612
- 613

614

Table 3. Parameters of conceptual rainfall-runoff model in the two catchments

		Cł	La Lechere			
parameter	Search	Median	Mean	Search	Median	Mean
	range	[min-	(standard	range	[min-max]	(standard
		max]	deviation)			deviation)
Initial loss	[0, 2]	0.67 [0.0-	0.79 (0.59)	[0, 3]	1.6 [0.0-3.0]	1.6 (1.1)
(mm)		2.0]				
Constant	[0, 1]	0.65 [0.0-	0.64 (0.14)	[0, 1]	0.88	0.86 (0.09)
loss (-)		0.98]			[0.25,0.97]	
Reservoir	[1, 120]	15.5 [1.0-	19.0 (12.0)	[1,240]	26.6 [2.4-	36.3 (34.9)
constant		90.2]			240]	
(min)						
Time shift	[1,60]	14.0 [1.2-	16.1 (11.9)	[1,120]	2 [0-66]	4.5 (7.0)
(min)		60.0]				

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Table 4 Trend detection results for the parameters in the conceptual urban rainfall-runoff

618 model

		Parameter	Mann- Kendall test	P-value (%)	Slope (- /year)
	Model	Initial loss	Original	43.2	- / year)
~	parameters	Constant loss	Original	2.7×10 ⁻ 9**	-0.014
Chassieu (477		Reservoir constant	Modified	39.9	-
events)		Time shift	Modified	88.5	-
	Other	Time of	Modified	99.1	-
	parameters	concentration (min)			
	Model	Initial loss	Modified	19.2	-
	parameters	Constant loss	Modified	0.01**	-0.005
La Lechere		Reservoir constant	Modified	0.3**	-0.8
(398 events)		Time shift	Original	97.2	-
	Other	Time of	Modified	2.0^{*}	-3.5
	parameters	concentration (min)			

619

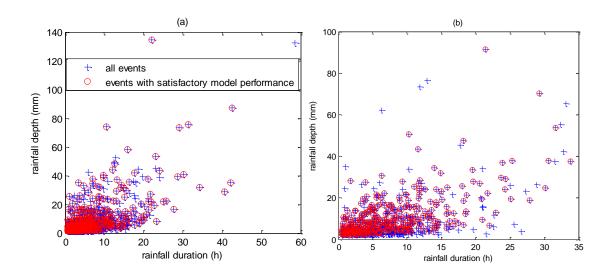




Fig. 1. All rainfall events and events that can be satisfactorily described by rainfall-runoff

623 model: (a) Chassieu; (b) La Lechere

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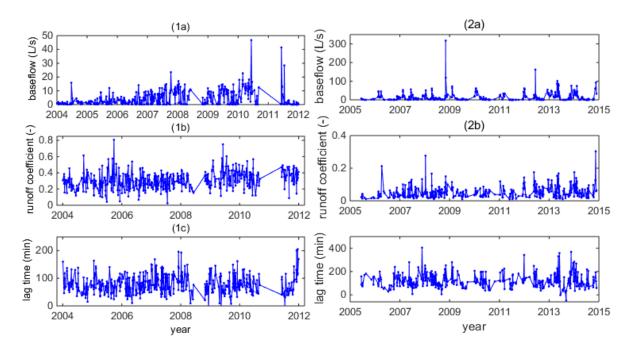


Fig. 2 Simple variables in chronological order: (1) Chassieu; (2) La Lechere

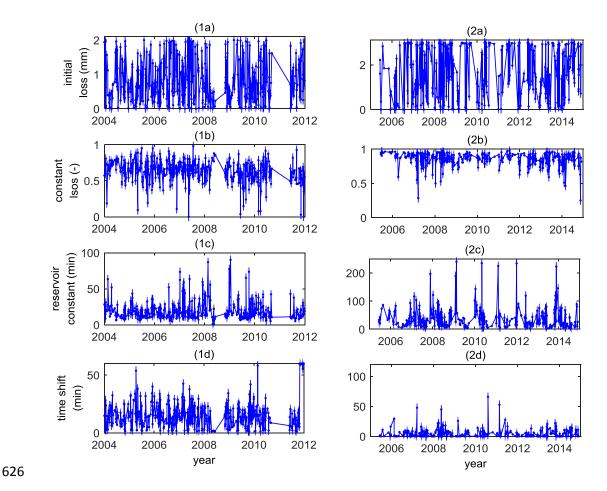
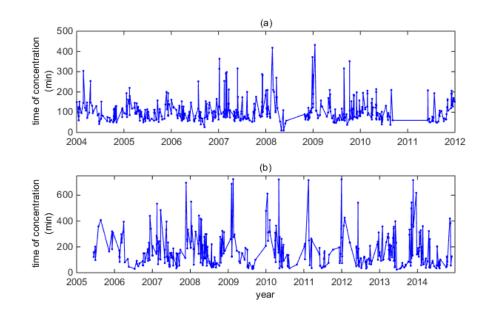


Fig. 3 Optimal model parameters of conceptual urban rainfall-runoff models using eventbased calibration: (1) Chassieu; (2) La Lechere

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630 Fig. 4 Time of concentration from calibrated model parameters: (a) Chassieu; (b) La Lechere

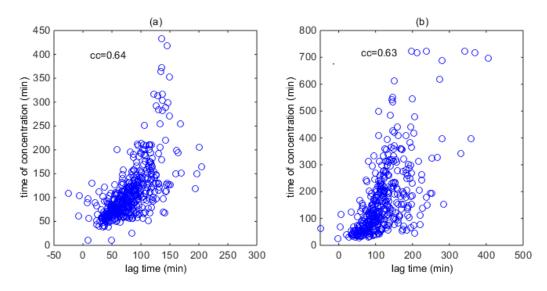






Fig. 5 Scatter plots of lag time and time of concentration: (a) Chassieu; (b) La Lechere

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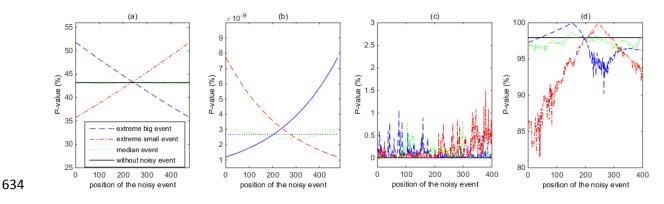
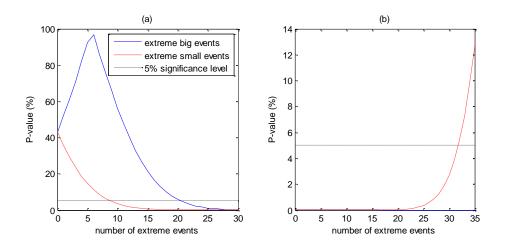


Fig. 6 Sensitivity of P-values to one noisy value: (a) the initial loss in Chassieu; (b) the

636 constant loss in Chassieu; (c) the constant loss in La Lechere; (d) the time shift in La Lechere

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Fig. 7 Sensitivity of *P*-values to a number of extreme events introduced in the beginning of
data series using the original Mann-Kendall method: (a) the initial loss in Chassieu; (b) the
constant loss in Chassieu.