Convergence rates in the central limit theorem for weighted sums of Bernoulli random fields

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CONVERGENCE RATES IN THE CENTRAL LIMIT THEOREM FOR WEIGHTED SUMS OF BERNOULLI RANDOM FIELDS

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Abstract. We prove moment inequalities for a class of functionals of i.i.d. random fields. We then derive rates in the central limit theorem for weighted sums of such randoms fields via an approximation by m-dependent random fields.

1. Introduction and main results

1.1. **Goal of the paper.** In its simplest form, the central limit theorem states that if $(X_i)_{i\geqslant 1}$ is an independent identically distributed (i.i.d.) sequence of centered random variables having variance one, then the sequence $\left(n^{-1/2}\sum_{i=1}^{n} X_i\right)_{n\geqslant 1}$ converges in distribution to a standard normal random variable. If X_1 has a finite moment of order three, Berry [Ber41] and Esseen [Ess42] gave the following convergence rate:

$$\sup_{t\in\mathbb{R}}\left|\mathbb{P}\left\{n^{-1/2}\sum_{i=1}^{n} X_i \leqslant t\right\} - \mathbb{P}\left\{N \leqslant t\right\}\right| \leqslant C\mathbb{E}\left[\left|X_1\right|^3\right] n^{-1/2}, \tag{1.1.1}$$

where C is a numerical constant and N has a standard normal distribution. The question of extending the previous result to a larger class of sequence have received a lot of attention. When X_i can be represented as a function of an i.i.d. sequence, optimal convergence rates are given in [Jir16].

In this paper, we will focus on random fields, that is, collection of random variables indexed by $\mathbb{Z}^d$ and more precisely in Bernoulli random fields, which are defined as follows.

Definition 1.1. *Let $d \geqslant 1$ be an integer. The random field $(X_{\boldsymbol{n}})_{\boldsymbol{n}\in\mathbb{Z}^d}$ is said to be Bernoulli if there exist an i.i.d. random field $(\varepsilon_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ and a measurable function $f\colon \mathbb{R}^{\mathbb{Z}^d} \to \mathbb{R}$ such that $X_{\boldsymbol{n}} = f\left((\varepsilon_{\boldsymbol{n}-\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}\right)$ for each $\boldsymbol{n}\in\mathbb{Z}^d$.*

We are interested in the asymptotic behavior of the sequence $(S_n)_{n\geqslant 1}$ defined by

$$S_n := \sum_{\boldsymbol{i}\in\mathbb{Z}^d} b_{n,\boldsymbol{i}} X_{\boldsymbol{i}}, \tag{1.1.2}$$

where $b_n := (b_{n,\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ is an element of $\ell^2\left(\mathbb{Z}^d\right)$. Under appropriated conditions on the dependence of the random field $(X_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ and the sequence of weights $(b_n)_{n\geqslant 1}$ that will be specified later, the sequence $(S_n/\left\|b_n\right\|_2)_{n\geqslant 1}$ converges in law to a normal distribution [KVW16]. The goal of this paper is to provide bounds of the type Berry-Esseen in order to give convergence rates in the central limit theorem.

Date: October 29, 2018.

Key words and phrases. random fields; moment inequalities; central limit theorem.

This type of question has been addressed for the so-called BL (θ)-dependent random fields [BK06], martingale differences random fields [NP04], positively and negatively dependent random fields [Bul96, Pav93] and mixing random fields [BS07, BD90].

In order to establish this kind of results, we need several ingredients. First, we need convergence rates for m-dependent random fields. Second, a Bernoulli random field can be decomposed as the sum of an m-dependent random field and a remainder. The control of the contribution of the remainder is done by a moment inequality in the spirit of Rosenthal's inequality [Ros70]. One of the main applications of such an inequality is the estimate of the convergence rates in the central limit theorem for random fields that can be expressed as a functional of a random field consisting of i.i.d. random variables. The method consists in approximating the considered random field by an m-dependent one, and in controlling the approximation with the help of the established moment inequality. In the one dimensional case, probability and moment inequalities have been established in [LXW13] for maxima of partial sums of Bernoulli sequences. The techniques used therein permit to derive results for weighted sums of such sequences.

The paper is organized as follows. In Subsections 1.2, we give the material which is necessary to understand the moment inequality stated in Theorem 1.4. We then give the results on convergence rates in Subsection 1.3 (for weighted sums, sums on subsets of $\mathbb{Z}^d$ and in a regression model) and compare the obtained results in the case of linear random fields with some existing ones. Section 2 is devoted to the proofs.

1.2. **Backgroud.** The following version of Rosenthal's inequality is due to Johnson, Schechtman and Zinn [JSZ85]: if $(Y_i)_{i=1}^n$ are independent centered random variables with a finite moment of order $p \geqslant 2$, then

$$\left\| \sum_{i=1}^n Y_i \right\|_p \leqslant \frac{14.5p}{\log p} \left(\left(\sum_{i=1}^n \|Y_i\|_2^2 \right)^{1/2} + \left(\sum_{i=1}^n \|Y_i\|_p^p \right)^{1/p} \right), \tag{1.2.1}$$

where $\|Y\|_q := (\mathbb{E}[|Y|^q])^{1/q}$ for $q \geqslant 1$.

It was first estalish without explicit constant in Theorem 3 of [Ros70].

Various extension of Rosenthal-type inequalities have been obtained under mixing conditions [Sha95, Rio00] or projective conditions [PUW07, Rio09, MP13]. We are interested by extensions of (1.2.1) to the setting of dependent random fields.

Throughout the paper, we shall use the following notations.

(N.1) For a positive integer d, the set $\{1, \ldots, d\}$ is denoted by $[d]$.
(N.2) The coordinatewise order is denoted by $\preccurlyeq$, that is, for $\boldsymbol{i} = (i_q)_{q=1}^d \in \mathbb{Z}^d$ and $\boldsymbol{j} = (j_q)_{q=1}^d \in \mathbb{Z}^d$, $\boldsymbol{i} \preccurlyeq \boldsymbol{j}$ means that $i_k \leqslant j_k$ for any $k \in [d]$.
(N.3) For $k \in [d]$, $\boldsymbol{e_k}$ denotes the element of $\mathbb{Z}^d$ whose qth coordinate is 1 and all the others are zero. Moreover, we write $\mathbf{0} = (0, \ldots, 0)$ and $\mathbf{1} = (1, \ldots, 1)$.
(N.4) For $\boldsymbol{n} = (n_k)_{k=1}^d \in \mathbb{N}^d$, we write the product $\prod_{k=1}^d n_q$ as $|\boldsymbol{n}|$.
(N.5) The cardinality of a set I is denoted by $|I|$.
(N.6) For a real number x, we denote by $[x]$ the unique integer such that $[x] \leqslant x < [x] + 1$.
(N.7) We write Φ for the cumulative distribution function of a standard normal law.

(N.8) If Λ is a subset of $\mathbb{Z}^d$ and $\boldsymbol{k} \in \mathbb{Z}^d$, then $\Lambda - \boldsymbol{k}$ is defined as $\{\boldsymbol{l} - \boldsymbol{k}, \boldsymbol{l} \in \Lambda\}$.

(N.9) For $q \geqslant 1$, we denote by $\ell^q\left(\mathbb{Z}^d\right)$ the space of sequences $\boldsymbol{a} := (a_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ such that $\|\boldsymbol{a}\|_{\ell^q} := \left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d} |a_{\boldsymbol{i}}|^q\right)^{1/q} < +\infty$.

(N.10) For $\boldsymbol{i} = (i_q)_{q=1}^d$, the quantity $\|\boldsymbol{i}\|_\infty$ is defined as $\max_{1\leqslant q\leqslant d} |i_q|$.

Let $(Y_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ be a random field. The sum $\sum_{\boldsymbol{i}\in\mathbb{Z}^d} Y_{\boldsymbol{i}}$ is understood as the $\mathbb{L}^1$-limit of the sequence $(S_k)_{k\geqslant 1}$ where $S_k = \sum_{\boldsymbol{i}\in\mathbb{Z}^d, \|\boldsymbol{i}\|_\infty \leqslant k} Y_{\boldsymbol{i}}$.

Following [Wu05] we define the physical dependence measure.

Definition 1.2. *Let $(X_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d} := \left(f\left((\varepsilon_{\boldsymbol{i}-\boldsymbol{j}})_{\boldsymbol{j}\in\mathbb{Z}^d}\right)\right)_{\boldsymbol{i}\in\mathbb{Z}^d}$ be a Bernoulli random field, $p \geqslant 1$ and $(\varepsilon'_{\boldsymbol{u}})_{\boldsymbol{u}\in\mathbb{Z}^d}$ be an i.i.d. random field which is independent of the i.i.d. random field $(\varepsilon_{\boldsymbol{u}})_{\boldsymbol{u}\in\mathbb{Z}^d}$ and has the same distribution as $(\varepsilon_{\boldsymbol{u}})_{\boldsymbol{u}\in\mathbb{Z}^d}$. For $\boldsymbol{i} \in \mathbb{Z}^d$, we introduce the physical dependence measure*

$$\delta_{\boldsymbol{i},p} := \|X_{\boldsymbol{i}} - X^*_{\boldsymbol{i}}\|_p \tag{1.2.2}$$

where $X^*_{\boldsymbol{i}} = f\left(\left(\varepsilon^*_{\boldsymbol{i}-\boldsymbol{j}}\right)_{\boldsymbol{j}\in\mathbb{Z}^d}\right)$ and $\varepsilon^*_{\boldsymbol{u}} = \varepsilon_{\boldsymbol{u}}$ if $\boldsymbol{u} \neq \boldsymbol{0}$, $\varepsilon^*_{\boldsymbol{0}} = \varepsilon'_{\boldsymbol{0}}$.

In [EVW13, BD14], various examples of Bernoulli random fields are given, for which the physical dependence measure is either computed or estimated. Proposition 1 of [EVW13] also gives the following moment inequality: if Γ is a finite subset of $\mathbb{Z}^d$, $(a_{\boldsymbol{i}})_{\boldsymbol{i}\in\Gamma}$ is a family of real numbers and $p \geqslant 2$, then for any Bernoulli random field $(X_{\boldsymbol{n}})_{\boldsymbol{n}\in\mathbb{Z}^d}$,

$$\left\|\sum_{\boldsymbol{i}\in\Gamma} a_{\boldsymbol{i}} X_{\boldsymbol{i}}\right\|_p \leqslant \left(2p \sum_{\boldsymbol{i}\in\Gamma} a_{\boldsymbol{i}}^2\right)^{1/2} \cdot \sum_{\boldsymbol{j}\in\mathbb{Z}^d} \delta_{\boldsymbol{j},p}. \tag{1.2.3}$$

This was used in [EVW13, BD14] in order to establish functional central limit theorems. Truquet [Tru10] also obtained an inequality in this spirit. If $(X_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ is i.i.d. and centered, (1.2.1) would give

$$\left\|\sum_{\boldsymbol{i}\in\Gamma} a_{\boldsymbol{i}} X_{\boldsymbol{i}}\right\|_p \leqslant C \left(\sum_{\boldsymbol{i}\in\Gamma} a_{\boldsymbol{i}}^2\right)^{1/2} \|X_{\boldsymbol{1}}\|_p, \tag{1.2.4}$$

while Rosenthal's inequality (1.2.1) would give

$$\left\|\sum_{\boldsymbol{i}\in\Gamma} a_{\boldsymbol{i}} X_{\boldsymbol{i}}\right\|_p \leqslant C \left(\sum_{\boldsymbol{i}\in\Gamma} a_{\boldsymbol{i}}^2\right)^{1/2} \|X_{\boldsymbol{1}}\|_2 + C \left(\sum_{\boldsymbol{i}\in\Gamma} |a_{\boldsymbol{i}}|^p\right)^{1/p} \|X_{\boldsymbol{1}}\|_p, \tag{1.2.5}$$

a better result in this context.

In the case of linear processes, equality $\delta_{\boldsymbol{j},p} \leqslant K\delta_{\boldsymbol{j},2}$ holds for a constant K which does not depend on $\boldsymbol{j}$. However, there are processes for which such an inequality does not hold.

Example 1.3. We give an example of a random field such that there is no constant K such that $\delta_{\boldsymbol{j},p} \leqslant K\delta_{\boldsymbol{j},2}$ holds for all $\boldsymbol{j} \in \mathbb{Z}^d$. Let $p \geqslant 2$ and let $(\varepsilon_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ be an i.i.d. random field and for each $\boldsymbol{k} \in \mathbb{Z}^d$, let $f_{\boldsymbol{k}}\colon \mathbb{R} \to \mathbb{R}$ be a function such that the random variable $Z_{\boldsymbol{k}} := f_{\boldsymbol{k}}(\varepsilon_{\boldsymbol{0}})$ is centered and has a finite moment of order p, and $\sum_{\boldsymbol{k}\in\mathbb{Z}^d} \|Z_{\boldsymbol{k}}\|_2^2 < +\infty$. Define $X_{\boldsymbol{n}} := \lim_{N\to+\infty} \sum_{-N\boldsymbol{1}\preccurlyeq\boldsymbol{j}\preccurlyeq N\boldsymbol{1}} f_{\boldsymbol{k}}(\varepsilon_{\boldsymbol{n}-\boldsymbol{k}})$, where the limit is taken in $\mathbb{L}^2$. Then $X_{\boldsymbol{i}} - X^*_{\boldsymbol{i}} = f_{\boldsymbol{i}}(\varepsilon_{\boldsymbol{0}}) - f_{\boldsymbol{i}}(\varepsilon'_{\boldsymbol{0}})$ hence $\delta_{\boldsymbol{i},2}$ is of order $\|Z_{\boldsymbol{i}}\|_2$ while $\delta_{\boldsymbol{i},p}$ is of order $\|Z_{\boldsymbol{i}}\|_p$.

Consequently, having the ℓ^p-norm instead of the ℓ^2-norm of the $(a_{\boldsymbol{i}})_{\boldsymbol{i}\in\Gamma}$ is more suitable.

1.3. **Mains results.** We now give a Rosenthal-like inequality for weighted sums of Bernoulli random fields in terms of the physical dependence measure.

Theorem 1.4. *Let $\left\{\varepsilon_{\boldsymbol{i}}, \boldsymbol{i} \in \mathbb{Z}^d\right\}$ be an i.i.d. set of random variables. Then for any measurable function $f\colon \mathbb{R}^{\mathbb{Z}^d} \to \mathbb{R}$ such that $X_{\boldsymbol{j}} := f\left((X_{\boldsymbol{j}-\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}\right)$ has a finite moment of order $p \geqslant 2$ and is centered, and any $(a_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d} \in \ell^2\left(\mathbb{Z}^d\right)$,*

$$\left\|\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}} X_{\boldsymbol{i}}\right\|_p \leqslant \frac{14.5p}{\log p}\left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}}^2\right)^{1/2} \sum_{j=0}^{+\infty} (4j+4)^{d/2} \left\|X_{\boldsymbol{0},j}\right\|_2 + \frac{14.5p}{\log p}\left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d} |a_{\boldsymbol{i}}|^p\right)^{1/p} \sum_{j=0}^{+\infty} (4j+4)^{d(1-1/p)} \left\|X_{\boldsymbol{0},j}\right\|_p, \quad (1.3.1)$$

where for $j \geqslant 1$,

$$X_{\boldsymbol{0},j} = \mathbb{E}\left[X_{\boldsymbol{0}} \mid \sigma\left\{\varepsilon_{\boldsymbol{u}}, \|\boldsymbol{u}\|_\infty \leqslant j\right\}\right] - \mathbb{E}\left[X_{\boldsymbol{0}} \mid \sigma\left\{\varepsilon_{\boldsymbol{u}}, \|\boldsymbol{u}\|_\infty \leqslant j-1\right\}\right] \quad (1.3.2)$$

and $X_{\boldsymbol{0},0} = \mathbb{E}\left[X_{\boldsymbol{0}} \mid \sigma\left\{\varepsilon_{\boldsymbol{0}}\right\}\right]$.

We can formulate a version of inequality (1.3.1) where the right hand side is expressed in terms of the coefficients of physical dependence measure. The obtained result is not directly comparable to (1.2.3) because of the presence of the ℓ^p-norm of the coefficients.

Corollary 1.5. *Let $\left\{\varepsilon_{\boldsymbol{i}}, \boldsymbol{i} \in \mathbb{Z}^d\right\}$ be an i.i.d. set of random variables. Then for any measurable function $f\colon \mathbb{R}^{\mathbb{Z}^d} \to \mathbb{R}$ such that $X_{\boldsymbol{j}} := f\left((X_{\boldsymbol{j}-\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}\right)$ has a finite moment of order $p \geqslant 2$ and is centered, and any $(a_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d} \in \ell^2\left(\mathbb{Z}^d\right)$,*

$$\left\|\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}} X_{\boldsymbol{i}}\right\|_p \leqslant \sqrt{2}\frac{14.5p}{\log p}\left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}}^2\right)^{1/2} \sum_{j=0}^{+\infty} (4j+4)^{d/2} \left(\sum_{\|\boldsymbol{i}\|_\infty = j} \delta_{\boldsymbol{i},2}^2\right)^{1/2} + \sqrt{2}\frac{14.5p}{\log p}\sqrt{p-1}\left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d} |a_{\boldsymbol{i}}|^p\right)^{1/p} \sum_{j=0}^{+\infty} (4j+4)^{d(1-1/p)} \left(\sum_{\|\boldsymbol{i}\|_\infty = j} \delta_{\boldsymbol{i},p}^2\right)^{1/2}. \quad (1.3.3)$$

Let $(X_{\boldsymbol{j}})_{\boldsymbol{j}\in\mathbb{Z}^d} = f\left((\varepsilon_{\boldsymbol{j}-\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}\right)$ be a centered square integrable Bernoulli random field and for any positive integer n, let $b_n := (b_{n,\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ be an element of $\ell^2\left(\mathbb{Z}^d\right)$. We are interested in the asymptotic behavior of the sequence $(S_n)_{n\geqslant 1}$ defined by

$$S_n := \sum_{\boldsymbol{i}\in\mathbb{Z}^d} b_{n,\boldsymbol{i}} X_{\boldsymbol{i}}. \quad (1.3.4)$$

Let us denote for $\boldsymbol{k} \in \mathbb{Z}^d$ the map $\tau_{\boldsymbol{k}}\colon \ell^2\left(\mathbb{Z}^d\right) \to \ell^2\left(\mathbb{Z}^d\right)$ defined by $\tau_{\boldsymbol{k}}\left((x_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}\right) := (x_{\boldsymbol{i}+\boldsymbol{k}})_{\boldsymbol{i}\in\mathbb{Z}^d}$.

In [KVW16], Corollary 2.6 gives the following result: under a Hannan type condition on the random field $(X_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ and under the following condition on the weights: for any $q \in [d]$,

$$\frac{1}{\|b_n\|_{\ell^2}} \left\|\tau_{\boldsymbol{e}_q}(b_n) - b_n\right\|_{\ell^2} = 0, \quad (1.3.5)$$

the series $\sum_{\boldsymbol{i}\in\mathbb{Z}^d}|\mathrm{Cov}\left(X_{\boldsymbol{0}},X_{\boldsymbol{i}}\right)|$ converges and with

$$\sigma:=\left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d}\mathrm{Cov}\left(X_{\boldsymbol{0}},X_{\boldsymbol{i}}\right)\right)^{1/2}, \tag{1.3.6}$$

the sequence $\left(S_n/\left\|b_n\right\|_{\ell^2}\right)_{n\geqslant 1}$ converges in distribution to a centered normal distribution with variance σ^2. The argument relies on an approximation by an m-dependent random field.

The purpose of the next theorem is to give a general speed of convergence. In order to measure it, we define

$$\Delta_n:=\sup_{t\in\mathbb{R}}\left|\mathbb{P}\left\{\frac{S_n}{\left\|b_n\right\|_{\ell^2}}\leqslant t\right\}-\Phi\left(t/\sigma\right)\right|. \tag{1.3.7}$$

The following quantity will also play an important role in the estimation of convergence rates.

$$\varepsilon_n:=\sum_{\boldsymbol{j}\in\mathbb{Z}^d}\left|\mathbb{E}\left[X_{\boldsymbol{0}}X_{\boldsymbol{j}}\right]\right|\sum_{\boldsymbol{i}\in\mathbb{Z}^d}\left|\frac{b_{n,\boldsymbol{i}}b_{n,\boldsymbol{i}+\boldsymbol{j}}}{\left\|b_n\right\|_{\ell^2}}-1\right|. \tag{1.3.8}$$

Theorem 1.6. *Let $p>2$, $p':=\min\{p,3\}$ and let $\left(X_{\boldsymbol{j}}\right)_{\boldsymbol{j}\in\mathbb{Z}^d}=\left(f\left(\left(\varepsilon_{\boldsymbol{j}-\boldsymbol{i}}\right)_{\boldsymbol{i}\in\mathbb{Z}^d}\right)\right)_{\boldsymbol{j}\in\mathbb{Z}^d}$ be a centered Bernoulli random field with a finite moment of order p and for any positive integer n, let $b_n:=\left(b_{n,\boldsymbol{i}}\right)_{\boldsymbol{i}\in\mathbb{Z}^d}$ be an element of $\ell^2\left(\mathbb{Z}^d\right)$ such that for any $n\geqslant 1$, the set $\left\{\boldsymbol{k}\in\mathbb{Z}^d,b_{n,\boldsymbol{k}}\neq 0\right\}$ is finite and nonempty, $\lim_{n\to+\infty}\left\|b_n\right\|_{\ell^2}=+\infty$ and (1.3.5) holds for any $q\in[d]$. Assume that for some positive α and β, the following series are convergent:*

$$C_2\left(\alpha\right):=\sum_{i=0}^{+\infty}\left(i+1\right)^{d/2+\alpha}\left\|X_{\boldsymbol{0},i}\right\|_2 \text{ and } C_p\left(\beta\right):=\sum_{i=0}^{+\infty}\left(i+1\right)^{d(1-1/p)+\beta}\left\|X_{\boldsymbol{0},i}\right\|_p. \tag{1.3.9}$$

Let S_n be defined by (1.3.4),

Assume that $\sum_{\boldsymbol{i}\in\mathbb{Z}^d}|\mathrm{Cov}\left(X_{\boldsymbol{0}},X_{\boldsymbol{i}}\right)|$ is finite and that σ be given by (1.3.6) is positive. Let $\gamma>0$ and let

$$n_0:=\inf\left\{N\geqslant 1\mid\forall n\geqslant N,\sqrt{\sigma^2+\varepsilon_n}-29\left(\log 2\right)^{-1}C_2\left(\alpha\right)\left(\left[\left\|b_n\right\|_{\ell^2}\right]^{\gamma}\right)^{-\alpha}\geqslant\sigma/2\right\}. \tag{1.3.10}$$

Then for each $n\geqslant n_0$,

$$\begin{aligned}\widetilde{\Delta_n}\leqslant{}& 150\left(29\left(\left[\left\|b_n\right\|_{\ell^2}\right]+21\right)^{\gamma}+21\right)^{\left(p'-1\right)d}\left\|X_{\boldsymbol{0}}\right\|_{p'}^{p'}\left(\frac{\left\|b_n\right\|_{\ell^{p'}}}{\left\|b_n\right\|_{\ell^2}}\right)^{p'}\left(\sigma/2\right)^{-p'}\\&+\left(2\frac{\left|\varepsilon_n\right|}{\sigma^2}+80\left(\log 2\right)^{-1}\frac{\left\|b_n\right\|_{\ell^2}^{-\gamma\alpha}}{\sigma^2}C_2\left(\alpha\right)^2\right)\left(2\pi e\right)^{-1/2}\\&+\left(\frac{14.5p}{\sigma\log p}4^{d/2}\left\|b_n\right\|_{\ell^2}^{-\gamma\alpha}C_2\left(\alpha\right)\right)^{\frac{p}{p+1}}+\left(\frac{\left\|b_n\right\|_{\ell^p}}{\sigma\left\|b_n\right\|_{\ell^2}}\frac{14.5p}{\log p}4^{d(1-1/p)}\left\|b_n\right\|_{\ell^2}^{-\gamma\beta}C_p\left(\beta\right)\right)^{\frac{p}{p+1}}.\end{aligned} \tag{1.3.11}$$

In particular, there exists a constant κ such that for all $n\geqslant n_0$,

$$\Delta_n\leqslant\kappa\left(\left\|b_n\right\|_{\ell^2}^{\gamma\left(p'-1\right)d-p'}\left\|b_n\right\|_{\ell^{p'}}^{p'}+\left|\varepsilon_n\right|+\left\|b_n\right\|_{\ell^2}^{-\gamma\alpha\frac{p}{p+1}}+\left\|b_n\right\|_{\ell^p}^{\frac{p}{p+1}}\left\|b_n\right\|_{\ell^2}^{-\frac{p}{p+1}(\gamma\beta+1)}\right). \tag{1.3.12}$$

Remark 1.7. If (1.3.5), $\lim_{n\to+\infty}\left\|b_n\right\|_{\ell^2}=+\infty$ and the family $\left(\delta_{\boldsymbol{i},2}\right)_{\boldsymbol{i}\in\mathbb{Z}^d}$ is summable, then the sequence $\left(\varepsilon_n\right)_{n\geqslant 1}$ converges to 0 hence n_0 is well-defined. However, it is not clear to us whether the finiteness of $C_2\left(\alpha\right)$ combined with (1.3.5) and $\lim_{n\to+\infty}\left\|b_n\right\|_{\ell^2}=+\infty$ imply that

$\sum_{\boldsymbol{j}\in\mathbb{Z}^d}\left|\mathbb{E}\left[X_{\boldsymbol{0}}X_{\boldsymbol{j}}\right]\right|$ is finite. Nevertheless, we can show an analogous result in terms of $\delta_{\boldsymbol{i},p}$ coefficients by changing the following in the statement of Theorem 1.6:

(1) the definition of $C_2(\alpha)$ should be replaced by

$$C_2(\alpha) := \sqrt{2}\sum_{j=0}^{+\infty}(j+1)^{d/2+\alpha}\left(\sum_{\|\boldsymbol{i}\|_\infty=j}\delta_{\boldsymbol{i},2}^2\right)^{1/2}; \tag{1.3.13}$$

(2) the definition of $C_p(\beta)$ should be replaced by

$$C_p(\beta) := \sqrt{2(p-1)}\sum_{j=0}^{+\infty}(j+1)^{d(1-1/p)+\beta}\left(\sum_{\|\boldsymbol{i}\|_\infty=j}\delta_{\boldsymbol{i},2}^2\right)^{1/2}. \tag{1.3.14}$$

In this case, the convergence of $\sum_{\boldsymbol{i}\in\mathbb{Z}^d}\left|\mathrm{Cov}\left(X_{\boldsymbol{0}},X_{\boldsymbol{i}}\right)\right|$ holds (cf. Proposition 2 in [EVW13]).

Recall notation (N.8). Let $(\Lambda_n)_{n\geqslant 1}$ be a sequence of subsets of $\mathbb{Z}^d$. The choice $b_{n,\boldsymbol{j}}=1$ if $\boldsymbol{j}\in\Lambda_n$ and 0 otherwise yields the following corollary for set-indexed partial sums.

Corollary 1.8. *Let $(X_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ be a centered Bernoulli random field with a finite moment of order $p\geqslant 2$, $p':=\min\{p,3\}$ and let $(\Lambda_n)_{n\geqslant 1}$ be a sequence of subset of $\mathbb{Z}^d$ such that $|\Lambda_n|\to+\infty$ and for any $\boldsymbol{k}\in\mathbb{Z}^d$, $\lim_{n\to+\infty}\left|\Lambda_n\cap(\Lambda_n-\boldsymbol{k})\right|/|\Lambda_n|=1$. Assume that the series defined in (1.3.9) are convergent for some positive α and β, that $\sum_{\boldsymbol{i}\in\mathbb{Z}^d}\left|\mathrm{Cov}\left(X_{\boldsymbol{0}},X_{\boldsymbol{i}}\right)\right|$ is finite and that σ defined by (1.3.6) is positive. Let $\gamma>0$ and n_0 be defined by (1.3.10). There exists a constant κ such that for any $n\geqslant n_0$,*

$$\sup_{t\in\mathbb{R}}\left|\mathbb{P}\left\{\frac{\sum_{\boldsymbol{i}\in\Lambda_n}X_{\boldsymbol{i}}}{|\Lambda_n|^{1/2}}\leqslant t\right\}-\Phi(t/\sigma)\right| \leqslant \kappa\left(|\Lambda_n|^q+\sum_{\boldsymbol{j}\in\mathbb{Z}^d}\left|\mathbb{E}\left[X_{\boldsymbol{0}}X_{\boldsymbol{j}}\right]\right|\left|\frac{|\Lambda_n\cap(\Lambda_n-\boldsymbol{j})|}{|\Lambda_n|}-1\right|\right), \tag{1.3.15}$$

where

$$q:=\max\left\{\frac{\gamma(p'-1)d-p'}{2}+1;-\gamma\alpha\frac{p}{2(p+1)};\frac{2-p-p\gamma\beta}{2(p+1)}\right\}. \tag{1.3.16}$$

We consider now the following regression model:

$$Y_{\boldsymbol{i}}=g\left(\frac{\boldsymbol{i}}{n}\right)+X_{\boldsymbol{i}},\quad \boldsymbol{i}\in\Lambda_n:=\{1,\ldots,n\}^d, \tag{1.3.17}$$

where $g\colon[0,1]^d\to\mathbb{R}$ is an unknown smooth function and $(X_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ is a zero mean stationary Bernoulli random field. Let K be a probability kernel defined on $\mathbb{R}^d$ and let $(h_n)_{n\geqslant 1}$ be a sequence of positive numbers which converges to zero and which satisfies

$$\lim_{n\to+\infty}nh_n=+\infty \text{ and } \lim_{n\to+\infty}nh_n^{d+1}=0. \tag{1.3.18}$$

We estimate the function g by the kernel estimator g_n defined by

$$g_n(\boldsymbol{x})=\frac{\sum_{\boldsymbol{i}\in\Lambda_n}Y_{\boldsymbol{i}}K\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right)}{\sum_{\boldsymbol{i}\in\Lambda_n}K\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right)},\quad x\in[0,1]^d. \tag{1.3.19}$$

We make the following assumptions on the regression function g and the probability kernel K:

(A) The probability kernel K fulfills $\int_{\mathbb{R}^d} K(\boldsymbol{u})\,\mathrm{d}\boldsymbol{u} = 1$, is symmetric, non-negative, supported by $[-1,1]^d$. Moreover, there exist positive constants r, c and C such that for any $\boldsymbol{x},\boldsymbol{y} \in [-1,1]^d$, $|K(\boldsymbol{x}) - K(\boldsymbol{y})| \leqslant r\,\|\boldsymbol{x}-\boldsymbol{y}\|_\infty$ and $c \leqslant K(\boldsymbol{x}) \leqslant C$.

We measure the speed of convergence of $\left((nh_n)^{d/2}\left(g_n(\boldsymbol{x}) - \mathbb{E}\left[g_n(\boldsymbol{x})\right]\right)\right)_{n\geqslant 1}$ to a normal distribution by the use of the quantity

$$\widetilde{\Delta_n} := \sup_{t\in\mathbb{R}} \left| \mathbb{P}\left\{(nh_n)^{d/2}\left(g_n(\boldsymbol{x}) - \mathbb{E}\left[g_n(\boldsymbol{x})\right]\right) \leqslant t\right\} - \Phi\left(\frac{t}{\sigma\,\|K\|_2}\right)\right|. \tag{1.3.20}$$

Two other quantities will be involved, namely,

$$A_n := (nh_n)^{d/2}\left(\sum_{\boldsymbol{i}\in\Lambda_n} K^2\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right)\right)^{1/2} \|K\|^{-1}_{\mathbb{L}^2(\mathbb{R}^d)} \left(\sum_{\boldsymbol{i}\in\Lambda_n} K\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right)\right)^{-1/2} \quad \text{and} \tag{1.3.21}$$

$$\varepsilon_n := \sum_{\boldsymbol{j}\in\mathbb{Z}^d} \left|\mathbb{E}\left[X_{\boldsymbol{0}}X_{\boldsymbol{j}}\right]\right| \left(\sum_{\boldsymbol{i}\in\Lambda_n\cap(\Lambda_n-\boldsymbol{j})} \frac{K\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right) K\left(\frac{\boldsymbol{x}-(\boldsymbol{i}-\boldsymbol{j})/n}{h_n}\right)}{\sum_{\boldsymbol{k}\in\Lambda_n} K^2\left(\frac{\boldsymbol{x}-\boldsymbol{k}/n}{h_n}\right)} - 1\right). \tag{1.3.22}$$

Theorem 1.9. *Let $p > 2$, $p' := \min\{p,3\}$ and let $(X_{\boldsymbol{j}})_{\boldsymbol{j}\in\mathbb{Z}^d} = \left(f\left((\varepsilon_{\boldsymbol{j}-\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}\right)\right)_{\boldsymbol{j}\in\mathbb{Z}^d}$ be a centered Bernoulli random field with a finite moment of order p. Assume that for some positive α and β, the following series are convergent:*

$$C_2(\alpha) := \sum_{i=0}^{+\infty} (i+1)^{d/2+\alpha}\,\|X_{\boldsymbol{0},i}\|_2 \ \text{ and } C_p(\beta) := \sum_{i=0}^{+\infty} (i+1)^{d(1-1/p)+\beta}\,\|X_{\boldsymbol{0},i}\|_p. \tag{1.3.23}$$

Let $g_n(\boldsymbol{x})$ be defined by (1.3.19), $(h_n)_{n\geqslant 1}$ be a sequence which converges to 0 and satisfies (1.3.18),

Assume that $\sum_{\boldsymbol{i}\in\mathbb{Z}^d} |\mathrm{Cov}(X_{\boldsymbol{0}}, X_{\boldsymbol{i}})|$ is finite and that $\sigma := \sum_{\boldsymbol{j}\in\mathbb{Z}^d} \mathrm{Cov}(X_{\boldsymbol{0}}, X_{\boldsymbol{j}}) > 0$. Let $n_1 \in \mathbb{N}$ be such that for each $n \geqslant n_1$,

$$\frac{1}{2} \leqslant (nh_n)^{-d} K\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right) \leqslant \frac{3}{2} \ \text{ and} \tag{1.3.24}$$

$$\frac{1}{2}\|K\|_{\mathbb{L}^2(\mathbb{R}^d)} \leqslant (nh_n)^{-d} K^2\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right) \leqslant \frac{3}{2}\|K\|_{\mathbb{L}^2(\mathbb{R}^d)}. \tag{1.3.25}$$

Let n_0 be the smallest integer for which for all $n \geqslant n_0$,

$$\sqrt{\sigma^2+\varepsilon_n} - 29(\log 2)^{-1} C_2(\alpha)\left(\left[\left(\sum_{\boldsymbol{i}\in\Lambda_n} K\left(\frac{1}{h_n}\left(\boldsymbol{x}-\frac{\boldsymbol{i}}{n}\right)\right)^2\right)^{1/2}\right]^{\gamma}\right)^{-\alpha} \geqslant \sigma/2. \tag{1.3.26}$$

Then there exists a constant κ such that for each $n \geqslant \max\{n_0, n_1\}$,

$$\Delta_n \leqslant \kappa\,|A_n - 1|^{\frac{p}{p+1}} + |\varepsilon_n| + \kappa\,(nh_n)^{\frac{d}{2}\left(\gamma\left(p'-1\right)d-p'+2\right)} + (nh_n)^{-\frac{d}{2}\gamma\alpha\frac{p}{p+1}} + (nh_n)^{\frac{2d-p(\gamma\beta+1)}{2(p+1)}}. \tag{1.3.27}$$

Lemma 1 in [EMS10] shows that under (1.3.18), the sequence $(A_n)_{n\geqslant 1}$ goes to 1 as n goes to infinity and that the integer n_1 is well-defined.

We now consider the case of linear random fields in dimension 2, that is,

$$X_{j_1,j_2} = \sum_{i_1,i_2\in\mathbb{Z}} a_{i_1,i_2}\varepsilon_{j_1-i_1,j_2-i_2}, \tag{1.3.28}$$

where $(a_{i_1,i_2})_{i_1,i_2\mathbb{Z}} \in \ell^1\left(\mathbb{Z}^2\right)$ and $(\varepsilon_{u_1,u_2})_{u_1,u_2\in\mathbb{Z}^2}$ is i.i.d., centered and $\varepsilon_{0,0}$ has a finite variance. We will focus on the case where the weights are of the form $b_{n,i_1,i_2} = 1$ if $1 \leqslant i_1, i_2 \leqslant n$ and $b_{n,i_1,i_2} = 0$ otherwise.

Mielkaitis and Paulauskas [MP11] established the following convergence rate. Denoting

$$\Delta'_n := \sup_{r\geqslant 0}\left|\mathbb{P}\left\{\left|\frac{1}{n}\sum_{i_1,i_2=1}^{n} X_{i_1,i_2}\right| \leqslant r\right\} - \mathbb{P}\left\{|N| \leqslant r\right\}\right| \tag{1.3.29}$$

and assuming that $\mathbb{E}\left[\left|\varepsilon_{0,0}\right|^{2+\delta}\right]$ is finite and

$$\sum_{k_1,k_2\in\mathbb{Z}} \left(|k_1|+1\right)^2\left(|k_2|+1\right)^2 a_{k_1,k_2}^2 < +\infty, \tag{1.3.30}$$

the following estimate holds for Δ'_n:

$$\Delta'_n = O\left(n^{-r}\right), \quad r := \frac{1}{2}\min\left\{\delta, 1-\frac{1}{3+\delta}\right\}. \tag{1.3.31}$$

In the context of Corollary 1.8, the condition on the coefficients reads as follows:

$$\sum_{i=0}^{+\infty}\left((i+1)^{1+\alpha} + (i+1)^{2-2/p+\beta}\right)\left(\sum_{(j_1,j_2):\|(j_1,j_2)\|_\infty = i} a_{j_1,j_2}^2\right)^{1/2} < +\infty, \tag{1.3.32}$$

where $p = 2+\delta$. Let us compare (1.3.30) with (1.3.32). Let $s := \max\left\{1+\alpha, 2-2/p+\beta\right\}$. When $s \geqslant 2$, (1.3.32) implies (1.3.30). However, this implication does not hold if $s < 3/2$. Indeed, let $r \in (s+1, 5/2)$ and define $a_{k_1,k_2} := k_1^{-r}$ if $k_1 = k_2 \geqslant 1$ and $a_{k_1,k_2} := 0$ otherwise. Then (1.3.32) holds whereas (1.3.30) does not.

Let us discuss the convergence rates in the following example. Let $a_{k_1,k_2} := 2^{-|k_1|-|k_2|}$ and let $p = 2+\delta$, where $\delta \in (0,1]$. In our context,

$$\left|\frac{|\Lambda_n \cap (\Lambda_n - \boldsymbol{j})|}{|\Lambda_n|} - 1\right| \leqslant \frac{n^2 - (n-j_1)(n-j_2)}{n^2} \leqslant \frac{j_1+j_2}{n} \tag{1.3.33}$$

hence the convergence of $\sum_{j_1,j_2\in\mathbb{Z}} \left|\mathrm{Cov}\left(X_{0,0}, X_{j_1,j_2}\right)\right|(j_1+j_2)$ guarantees that ε_n in Corollary 1.8 is of order $1/n$. Moreover, since (1.3.32) holds for all α and β, the choice of γ allows to reach rates of the form $n^{-\delta+r_0}$ for any fixed r_0. In particular, when $\delta = 1$, one can reach for any fixed r_0 rates of the form n^{-1+r_0}. In comparison, with the same assumptions, the result of [MP11] gives $n^{-3/8}$.

2. Proofs

2.1. **Proof of Theorem 1.4.** We define for $j \geqslant 1$ and $\boldsymbol{i} \in \mathbb{Z}^d$,

$$X_{\boldsymbol{i},j} = \mathbb{E}\left[X_i \mid \sigma\left(\varepsilon_{\boldsymbol{u}}, \|\boldsymbol{u}-\boldsymbol{i}\|_\infty \leqslant j\right)\right] - \mathbb{E}\left[X_i \mid \sigma\left(\varepsilon_{\boldsymbol{u}}, \|\boldsymbol{u}-\boldsymbol{i}\|_\infty \leqslant j-1\right)\right]. \tag{2.1.1}$$

In this way, by the martingale convergence theorem,

$$X_{\boldsymbol{i}} - \mathbb{E}\left[X_{\boldsymbol{i}} \mid \varepsilon_{\boldsymbol{i}}\right] = \lim_{N\to+\infty} \sum_{j=1}^{N} X_{\boldsymbol{i},j} \tag{2.1.2}$$

hence

$$\left\|\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}} X_{\boldsymbol{i}}\right\|_p \leqslant \sum_{j=1}^{+\infty} \left\|\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}\right\|_p + \left\|\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}} \mathbb{E}\left[X_{\boldsymbol{i}} \mid \varepsilon_{\boldsymbol{i}}\right]\right\|_p. \tag{2.1.3}$$

Let us fix $j \geqslant 1$. We divide $\mathbb{Z}^d$ into blocks. For $\boldsymbol{v} \in \mathbb{Z}^d$, we define

$$A_{\boldsymbol{v}} := \prod_{q=1}^{d} \left(\left[(2j+2)\, v_q, (2j+2)\,(v_q+1) - 1\right] \cap \mathbb{Z}\right), \tag{2.1.4}$$

and if K is a subset of $[d]$, we define

$$E_K := \left\{\boldsymbol{v} \in \mathbb{Z}^d, v_q \text{ is even if and only if } q \in K\right\}. \tag{2.1.5}$$

Therefore, the following inequality takes place

$$\left\|\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}\right\|_p \leqslant \sum_{K\subset[d]} \left\|\sum_{\boldsymbol{v}\in E_K} \sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}\right\|_p. \tag{2.1.6}$$

Observe that the random variable $\sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}$ is measurable for the σ-algebra generated by $\varepsilon_{\boldsymbol{u}}$, where $\boldsymbol{u}$ satisfies $(2j+2)\, v_q - (j+1) \leqslant u_q \leqslant j+1+(2j+2)\,(v_q+1)-1$ for all $q \in [d]$. Since the family $\left\{\varepsilon_{\boldsymbol{u}}, \boldsymbol{u} \in \mathbb{Z}^d\right\}$ is independent, the family $\left\{\sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}, \boldsymbol{v} \in E_K\right\}$ is independent for each fixed $K \subset [d]$. Using inequality (1.2.1), it thus follows that

$$\begin{aligned}\left\|\sum_{\boldsymbol{v}\in E_K} \sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}\right\|_p \leqslant{}& \frac{14.5p}{\log p} \left(\sum_{\boldsymbol{v}\in E_K} \left\|\sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}\right\|_2^2\right)^{1/2} \\ &+ \frac{14.5p}{\log p} \left(\sum_{\boldsymbol{v}\in E_K} \left\|\sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}\right\|_p^p\right)^{1/p}.\end{aligned} \tag{2.1.7}$$

By stationarity, one can see that $\|X_{\boldsymbol{i},j}\|_q = \|X_{\boldsymbol{0},j}\|_q$ for $q \in \{2, p\}$, hence the triangle inequality yields

$$\begin{aligned}\left\|\sum_{\boldsymbol{v}\in E_K} \sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} a_{\boldsymbol{i}} X_{\boldsymbol{i},j}\right\|_p \leqslant{}& \frac{14.5p}{\log p} \|X_{\boldsymbol{0},j}\|_2 \left(\sum_{\boldsymbol{v}\in E_K} \left(\sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} |a_{\boldsymbol{i}}|\right)^2\right)^{1/2} \\ &+ \frac{14.5p}{\log p} \|X_{\boldsymbol{0},j}\|_p \left(\sum_{\boldsymbol{v}\in E_K} \left(\sum_{\boldsymbol{i}\in A_{\boldsymbol{v}}} |a_{\boldsymbol{i}}|\right)^p\right)^{1/p}.\end{aligned} \tag{2.1.8}$$

By Jensen's inequality, for $q \in \{2, p\}$,

$$\left(\sum_{\boldsymbol{i} \in A_{\boldsymbol{v}}} |a_{\boldsymbol{i}}|\right)^q \leqslant |A_{\boldsymbol{v}}|^{q-1} \sum_{\boldsymbol{i} \in A_{\boldsymbol{v}}} |a_{\boldsymbol{i}}|^q \leqslant (2j+2)^{d(q-1)} \sum_{\boldsymbol{i} \in A_{\boldsymbol{v}}} |a_{\boldsymbol{i}}|^q \tag{2.1.9}$$

and using $\sum_{i=1}^N x_i^{1/q} \leqslant N^{\frac{q-1}{q}} \left(\sum_{i=1}^N x_i\right)^{1/q}$, it follows that

$$\sum_{K \subset [d]} \left\| \sum_{\boldsymbol{v} \in E_K} \sum_{\boldsymbol{i} \in A_{\boldsymbol{v}}} a_{\boldsymbol{i}} X_{\boldsymbol{i},j} \right\|_p \leqslant \frac{14.5p}{\log p} \left\| X_{\boldsymbol{0},j} \right\|_2 \left(\sum_{\boldsymbol{i} \in \mathbb{Z}^d} a_i^2 \right)^{1/2} (4j+4)^{d/2}$$
$$+ \frac{14.5p}{\log p} \left\| X_{\boldsymbol{0},j} \right\|_p \left(\sum_{\boldsymbol{i} \in \mathbb{Z}^d} |a_i|^p \right)^{1/p} (4j+4)^{d(1-1/p)} . \tag{2.1.10}$$

Combining (2.1.3), (2.1.6) and (2.1.10), we derive that

$$\left\| \sum_{\boldsymbol{i} \in \mathbb{Z}^d} a_{\boldsymbol{i}} X_{\boldsymbol{i}} \right\|_p \leqslant \frac{14.5p}{\log p} \sum_{j=1}^{+\infty} \left\| X_{\boldsymbol{0},j} \right\|_2 \left(\sum_{\boldsymbol{i} \in \mathbb{Z}^d} a_i^2 \right)^{1/2} (4j+4)^{d/2}$$
$$+ \frac{14.5p}{\log p} \sum_{j=1}^{+\infty} \left\| X_{\boldsymbol{0},j} \right\|_p \left(\sum_{\boldsymbol{i} \in \mathbb{Z}^d} |a_i|^p \right)^{1/p} (4j+4)^{d(1-1/p)} + \left\| \sum_{\boldsymbol{i} \in \mathbb{Z}^d} a_{\boldsymbol{i}} \mathbb{E}\left[X_{\boldsymbol{i}} \mid \varepsilon_{\boldsymbol{i}} \right] \right\|_p . \tag{2.1.11}$$

In order to control the last term, we use inequality (1.2.1) and bound $\left\| \mathbb{E}\left[X_{\boldsymbol{i}} \mid \varepsilon_{\boldsymbol{i}} \right] \right\|_q$ by $\left\| X_{\boldsymbol{0},0} \right\|_q$ for $q \in \{1, 2\}$. This ends the proof of Theorem 1.4.

Proof of Corollary 1.5. The following lemma gives a control of the $\mathbb{L}^q$-norm of $X_{\boldsymbol{0},j}$ in terms of the physical measure dependence.

Lemma 2.1. *For $q \in \{2, p\}$ and $j \in \mathbb{N}$, the following inequality holds*

$$\left\| X_{\boldsymbol{0},j} \right\|_q \leqslant \left(2\left(q-1\right) \sum_{\boldsymbol{i} \in \mathbb{Z}^d, \left\| \boldsymbol{i} \right\|_\infty = j} \delta_{\boldsymbol{i},q}^2 \right)^{1/2} . \tag{2.1.12}$$

Proof. Let j be fixed. Let us write the set of elements of $\mathbb{Z}^d$ whose infinite norm is equal to j as $\{\boldsymbol{v_s}, 1 \leqslant s \leqslant N_j\}$ where $N_j \in \mathbb{N}$. We also assume that $\boldsymbol{v_s} - \boldsymbol{v_{s-1}} \in \{\boldsymbol{e_k}, 1 \leqslant k \leqslant d\}$ for all $s \in \{2, \ldots, N_j\}$.

Denote

$$\mathcal{F}_s := \sigma\left(\varepsilon_{\boldsymbol{u}}, \left\| \boldsymbol{u} \right\|_\infty \leqslant j, \varepsilon_{\boldsymbol{v_t}}, 1 \leqslant t \leqslant s\right), \tag{2.1.13}$$

and $\mathcal{F}_0 := \sigma\left(\varepsilon_{\boldsymbol{u}}, \left\| \boldsymbol{u} \right\|_\infty \leqslant j\right)$. Then $X_{\boldsymbol{0},j} = \sum_{s=1}^{N_j} \mathbb{E}\left[X_{\boldsymbol{0}} \mid \mathcal{F}_s\right] - \mathbb{E}\left[X_{\boldsymbol{0}} \mid \mathcal{F}_{s-1}\right]$, from which it follows, by Theorem 2.1 in [Rio09], that

$$\left\| X_{\boldsymbol{0},j} \right\|_q^2 \leqslant (q-1) \sum_{s=1}^{N_j} \left\| \mathbb{E}\left[X_{\boldsymbol{0}} \mid \mathcal{F}_s\right] - \mathbb{E}\left[X_{\boldsymbol{0}} \mid \mathcal{F}_{s-1}\right] \right\|_q^2 . \tag{2.1.14}$$

Then using similar arguments as in the proof of Theorem 1 (i) in [Wu05] give the bound $\left\| \mathbb{E}\left[X_{\boldsymbol{0}} \mid \mathcal{F}_s\right] - \mathbb{E}\left[X_{\boldsymbol{0}} \mid \mathcal{F}_{s-1}\right] \right\|_q \leqslant \delta_{\boldsymbol{v_s},q} + \delta_{\boldsymbol{v_{s-1}},q}$. This ends the proof of Lemma 2.1. □

Now, Corollary 1.5 follows from an application of Lemma 2.1 with $q = 2$ and $q = p$ respectively. □

2.2. **Proof of Theorem 1.6.** Denote for a random variable Z the quantity

$$\delta(Z) := \sup_{t\in\mathbb{R}} \left|\mathbb{P}\{Z \leqslant t\} - \Phi(t)\right|. \tag{2.2.1}$$

We say that a random field $(Y_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ is m-dependent if the collections of random variables $(Y_{\boldsymbol{i}}, \boldsymbol{i} \in A)$ and $(Y_{\boldsymbol{i}}, \boldsymbol{i} \in B)$ are independent whenever $\inf\{\|\boldsymbol{a}-\boldsymbol{b}\|_\infty, \boldsymbol{a} \in A, \boldsymbol{b} \in B\} > m$. The proof of Theorem 1.6 will use the following tools.

(T.1) By Theorem 2.6 in [CS04], if I is a finite subset of $\mathbb{Z}^d$, $(Y_{\boldsymbol{i}})_{\boldsymbol{i}\in I}$ an m-dependent centered random field such that $\mathbb{E}\left[|Y_{\boldsymbol{i}}|^p\right] < +\infty$ for each $\boldsymbol{i} \in I$ and some $p \in (2,3]$ and $\operatorname{Var}\left(\sum_{\boldsymbol{i}\in I} Y_{\boldsymbol{i}}\right) = 1$, then

$$\delta\left(\sum_{\boldsymbol{i}\in I} Y_{\boldsymbol{i}}\right) \leqslant 75\,(10m+1)^{(p-1)d} \sum_{\boldsymbol{i}\in I} \mathbb{E}\left[|Y_{\boldsymbol{i}}|^p\right]. \tag{2.2.2}$$

(T.2) By Lemma 1 in [EMO07], for any two random variables Z and Z' and $p \geqslant 1$,

$$\delta(Z+Z') \leqslant 2\delta(Z) + \|Z'\|_p^{\frac{p}{p+1}}. \tag{2.2.3}$$

Let $(\varepsilon_{\boldsymbol{u}})_{\boldsymbol{u}\in\mathbb{Z}^d}$ be an i.i.d. random field and let $f\colon \mathbb{R}^{\mathbb{Z}^d} \to \mathbb{R}$ be a measurable function such that for each $\boldsymbol{i} \in \mathbb{Z}^d$, $X_{\boldsymbol{i}} = f\left((\varepsilon_{\boldsymbol{i}-\boldsymbol{u}})_{\boldsymbol{u}\in\mathbb{Z}^d}\right)$. Let $\gamma > 0$ and n_0 defined by (1.3.10).

Let $m := ([\|b_n\|_{\ell^2}] + 1)^\gamma$ and let us define

$$X_{\boldsymbol{i}}^{(m)} := \mathbb{E}\left[X_{\boldsymbol{i}} \mid \sigma\left(\varepsilon_{\boldsymbol{u}}, \boldsymbol{i} - m\mathbf{1} \preccurlyeq \boldsymbol{u} \preccurlyeq \boldsymbol{i} + m\mathbf{1}\right)\right]. \tag{2.2.4}$$

Since the random field $(\varepsilon_{\boldsymbol{u}})_{\boldsymbol{u}\in\mathbb{Z}^d}$ is independent, the following properties hold.

(P.1) The random field $\left(X_{\boldsymbol{i}}^{(m)}\right)_{\boldsymbol{i}\in\mathbb{Z}^d}$ is $(2m+1)$-dependent.

(P.2) The random field $\left(X_{\boldsymbol{i}}^{(m)}\right)_{\boldsymbol{i}\in\mathbb{Z}^d}$ is identically distributed and $\left\|X_{\boldsymbol{i}}^{(m)}\right\|_{p'} \leqslant \|X_{\mathbf{0}}\|_{p'}$.

(P.3) For any $(a_{\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d} \in \ell^2\left(\mathbb{Z}^d\right)$ and $q \geqslant 2$, the following inequality holds:

$$\begin{aligned}\left\|\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}}\left(X_{\boldsymbol{i}} - X_{\boldsymbol{i}}^{(m)}\right)\right\|_q &\leqslant \frac{14.5q}{\log q}\left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d} a_{\boldsymbol{i}}^2\right)^{1/2} \sum_{j\geqslant m} (4j+4)^{d/2}\,\|X_{\mathbf{0},j}\|_2 \\ &\quad + \frac{14.5q}{\log q}\left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d} |a_{\boldsymbol{i}}|^q\right)^{1/q} \sum_{j\geqslant m} (4j+4)^{d(1-1/q)}\,\|X_{\mathbf{0},j}\|_q.\end{aligned} \tag{2.2.5}$$

In order to prove (2.2.5), we follow the proof of Theorem 1.4 and start from the decomposition $X_{\boldsymbol{i}} - X_{\boldsymbol{i}}^{(m)} = \lim_{N\to+\infty} \sum_{j=m}^N X_{\boldsymbol{i},j}$ instead of (2.1.2).

Define $S_n^{(m)} := \sum_{\boldsymbol{i}\in\mathbb{Z}^d} b_{n,\boldsymbol{i}} X_{\boldsymbol{i}}^{(m)}$. An application of (T.2) to $Z := S_n^{(m)}\,\|b_n\|_{\ell^2}^{-1}\,\sigma^{-1}$ and $Z' := \left(S_n - S_n^{(m)}\right)\|b_n\|_{\ell^2}^{-1}\,\sigma^{-1}$ yields

$$\Delta_n \leqslant 2\delta\left(\frac{S_n^{(m)}}{\sigma\,\|b_n\|_{\ell^2}}\right) + \sigma^{-\frac{p}{p+1}} \frac{1}{\|b_n\|_{\ell^2}^{\frac{p}{p+1}}} \left\|S_n - S_n^{(m)}\right\|_p^{\frac{p}{p+1}}. \tag{2.2.6}$$

Moreover,

$$\delta\left(\frac{S_n^{(m)}}{\sigma\left\|b_n\right\|_{\ell^2}}\right)=\sup_{t\in\mathbb{R}}\left|\mathbb{P}\left\{\frac{S_n^{(m)}}{\sigma\left\|b_n\right\|_{\ell^2}}\leqslant t\right\}-\Phi\left(t\right)\right| \tag{2.2.7}$$

$$=\sup_{u\in\mathbb{R}}\left|\mathbb{P}\left\{\frac{S_n^{(m)}}{\left\|S_n^{(m)}\right\|_2}\leqslant u\right\}-\Phi\left(u\frac{\left\|S_n^{(m)}\right\|_2}{\sigma\left\|b_n\right\|_{\ell^2}}\right)\right| \tag{2.2.8}$$

$$\leqslant\delta\left(\frac{S_n^{(m)}}{\left\|S_n^{(m)}\right\|_2}\right)+\sup_{u\in\mathbb{R}}\left|\Phi\left(u\frac{\left\|S_n^{(m)}\right\|_2}{\sigma\left\|b_n\right\|_{\ell^2}}\right)-\Phi\left(u\right)\right|, \tag{2.2.9}$$

hence, by (P.1) and (T.1) applied with $Y_{\boldsymbol{i}}:=X_{\boldsymbol{i}}^{(m)}/\left\|S_n^{(m)}\right\|_2$, p' instead of p and $2m+1$ instead of m, we derive that

$$\Delta_n\leqslant(I)+(II)+(III) \tag{2.2.10}$$

where

$$(I):=150\left(20m+21\right)^{\left(p'-1\right)d}\sum_{\boldsymbol{i}\in\mathbb{Z}^d}\left|b_{n,\boldsymbol{i}}\right|^{p'}\left\|X_i^{(m)}\right\|_{p'}^{p'}\left\|S_n^{(m)}\right\|_2^{-p'}, \tag{2.2.11}$$

$$(II):=2\sup_{u\in\mathbb{R}}\left|\Phi\left(u\frac{\left\|S_n^{(m)}\right\|_2}{\sigma\left\|b_n\right\|_{\ell^2}}\right)-\Phi\left(u\right)\right|\text{ and} \tag{2.2.12}$$

$$(III):=\sigma^{-\frac{p}{p+1}}\frac{1}{\left\|b_n\right\|_{\ell^2}^{\frac{p}{p+1}}}\left\|S_n-S_n^{(m)}\right\|_p^{\frac{p}{p+1}}. \tag{2.2.13}$$

By (P.2) and the reversed triangular inequality, the term (I) can be bounded in the following way

$$(I)\leqslant150\left(20m+21\right)^{\left(p'-1\right)d}\left\|X_{\mathbf{0}}\right\|_{p'}^{p'}\left\|b_n\right\|_{\ell^{p'}}^{p'}\left(\left\|S_n\right\|_2-\left\|S_n-S_n^{(m)}\right\|_2\right)^{-p'} \tag{2.2.14}$$

and by (P.3) with $q=2$, we obtain that

$$\left(\left\|S_n\right\|_2-\left\|S_n-S_n^{(m)}\right\|_2\right)^{-p'}\leqslant\left(\left\|S_n\right\|_2-29\left(\log 2\right)^{-1}m^{-\alpha}\left\|b_n\right\|_{\ell^2}C_2\left(\alpha\right)\right)^{-p'}. \tag{2.2.15}$$

By (1.3.8), we have

$$\frac{\left\|S_n\right\|_2^2}{\left\|b_n\right\|_{\ell^2}^2}=\sigma^2+\varepsilon_n, \tag{2.2.16}$$

and we eventually get

$$(I)\leqslant150\left(20m+21\right)^{\left(p'-1\right)d}\left\|X_{\mathbf{0}}\right\|_{p'}^{p'}\left(\frac{\left\|b_n\right\|_{\ell^{p'}}}{\left\|b_n\right\|_{\ell^2}}\right)^{p'}\left(\sqrt{\sigma^2+\varepsilon_n}-29\left(\log 2\right)^{-1}m^{-\alpha}C_2\left(\alpha\right)\right)^{-p'}.$$

Since $n\geqslant n_0$, we derive, in view of (1.3.10),

$$(I)\leqslant150\left(20m+21\right)^{\left(p'-1\right)d}\left\|X_{\mathbf{0}}\right\|_{p'}^{p'}\left(\frac{\left\|b_n\right\|_{\ell^{p'}}}{\left\|b_n\right\|_{\ell^2}}\right)^{p'}\left(\sigma/2\right)^{-p'} \tag{2.2.17}$$

In order to bound (II), we argue as in [YWLH12] (p. 456). Doing similar computations as in [EM14] (p. 272), we obtain that

$$(II) \leqslant (2\pi e)^{-1/2} \left(\inf_{k\geqslant 1} a_k\right)^{-1} \left|a_n^2 - 1\right|, \tag{2.2.18}$$

where $a_n := \left\|S_n^{(m)}\right\|_2 \sigma^{-1} \|b_n\|_{\ell^2}^{-1}$. Observe that for any n, by (P.3),

$$a_n \geqslant \frac{\|S_n\|_2 - \left\|S_n - S_n^{(m)}\right\|_2}{\sigma \|b_n\|_{\ell^2}} \geqslant \frac{\sqrt{\sigma^2+\varepsilon_n} - 29 (\log 2)^{-1} C_2(\alpha) m^{-\alpha}}{\sigma} \tag{2.2.19}$$

and using again (P.3) combined with Theorem 1.4 for $p = q = 2$,

$$\left|a_n^2 - 1\right| = \left|\frac{\left\|S_n^{(m)}\right\|_2^2}{\sigma^2 \|b_n\|_{\ell^2}^2} - 1\right| \tag{2.2.20}$$

$$\leqslant \left|\frac{\|S_n\|_2^2}{\sigma^2 \|b_n\|_{\ell^2}^2} - 1\right| + \frac{\left|\left\|S_n^{(m)}\right\|_2^2 - \|S_n\|_2^2\right|}{\sigma^2 \|b_n\|_{\ell^2}^2} \tag{2.2.21}$$

$$\leqslant \frac{|\varepsilon_n|}{\sigma^2} + \frac{\left|\left\|S_n^{(m)}\right\|_2 - \|S_n\|_2\right| \left(\left\|S_n^{(m)}\right\|_2 + \|S_n\|_2\right)}{\sigma^2 \|b_n\|_{\ell^2}^2} \tag{2.2.22}$$

$$\leqslant \frac{|\varepsilon_n|}{\sigma^2} + \frac{\left\|S_n^{(m)} - S_n\right\|_2 \left(\left\|S_n^{(m)}\right\|_2 + \|S_n\|_2\right)}{\sigma^2 \|b_n\|_{\ell^2}^2} \tag{2.2.23}$$

$$\leqslant \frac{|\varepsilon_n|}{\sigma^2} + 40 (\log 2)^{-1} \frac{m^{-\alpha}}{\sigma^2} C_2(\alpha)^2. \tag{2.2.24}$$

This leads to the estimate

$$(II) \leqslant \frac{(2\pi e)^{-1/2}}{\sqrt{\sigma^2+\varepsilon_n} - 29 (\log 2)^{-1} C_2(\alpha) m^{-\alpha}} \left(\frac{|\varepsilon_n|}{\sigma} + 40 (\log 2)^{-1} \frac{m^{-\alpha}}{\sigma} C_2(\alpha)^2\right), \tag{2.2.25}$$

and since $n \geqslant n_0$, we derive, in view of (1.3.10),

$$(II) \leqslant \left(2\frac{|\varepsilon_n|}{\sigma^2} + 80 (\log 2)^{-1} \frac{\|b_n\|_{\ell^2}^{-\gamma\alpha}}{\sigma^2} C_2(\alpha)^2\right) (2\pi e)^{-1/2}. \tag{2.2.26}$$

The estimate of (III) relies on (P.3):

$$(III) \leqslant \sigma^{-\frac{p}{p+1}} \left(\frac{14.5p}{\log p} \sum_{j\geqslant m} (4j+4)^{d/2} \|X_{\mathbf{0},j}\|_2\right)^{\frac{p}{p+1}}$$
$$+ \sigma^{-\frac{p}{p+1}} \|b_n\|_{\ell^2}^{-\frac{p}{p+1}} \|b_n\|_{\ell^p}^{\frac{p}{p+1}} \left(\frac{14.5p}{\log p} \sum_{j\geqslant m} (4j+4)^{d(1-1/p)} \|X_{\mathbf{0},j}\|_p\right)^{\frac{p}{p+1}} \tag{2.2.27}$$

hence

$$(III) \leqslant \left(\frac{14.5p}{\sigma \log p} 4^{d/2} \|b_n\|_{\ell^2}^{-\gamma\alpha} C_2(\alpha)\right)^{\frac{p}{p+1}} + \left(\frac{\|b_n\|_{\ell^p}}{\sigma \|b_n\|_{\ell^2}} \frac{14.5p}{\log p} 4^{d(1-1/p)} \|b_n\|_{\ell^2}^{-\gamma\beta} C_p(\beta)\right)^{\frac{p}{p+1}}. \tag{2.2.28}$$

The combination of (2.2.10), (2.2.17), (2.2.26) and (2.2.28) gives (1.3.11).

2.3. **Proof of Theorem 1.9.** Since the random variables $X_{\boldsymbol{i}}$ are centered, we derive by definition of $g_n(\boldsymbol{x})$ that

$$(nh_n)^{d/2}\left(g_n(\boldsymbol{x})-\mathbb{E}\left[g_n(\boldsymbol{x})\right]\right)=(nh_n)^{d/2}\frac{\sum_{\boldsymbol{i}\in\Lambda_n}X_{\boldsymbol{i}}K\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right)}{\sum_{\boldsymbol{i}\in\Lambda_n}K\left(\frac{\boldsymbol{x}-\boldsymbol{i}/n}{h_n}\right)}. \tag{2.3.1}$$

We define

$$b_{n,\boldsymbol{i}}=K\left(\frac{1}{h_n}\left(\boldsymbol{x}-\frac{\boldsymbol{i}}{n}\right)\right),\quad \boldsymbol{i}\in\Lambda_n \tag{2.3.2}$$

and $b_{n,\boldsymbol{i}}=0$ otherwise. Denote $b_n=(b_{n,\boldsymbol{i}})_{\boldsymbol{i}\in\mathbb{Z}^d}$ and $\|b_n\|_{\ell^2}:=\left(\sum_{\boldsymbol{i}\in\mathbb{Z}^d}b_{n,\boldsymbol{i}}^2\right)^{1/2}$. In this way, by (2.3.1) and (1.3.21),

$$\frac{1}{\|K\|_{\mathbb{L}^2(\mathbb{R}^d)}\sigma}(nh_n)^{d/2}\left(g_n(\boldsymbol{x})-\mathbb{E}\left[g_n(\boldsymbol{x})\right]\right)=\frac{1}{\sigma}\sum_{\boldsymbol{i}\in\mathbb{Z}^d}b_{n,\boldsymbol{i}}X_{\boldsymbol{i}}\|b_n\|_{\ell^2}^{-1}A_n. \tag{2.3.3}$$

Applying (T.2) to $Z=\sum_{\boldsymbol{i}\in\mathbb{Z}^d}b_{n,\boldsymbol{i}}X_{\boldsymbol{i}}\|b_n\|_{\ell^2}^{-1}$ and $Z'=\sum_{\boldsymbol{i}\in\mathbb{Z}^d}b_{n,\boldsymbol{i}}X_{\boldsymbol{i}}\|b_n\|_{\ell^2}^{-1}\sigma^{-1}(A_n-1)$ and using Theorem 1.4, we derive that

$$\widetilde{\Delta_n}\leqslant c_p\Delta_n'+c_p\left(\sigma^{-1}C_2(\alpha)+C_p(\beta)\right)^{\frac{p}{p+1}}|A_n-1|^{\frac{p}{p+1}}, \tag{2.3.4}$$

where

$$\Delta_n'=\sup_{t\in\mathbb{R}}\left|\mathbb{P}\left\{Z\leqslant t\right\}-\Phi\left(\frac{t}{\sigma}\right)\right|. \tag{2.3.5}$$

We then use Theorem 1.6 to handle Δ_n' (which is allowed, by (A)). Using boundedness of K, we control the ℓ^p and $\ell^{p'}$ norms by a constant times the ℓ^2-norm. This ends the proof of Theorem 1.9.

Acknowledgments This research was supported by the grand SFB 823.

The author would like to thank the referees for many suggestions which improved the presentation of the paper.

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