



**HAL**  
open science

## Reduced animal model for marker assisted selection using best linear unbiased prediction

Rjc Cantet, C Smith

► **To cite this version:**

Rjc Cantet, C Smith. Reduced animal model for marker assisted selection using best linear unbiased prediction. *Genetics Selection Evolution*, 1991, 23 (3), pp.221-233. hal-00893877

**HAL Id: hal-00893877**

**<https://hal.science/hal-00893877>**

Submitted on 11 May 2020

**HAL** is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

# Reduced animal model for marker assisted selection using best linear unbiased prediction

RJC Cantet<sup>\*,\*\*</sup>, C Smith

*University of Guelph, Centre for Genetic Improvement of Livestock,  
Department of Animal and Poultry Science, Guelph, Ontario, N1G 2W1, Canada*

(Received 15 October 1990; accepted 11 April 1991)

**Summary** – A reduced animal model (RAM) version of the animal model (AM) incorporating independent marked quantitative trait loci (MQTL's) of Fernando and Grossman (1989) is presented. Both AM and RAM permit obtaining Best Linear Unbiased Predictions of MQTL effects plus the remaining portion of the breeding value that is not accounted for by independent MQTL's. RAM reduces computational requirements by a reduction in the size of the system of equations. Non-parental MQTL effects are expressed as a linear function of parental MQTL effects using marker information and the recombination rate ( $r$ ) between the marker locus and the MQTL. The resulting fraction of the MQTL variance that is explained by the regression on parental MQTL effects is  $2[(1-r)^2 + r^2]/2$  when the individual is not inbred and both parents are known. Formulae are obtained to simplify the computations when backsolving for non-parental MQTL and breeding values in case all non-parents have one record. A small numerical example is also presented.

**maker assisted selection / best linear unbiased prediction / reduced animal model / genetic marker**

**Résumé** – Un modèle animal réduit pour la sélection assistée par marqueurs avec BLUP. Une version du modèle animal réduit (RAM) basée sur le modèle animal (AM) de Fernando et Grossman (1989) avec loci indépendants de caractères quantitatifs marqués (MQTL) est présentée. Dans les 2 cas, RAM et AM, on obtient les meilleures prédictions linéaires sans biais (BLUP) des effets des MQTL en plus de la portion restante de la valeur génétique inexpliquée par les MQTL indépendants. L'emploi de RAM diminue les exigences de calcul par une réduction de la taille du système d'équations. Les effets des MQTL non-parentaux sont exprimés sous la forme d'une fonction linéaire des effets des MQTL parentaux à l'aide de l'information provenant du marqueur et du taux de recombinaison ( $r$ ) entre le locus marqueur et le MQTL. La proportion résultante de la variance du MQTL

\* On leave from : Departamento de Zootecnia, Facultad de Agronomía, Universidad de Buenos Aires, Argentina

\*\* Correspondence and reprints

*expliquée par la régression des effets des MQTL parentaux est donnée par l'expression  $2[(1-r)^2 + r^2]/2$  dans le cas d'un individu non consanguin avec parents connus. Des formules sont dérivées pour simplifier les calculs lorsque l'on résout pour les effets des MQTL et des valeurs génétiques non parentaux dans le cas où tous les individus non parents possèdent une seule observation. Un exemple numérique est également donné.*

**sélection assistée par marqueurs / BLUP / modèle animal réduit / marqueur génétique**

## INTRODUCTION

In a recent paper, Fernando and Grossman (1989) obtained best linear unbiased predictors (Henderson, 1984) of the additive effects for alleles at a marked quantitative trait locus (MQTL) and of the remaining portion of the breeding value. They used an animal model (AM; Henderson, 1984) under a purely additive mode of inheritance. Letting  $p$  be the number of fixed effects in the model,  $n$  the number of animals in the pedigree file and  $m$  the number of MQTL's, the number of equations in the system for this AM is  $p + n(2m + 1)$ . For large  $m$ ,  $n$  or both, solving such a system may not always be feasible. The reduced animal model (RAM; Quaas and Pollak, 1980) is an equivalent model, in the sense of Henderson (1985), to the AM and provides the same results, but with a smaller number of equations to be solved. In this paper, the RAM version of the model of Fernando and Grossman (1989) is obtained. The resulting system of equation is of order  $p + s(2m + 1)$ ,  $s$  being the number of parents. In general  $s$  is much smaller than  $n$ . Therefore, the advantage due to the reduction in the number of equations by using RAM is considerable. A numerical example is included to illustrate the application.

## THEORY

For simplicity, derivations are presented for a model with one MQTL. The extension to the case of 2 or more independent MQTL's is covered in the section entitled *More than one MQTL*.

In the notation of Fernando and Grossman (1989),  $M_i^p$  and  $M_i^m$  are alleles at the marker locus that individual  $i$  inherited from its paternal ( $p$ ) and its maternal ( $m$ ) parents, and  $v_i^p$  and  $v_i^m$  are the additive effects of the paternal and maternal MQTL's, respectively. The recombination frequency between the marker allele and the MQTL is denoted as  $r$ . We will use the expression "breeding value" to refer to the additive effects of all genes that affect the trait excluding the MQTL(s).

### **Matrix expressions for the animal model with genetic marker information**

A matrix version of equation (3) in Fernando and Grossman (1989) is :

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \mathbf{W}\mathbf{v} + \mathbf{e} \quad (1)$$

where  $\mathbf{y}$  is an  $n \times 1$  vector of records,  $\mathbf{X}$ ,  $\mathbf{Z}$  and  $\mathbf{W}$  are  $n \times p$ ,  $n \times n$  and  $n \times 2n$  incidence matrices which relate data to the unknown vector of fixed effects  $\boldsymbol{\beta}$ , the

random vector of additive breeding values  $\mathbf{u}$  and the random vector  $\mathbf{v}$  of additive effects of the individual MQTL effects, respectively. The  $2n \times 1$  vector  $\mathbf{v}$  is ordered within animal such that  $v_i^p$  always precedes  $v_i^m$ . The matrices  $\mathbf{Z}$  and  $\mathbf{W}$  will have zero rows for animals that do not have records on themselves but that are related to animals with records. Non-zero rows of  $\mathbf{Z}$  and  $\mathbf{W}$  have 1 and 2 elements equal to 1, respectively, with the remaining elements being zero. First and second moments of  $\mathbf{y}$  are given by :

$$E(\mathbf{y}) = \mathbf{X}\boldsymbol{\beta} \quad \text{Var}(\mathbf{y}) = \mathbf{ZAZ}'\sigma_A^2 + \mathbf{W}\mathbf{G}_v\mathbf{W}'\sigma_v^2 + \mathbf{I}\sigma_e^2$$

where  $\mathbf{A}\sigma_A^2$  and  $\mathbf{G}_v\sigma_v^2$  are the variance-covariance matrices of  $\mathbf{u}$  and  $\mathbf{v}$ , respectively. The scalars  $\sigma_A^2$ ,  $\sigma_v^2$  and  $\sigma_e^2$  are the variance components of the additive effects of breeding values, the MQTL additive effects and of the environmental effects.

RAM requires partitioning the data vector  $\mathbf{y}$  into records of individuals with progeny ( $\mathbf{y}_P$ ; parents) and records of individuals without progeny ( $\mathbf{y}_N$ ; non-parents) so that  $\mathbf{y}' = [\mathbf{y}'_P, \mathbf{y}'_N]$ . A conformable partition can be used in  $\mathbf{X}$ ,  $\mathbf{Z}$ ,  $\mathbf{W}$ ,  $\mathbf{u}$ ,  $\mathbf{v}$  and  $\mathbf{e}$ . Using this idea (1) can be written as :

$$\begin{bmatrix} \mathbf{y}_P \\ \mathbf{y}_N \end{bmatrix} = \begin{bmatrix} \mathbf{X}_P \\ \mathbf{X}_N \end{bmatrix} \boldsymbol{\beta} + \begin{bmatrix} \mathbf{Z}_P & 0 \\ 0 & \mathbf{Z}_N \end{bmatrix} \begin{bmatrix} \mathbf{u}_P \\ \mathbf{u}_N \end{bmatrix} + \begin{bmatrix} \mathbf{W}_P & 0 \\ 0 & \mathbf{W}_N \end{bmatrix} \begin{bmatrix} \mathbf{v}_P \\ \mathbf{v}_N \end{bmatrix} + \begin{bmatrix} \mathbf{e}_P \\ \mathbf{e}_N \end{bmatrix} \quad (2)$$

To obtain RAM,  $\mathbf{u}_N$  and  $\mathbf{v}_N$  should be expressed as linear functions of  $\mathbf{u}_P$  and  $\mathbf{v}_P$ , respectively. Since an individual's breeding value can be described as the average of the breeding value of its parents plus an independently distributed Mendelian sampling residual ( $\boldsymbol{\phi}$ ) (Quaas and Pollak, 1980), for  $\mathbf{u}_N$  we can write :

$$\mathbf{u}_N = \mathbf{P}\mathbf{u}_P + \boldsymbol{\phi}_N \quad (3)$$

where  $\mathbf{P}$  is an  $(n - s) \times s$  matrix relating non-parental to parental breeding values. Each row of  $\mathbf{P}$  contains at most two 0.5 values in the columns pertaining to the  $BV$ 's of the sire and of the dam. Now,  $E(\boldsymbol{\phi}) = 0$  and  $\text{Var}(\boldsymbol{\phi}) = \mathbf{D}_A\sigma_A^2$ , where  $\mathbf{D}_A$  is a diagonal matrix with diagonal elements equal to :

- 1 - 0.25( $a_{ss} + a_{dd}$ ), if both sire and dam of the non-parent are known
- 1 - 0.25 $a_{ss}$ , if only the sire is known
- 1 - 0.25 $a_{dd}$ , if only the dam is known
- 1, if both parents are unknown

with  $a_{ss}$  and  $a_{dd}$  being the diagonal elements of  $\mathbf{A}$  corresponding to the sire and the dam, respectively.

A scalar version of the relationship between  $v_N$  and  $v_P$  can be obtained from equations (8a) and (8b) in Fernando and Grossman (1989) and these are :

$$\begin{aligned} v_0^p &= b_1 v_s^p + b_2 v_s^m + \varepsilon_0^p \\ v_0^m &= b_3 v_d^p + b_4 v_d^m + \varepsilon_0^m \end{aligned}$$

The subscripts  $o, s$  and  $d$  denote the individual, its sire and its dam, respectively. The coefficients  $b'_i s$  are either  $1 - r$  or  $r$  according to any of these 4 possible patterns of inheritance of the marker alleles :

<i>Paternal marker</i>	<i>Maternal marker</i>	$b_1$	$b_2$	$b_3$	$b_4$
$M_s^p$	$M_d^p$	$1 - r$	$r$	$1 - r$	$r$
$M_s^p$	$M_d^m$	$1 - r$	$r$	$r$	$1 - r$
$M_s^m$	$M_d^p$	$r$	$1 - r$	$1 - r$	$r$
$M_s^m$	$M_d^m$	$r$	$1 - r$	$r$	$1 - r$

The above developments lead us to the following relationship between  $\mathbf{v}_N$  and  $\mathbf{v}_p$  :

$$\mathbf{v}_N = \mathbf{F}\mathbf{v}_p + \boldsymbol{\varepsilon} \tag{4}$$

The  $2(n - s) \times 2s$  matrix  $\mathbf{F}$  relates the additive effects of the MQTL of non-parents to the additive effects of the MQTL of parents and  $\boldsymbol{\varepsilon}$  is the vector with element  $i$  equal to residual  $\boldsymbol{\varepsilon}_0^p$  and element  $i + 1$  equal to the residual  $\boldsymbol{\varepsilon}_0^m$ . Each row of  $\mathbf{F}$ , contains at most, 2 non-zero elements : the  $b'_i s$ . Let  $i$  and  $k$  be the row indices for the MQTL marked by  $M_0^p$  and  $M_0^m$  respectively. Let  $j$  and  $j + 1$  be the column indices corresponding to the additive effects of the MQTL for the sire that transmits  $i : j$  refers to the paternal grandsire and  $j + 1$  to the paternal granddam. Also, let  $l$  and  $l + 1$  be the column indices corresponding to the dam that transmits  $i + 1 : l$  corresponds to the maternal grandsire and  $l + 1$  to the maternal granddam. Then  $\mathbf{F}_{ij} = b_1$ ,  $\mathbf{F}_{i,j+1} = b_2$ ,  $\mathbf{F}_{k,l} = b_3$  and  $\mathbf{F}_{k,l+1} = b_4$ . All remaining elements of  $\mathbf{F}$  are 0. When marker information is unavailable,  $r$  is taken to be 0.5 (Fernando and Grossman, 1989) and all  $b'_i s$  are 0.5. To exemplify, consider individuals 1 (male), 2 (female) and 3 (progeny of 1 and 2). Animals 1 and 2 are unrelated and 3 has paternal and maternal marker alleles originating from the dams of 1 and 2, namely alleles  $M_1^d$  and  $M_2^d$ , respectively. Then,  $\mathbf{v} = [v_1^p, v_1^m, v_2^p, v_2^m, v_3^p, v_3^m]'$ , with  $\mathbf{v}^p = [v_1^p, v_1^m, v_2^p, v_2^m]'$  and  $\mathbf{v}^M = [v_3^p, v_3^m]'$ . The matrix  $\mathbf{W}$  is :

$$\mathbf{W} = \begin{bmatrix} \mathbf{W}_P & 0_{2 \times 2} \\ 0_{1 \times 4} & \mathbf{W}_N \end{bmatrix} = \begin{bmatrix} 1 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 1 \end{bmatrix}$$

For  $r = 0.2$ , the matrix  $\mathbf{F}$  is  $2 \times 4$  and equal to :

$$\mathbf{F} = \begin{bmatrix} 0.2 & 0.8 & 0.0 & 0.0 \\ 0.0 & 0.0 & 0.2 & 0.8 \end{bmatrix}$$

The residuals  $\boldsymbol{\varepsilon}$  have  $E(\boldsymbol{\varepsilon}) = 0$  and  $\text{Var}(\boldsymbol{\varepsilon}) = \mathbf{G}_\boldsymbol{\varepsilon}\sigma_v^2$ . Fernando and Grossman (1989) showed that  $\mathbf{G}_\boldsymbol{\varepsilon}\sigma_v^2$  is diagonal with non-zero elements equal to  $\text{Var}(\boldsymbol{\varepsilon}_0^p) = 2r(1 - r)(1 - f_s)\sigma_v^2$  and  $\text{Var}(\boldsymbol{\varepsilon}_0^m) = 2r(1 - r)(1 - f_d)\sigma_v^2$ , where  $f_s$ , and  $f_d$  are the inbreeding coefficients at the MQTL of the sire and of the dam, respectively. They express the probability that the paternal and maternal alleles of an individual for

a given MQTL are the same. These  $f$ 's are the off-diagonal elements in the  $2 \times 2$  diagonal blocks of the matrix  $\mathbf{G}_v$  (Fernando and Grossman, 1989).

Using (3) and (4) in (2) gives :

$$\begin{bmatrix} \mathbf{y}_P \\ \mathbf{y}_N \end{bmatrix} = \begin{bmatrix} \mathbf{X}_P \\ \mathbf{X}_N \end{bmatrix} \boldsymbol{\beta} + \begin{bmatrix} \mathbf{Z}_P & 0 \\ \mathbf{Z}_N \mathbf{P} & \mathbf{Z}_N \end{bmatrix} \begin{bmatrix} \mathbf{u}_P \\ \boldsymbol{\phi} \end{bmatrix} + \begin{bmatrix} \mathbf{W}_P & 0 \\ \mathbf{W}_N \mathbf{F} & \mathbf{W}_N \end{bmatrix} \begin{bmatrix} \mathbf{v}_P \\ \boldsymbol{\varepsilon} \end{bmatrix} + \begin{bmatrix} \mathbf{e}_P \\ \mathbf{e}_N \end{bmatrix} \quad (5)$$

or

$$\begin{bmatrix} \mathbf{y}_P \\ \mathbf{y}_N \end{bmatrix} = \begin{bmatrix} \mathbf{X}_P \\ \mathbf{X}_N \end{bmatrix} \boldsymbol{\beta} + \begin{bmatrix} \mathbf{Z}_P \\ \mathbf{Z}_N \mathbf{P} \end{bmatrix} \mathbf{u}_P + \begin{bmatrix} \mathbf{W}_P \\ \mathbf{W}_N \mathbf{F} \end{bmatrix} \mathbf{v}_P + \begin{bmatrix} \mathbf{e}_P \\ \mathbf{e}_N + \mathbf{Z}_N \boldsymbol{\phi} + \mathbf{W}_N \boldsymbol{\varepsilon} \end{bmatrix} \quad (6)$$

On letting  $\mathbf{e}^* = \mathbf{e}_N + \mathbf{Z}_N \boldsymbol{\phi} + \mathbf{W}_N \boldsymbol{\varepsilon}$ , we have that :

$$\text{Var} \begin{bmatrix} \mathbf{e}_P \\ \mathbf{e}^* \end{bmatrix} = \begin{bmatrix} \mathbf{I} & 0 \\ 0 & \mathbf{Q} \end{bmatrix} \sigma_e^2$$

where  $\mathbf{Q} = \mathbf{I}_{(n-s)} + \mathbf{Z}_N \mathbf{D}_A \mathbf{Z}'_N \boldsymbol{\alpha}_A + \mathbf{W}_N \mathbf{G}_\boldsymbol{\varepsilon} \mathbf{W}'_N \boldsymbol{\alpha}_v$ ,  $\boldsymbol{\alpha}_A = \sigma_A^2 / \sigma_e^2$  and  $\boldsymbol{\alpha}_v = \sigma_v^2 / \sigma_e^2$ . Mixed model equations for (6) are :

$$\begin{bmatrix} \mathbf{X}'_P \mathbf{X}_P + \mathbf{X}'_N \mathbf{Q}^{-1} \mathbf{X}_N & \mathbf{X}'_P \mathbf{Z}_P + \mathbf{X}'_N \mathbf{Q}^{-1} \mathbf{Z}_N \mathbf{P} & \mathbf{X}'_P \mathbf{W}_P + \mathbf{X}'_N \mathbf{Q}^{-1} \mathbf{W}_N \mathbf{F} \\ \mathbf{Z}'_P \mathbf{X}_P + \mathbf{P}' \mathbf{Z}'_N \mathbf{Q}^{-1} \mathbf{X}_N & \mathbf{Z}'_P \mathbf{Z}_P + \mathbf{P}' \mathbf{Z}'_N \mathbf{Q}^{-1} \mathbf{Z}_N \mathbf{P} + \mathbf{A}_P^{-1} \boldsymbol{\alpha}_A & \mathbf{Z}'_P \mathbf{W}_P + \mathbf{P}' \mathbf{Z}'_N \mathbf{Q}^{-1} \mathbf{W}_N \mathbf{F} \\ \mathbf{W}'_P \mathbf{X}_P + \mathbf{F}' \mathbf{W}'_N \mathbf{Q}^{-1} \mathbf{X}_N & \mathbf{W}'_P \mathbf{Z}_P + \mathbf{F}' \mathbf{W}'_N \mathbf{Q}^{-1} \mathbf{Z}_N \mathbf{P} & \mathbf{W}'_P \mathbf{W}_P + \mathbf{F}' \mathbf{W}'_N \mathbf{Q}^{-1} \mathbf{W}_N \mathbf{F} + \mathbf{G}_v^{-1} \boldsymbol{\alpha}_v \end{bmatrix} \begin{bmatrix} \hat{\boldsymbol{\beta}} \\ \hat{\mathbf{u}}_P \\ \hat{\mathbf{v}}_P \end{bmatrix} = \begin{bmatrix} \mathbf{X}'_P \mathbf{y}_P + \mathbf{X}'_N \mathbf{Q}^{-1} \mathbf{y}_N \\ \mathbf{Z}'_P \mathbf{y}_P + \mathbf{P}' \mathbf{Z}'_N \mathbf{Q}^{-1} \mathbf{y}_N \\ \mathbf{W}'_P \mathbf{y}_P + \mathbf{F}' \mathbf{W}'_N \mathbf{Q}^{-1} \mathbf{y}_N \end{bmatrix} \quad (7)$$

The matrices  $\mathbf{A}_P$  and  $\mathbf{G}_v$  are the corresponding submatrices of  $\mathbf{A}$  and  $\mathbf{G}_v$  that belong to parents. Equations (7) give the solutions for RAM with genetic markers. Of practical importance is the case where all non-parents have only one record so that  $\mathbf{Z}_N = \mathbf{I}$ . Then,  $\mathbf{W}_N \mathbf{G}_\boldsymbol{\varepsilon} \mathbf{W}'_N$  and  $\mathbf{Q}^{-1}$  are diagonal (see *Appendix A*). The diagonal elements of  $\mathbf{W}_N \mathbf{G}_\boldsymbol{\varepsilon} \mathbf{W}'_N$  are derived in *Appendix A* and they are equal to :

- $2r(1-r)(2-f_s-f_d)$ , when both the sire and the dam of the non-parent are known
- $2r(1-r)(1-f_s) + 1$ , when only the sire is known
- $2r(1-r)(1-f_d) + 1$ , when only the dam is known
- 2, if both the sire and the dam of the non-parent are unknown.

If there is zero probability that the paternal and maternal alleles at the MQTL of parent  $p$  are the same (*ie*  $f_p = 0$ ), the contribution to the diagonal element of  $\mathbf{W}_N \mathbf{G}_\boldsymbol{\varepsilon} \mathbf{W}'_N$  is  $2r(1-r)$  (if marker information is available) or  $1/2$  (if marker information is unavailable). This occurs because, in the absence of marker information, there is equal probability of receiving the MQTL from the grandsire and from the granddam, and  $r = 0.5$  (Fernando and Grossman, 1989).

A further simplification to (7) occurs when parents do not have records so that  $\mathbf{Z}_P$  and  $\mathbf{W}_P$  are zero and the model becomes a sire-dam model. A program for RAM, such as the one presented by Schaeffer and Wilton (1987) and modified to include marker information can be employed to solve equations (7).

### *More than one MQTL*

Multiple MQTL ( $k$ , say) can be dealt with assuming independence by the following modification of model (1) :

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + (\mathbf{j}'_k \otimes \mathbf{W}) \begin{bmatrix} \mathbf{v}_1 \\ \cdots \\ \mathbf{v}_k \end{bmatrix} + \mathbf{e} \quad (8)$$

where  $\mathbf{j}_k$  is a  $k \times 1$  vector with all elements equal to 1. We will assume that  $\text{Var}(\mathbf{v}_i) = \mathbf{G}_{v_i}\sigma_{v_i}^2$  and  $\text{Cov}(\mathbf{v}_i, \mathbf{v}_{i'}) = 0$ . For  $k = 2$  and letting  $\mathbf{Q}^* = \mathbf{I}_{(n-s)} + \mathbf{Z}'_N \mathbf{D}_A \mathbf{Z}'_N \boldsymbol{\alpha}_A + \mathbf{W}'_N (\mathbf{G}_{\boldsymbol{\varepsilon}1} \boldsymbol{\alpha}_{v1} + \mathbf{G}_{\boldsymbol{\varepsilon}2} \boldsymbol{\alpha}_{v2}) \mathbf{W}'_N$ , RAM equations for (8) are :

### *Backsolving for non-parents*

After solving for fixed effects, parental breeding values and parental effects of the MQTL, the breeding values and additive MQTL effects of non-parents can be calculated. This is accomplished by writing the equations for  $\hat{\boldsymbol{\phi}}$  and  $\hat{\boldsymbol{\varepsilon}}$  from the mixed model equations of (5). This gives :

$$\begin{aligned} \mathbf{Z}'_N \mathbf{X}_N \hat{\boldsymbol{\beta}} + \mathbf{Z}'_N \mathbf{Z}_N \mathbf{P} \hat{\mathbf{u}}_P + (\mathbf{Z}'_N \mathbf{Z}_N + \mathbf{D}_A^{-1} \boldsymbol{\alpha}_A) \hat{\boldsymbol{\phi}} + \mathbf{Z}'_N \mathbf{W}_N \mathbf{F} \hat{\mathbf{v}}_P + \mathbf{Z}'_N \mathbf{W}_N \hat{\boldsymbol{\varepsilon}} &= \mathbf{Z}'_N \mathbf{y}_N \\ \mathbf{W}'_N \mathbf{X}_N \hat{\boldsymbol{\beta}} + \mathbf{W}'_N \mathbf{Z}_N \mathbf{P} \hat{\mathbf{u}}_P + \mathbf{W}'_N \mathbf{Z}_N \hat{\boldsymbol{\phi}} + \mathbf{W}'_N \mathbf{W}_N \mathbf{F} \hat{\mathbf{v}}_P & \\ + (\mathbf{W}'_N \mathbf{W}_N + \mathbf{G}_{\boldsymbol{\varepsilon}}^{-1} \boldsymbol{\alpha}_v) \hat{\boldsymbol{\varepsilon}} &= \mathbf{W}'_N \mathbf{y}_N \end{aligned}$$

and after a little algebra :

$$\begin{aligned} \begin{bmatrix} \mathbf{Z}'_N \mathbf{Z}_N + \mathbf{D}_A^{-1} \boldsymbol{\alpha}_A & \mathbf{Z}'_N \mathbf{W}_N \\ \mathbf{W}'_N \mathbf{Z}_N & \mathbf{W}'_N \mathbf{W}_N + \mathbf{G}_{\boldsymbol{\varepsilon}}^{-1} \boldsymbol{\alpha}_v \end{bmatrix} \begin{bmatrix} \hat{\boldsymbol{\phi}} \\ \hat{\boldsymbol{\varepsilon}} \end{bmatrix} & \\ = \begin{bmatrix} \mathbf{Z}'_N \\ \mathbf{W}'_N \end{bmatrix} (\mathbf{y}_N - \mathbf{X}_N \hat{\boldsymbol{\beta}} - \mathbf{Z}_N \mathbf{P} \hat{\mathbf{u}}_P - \mathbf{W}_N \mathbf{F} \hat{\mathbf{v}}_P) & \quad (10) \end{aligned}$$

*Appendix B* shows how to obtain solutions of equations (10), when all non-parents have one record, by solving  $(n-s)$  independent systems of order 2. Using the predictors obtained from (7) and (10) in (3) and (4), solutions for non-parents are :

$$\hat{\mathbf{u}}_N = \mathbf{P} \hat{\mathbf{u}}_P + \hat{\boldsymbol{\phi}}_N \quad \hat{\mathbf{v}}_N = \mathbf{F} \hat{\mathbf{v}}_P + \hat{\boldsymbol{\varepsilon}}$$

$$\begin{aligned}
 \begin{bmatrix}
 X'_P X_P + X'_N Q_*^{-1} X_N \\
 X_P + P' Z'_N Q_*^{-1} X_N \\
 X_P + F'_1 W'_N Q_*^{-1} X_N \\
 X_P + F'_2 W'_N Q_*^{-1} X_N
 \end{bmatrix}
 &=
 \begin{bmatrix}
 X'_P Z_P + X'_N Q_*^{-1} Z_N P \\
 Z'_P Z_P + P' Z'_N Q_*^{-1} Z_N P + A^{-1}_{PP} \alpha_A \\
 W'_P Z_P + F'_1 W'_N Q_*^{-1} Z_N P \\
 W'_P Z_P + F'_2 W'_N Q_*^{-1} Z_N P
 \end{bmatrix}
 \\
 &=
 \begin{bmatrix}
 \hat{\beta} \\
 \hat{u}_P \\
 \hat{v}_{1P} \\
 \hat{v}_{2P}
 \end{bmatrix}
 =
 \begin{bmatrix}
 X'_P y_P + X'_N Q_*^{-1} y_N \\
 Z'_P y_P + P' Z'_N Q_*^{-1} y_N \\
 W'_P y_P + F'_1 W'_N Q_*^{-1} y_N \\
 W'_P y_P + F'_2 W'_N Q_*^{-1} y_N
 \end{bmatrix}
 \\
 &=
 \begin{bmatrix}
 X'_P W_P + X'_N Q_*^{-1} W_N F_2 \\
 Z'_P W_P + P' Z'_N Q_*^{-1} W_N F_2 \\
 W'_P W_P + F'_1 W'_N Q_*^{-1} W_N F_1 + G^{-1}_{v1} P \alpha_{v1} \\
 W'_P W_P + F'_2 W'_N Q_*^{-1} W_N F_2 + G^{-1}_{v2} P \alpha_{v2}
 \end{bmatrix}
 \quad (9)
 \end{aligned}$$



**EXAMPLE**

We use the same data that Fernando and Grossman (1989) employed. There are 4 individuals, 3 of them are parents and 1 is a non-parent. The file is :

Individual	Sire	Dam	Marker inheritance	
			Sire	Dam
1	-	-	-	-
2	-	-	-	-
3	1	2	$M_1^p$	$M_2^m$
4	1	3	-	$M_3^p$

Notice that individual 4 is inbred. A fixed effect was included and the matrix resulting from adjoining the incidence matrix  $\mathbf{X}$  and the vector of observations  $\mathbf{y}$ , *ie*  $[\mathbf{X}|\mathbf{y}]$  is :

$$[\mathbf{X}|\mathbf{y}] = \begin{matrix} 1 \\ 2 \\ 3 \\ 4 \end{matrix} \left[ \begin{array}{c|c} 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \end{array} \right] \begin{matrix} 235 \\ 210 \\ 250 \\ 255 \end{matrix}$$

Variance components used were  $\sigma_A^2 = 100$ ,  $\sigma_v^2 = 10$  and  $\sigma_e^2 = 500$  and  $r = 0.1$ . The matrices  $\mathbf{G}_v$  and  $\mathbf{G}_v^{-1}$  are presented in Fernando and Grossman (1989).

First, solutions for AM were obtained. The coefficient matrix for AM is :

2.0	0.0	1.0	0.0	0.0	1.0	1.0	1.0	1.0	1.0	0.0	0.0	0.0	0.0
0.0	2.0	0.0	0.0	1.0	1.0	0.0	0.0	0.0	0.0	1.0	1.0	1.0	1.0
1.0	0.0	11.0	2.5	-2.5	-5.0	1.0	1.0	0.0	0.0	0.0	0.0	0.0	0.0
1.0	0.0	2.5	8.5	-5.0	0.0	0.0	0.0	1.0	1.0	0.0	0.0	0.0	0.0
0.0	1.0	-2.5	-5.0	13.5	-5.0	0.0	0.0	0.0	0.0	1.0	1.0	0.0	0.0
0.0	1.0	-5.0	0.0	-5.0	11.0	0.0	0.0	0.0	0.0	0.0	0.0	1.0	1.0
1.0	0.0	1.0	0.0	0.0	0.0	301.0	51.0	0.0	0.0	-250.0	0.0	-50.0	0.0
1.0	0.0	1.0	0.0	0.0	0.0	51.0	78.7	0.0	0.0	-27.7	0.0	-50.0	0.0
1.0	0.0	0.0	1.0	0.0	0.0	0.0	0.0	53.7	26.0	0.0	-27.7	0.0	0.0
1.0	0.0	0.0	1.0	0.0	0.0	0.0	0.0	26.0	276.0	0.0	-250.0	0.0	0.0
0.0	1.0	0.0	0.0	1.0	0.0	-25.0	-27.7	0.0	0.0	503.7	26.0	0.0	-250.0
0.0	1.0	0.0	0.0	1.0	0.0	0.0	0.0	-27.7	-250.0	26.0	281.5	0.0	-27.7
0.0	1.0	0.0	0.0	0.0	1.0	-50.0	-50.0	0.0	0.0	0.0	0.0	101.0	1.0
0.0	1.0	0.0	0.0	0.0	1.0	0.0	0.0	0.0	0.0	-250.0	-27.7	1.0	278.7

and the right-hand site vector is [445, 505, 235, 210, 250, 255, 235, 235, 210, 210, 250, 250, 255, 255]'. The vector of solutions is [222.5, 251.764, 2.08109, -2.08109, -0.083214, 1.16537, 0.213435, 0.216098, -0.202783, -0.226749, 0.213102, -0.229745, 0.231409, 0.174809].

There are 11 equations in the system for RAM (as compared to 14 in AM) since there is only 1 non-parent (individual 4) and  $\mathbf{Q}$  is a scalar :  $1.1136 = 1 + (0.5/\alpha_A) + 2[0.5(0.5) + (0.9)(0.1)]/\alpha_v$ . The vector of right-hand sides for equations

(7) is [445, 478.987, 349.494, 210, 364.494, 349.494, 349.494, 210, 210, 456.088, 272.899]' and the coefficient matrix is :

2.000	0.000	1.000	1.000	0.000	1.000	1.000	1.000	1.000	0.000	0.000
0.000	1.898	0.449	0.000	1.449	0.449	0.449	0.000	0.000	1.808	1.089
1.000	0.449	8.724	2.500	-4.775	1.224	1.224	0.000	0.000	0.404	0.044
1.000	0.000	2.500	8.500	-5.000	0.000	0.000	1.000	1.000	0.000	0.000
0.000	1.449	-4.775	-5.000	11.224	0.224	0.224	0.000	0.000	1.404	1.044
1.000	0.499	1.224	0.000	0.224	276.224	26.224	0.000	0.000	-249.596	0.044
1.000	0.449	1.224	0.000	0.224	26.224	54.002	0.000	0.000	-27.373	0.044
1.000	0.000	0.000	1.000	0.000	0.000	0.000	53.777	26.000	0.000	-27.777
1.000	0.000	0.000	1.000	0.000	0.000	0.000	26.000	276.000	0.000	-250.000
0.000	1.808	0.404	0.000	1.404	-249.596	-27.373	0.000	0.000	279.505	1.080
0.000	1.089	0.044	0.000	1.044	0.044	0.044	-27.777	-250.0	1.080	278.787

Solutions for RAM are 222.5, 251.764, 2.08109, -2.08109, -0.083214, 0.213435, 0.216098, -0.202783, -0.226749, 0.213102 and -0.229745. The next step is to backsolve for individual 4 (non-parent) using equation (B.2). Since both parents of 4 are known,  $d_{A44} = 0.5$  and  $d_{o44} = 5/(5 + 0.5) = 10/11 = 0.90909...$  The diagonal elements of the  $2 \times 2$  system in (B.2) are functions of  $r$ . However, as the information from the sire marker is unavailable,  $r = 0.5$  for the first diagonal element. Also,  $F_1 = F_3 = 0$  and  $d_{o44} y_4^* = 0.90909[255 - 251.764 - 0.5(2.081090 + 0.083214 + 0.213435 + 0.216098) - 0.9(0.213102) - 0.1(-0.229745)] = 1.6848545$ . For animal 4, we then have :

$$\begin{bmatrix} 0.90909 + 100 & 0.90909 \\ 0.90909 & 0.90909 + 277.7777 \end{bmatrix} \begin{bmatrix} \hat{\epsilon}_1^p \\ \hat{\epsilon}_3^m \end{bmatrix} = \begin{bmatrix} 1.6848545 \\ 1.6848545 \end{bmatrix}$$

which has solutions  $\hat{\epsilon}_1^p = 0.0166428$  and  $\hat{\epsilon}_3^m = 0.00599141$ . Putting these into (B.3) gives  $\hat{\phi}_4 = 0.166428$ . Therefore,  $BLUP(u_4) = 0.5 BLUP(u_1) + 0.5 BLUP(u_3) + BLUP(\phi_4) = 0.5[2.08109 + (-0.083214)] + 0.166428 = 1.16537$ . Also,  $BLUP(v_4^p) = 0.5 BLUP(v_1^p) + 0.5 BLUP(v_1^m) + BLUP(\epsilon_4^p) = 0.5[0.213435 + 0.216098] + 0.0166428 = 0.231409$  and  $BLUP(v_4^m) = 0.9 BLUP(v_3^p) + 0.1 BLUP(v_3^m) + BLUP(\epsilon_4^m) = 0.9(0.213102) + 0.1(-0.229745) + 0.00599141 = 0.174809$ . As expected, solutions obtained by both AM and RAM are the same.

**DISCUSSION**

The advantage of RAM over AM increases as both the ratio between the number of non-parents and the number of parents and the number of independent MQTL increase. Goddard (1991) suggested the use of RAM to decrease the size of the resulting system of equations when working with information on flanking markers.

As shown in Appendix A and for a non-inbred individual, the fraction of the variance of the MQTL that is due to Mendelian segregation is  $4r(1 - r)/2$ . Now,  $1 = (r + 1 - r)^2 = r^2 + 2r(1 - r) + (1 - r)^2$ , so that  $2[1 - 2r(1 - r)] = 2[r^2 + (1 - r)^2]$ . Therefore, the fraction of the variance of the MQTL that is explained by parental segregation is  $2[r^2 + (1 - r)^2]/2$ . These proportions can also be worked out from equations (8a) and (8b) in Fernando and Grossman (1989) and they agree with

formulae derived by Dekkers and Dentine (1991). A slight difference between their result and the one obtained here stems from the fact that they define the variance of the MQTL as one half the variance as defined by Fernando and Grossman (1989) ( $0.5\sigma_v^2$ ).

Both AM and RAM rest on knowing the variance components as well as the recombination rate between the marker gene and the QTL. As the latter parameter enters into the variance-covariance matrix of QTL effects in a rather complex manner, its estimation by the classical methods employed in animal breeding seems to be difficult, as discussed by Fernando (1990).

When more than one MQTL is being considered, covariances between pairs of MQTL effects are likely to be non-zero due to linkage disequilibrium caused by selection (Bulmer, 1985). Model (8) assumes that these covariances are zero. The extent of the error in predicting  $\mathbf{v}$  (or functions of  $\mathbf{v}$ ) due to incorrectly assuming null covariances between MQTL effects will depend on the magnitude and sign of the covariance. If the covariances are mostly negative, which is likely to happen on a trait undergoing selection (Bulmer, 1985), MQTL effects may be overpredicted. Research is in progress to overcome this restriction of model (8).

## APPENDIX A

### Derivation of the diagonal elements of $\mathbf{W}_N \mathbf{G}_e \mathbf{W}'_N$ when all non-parents have one record

First we show that  $\mathbf{W}_N \mathbf{G}_e \mathbf{W}'_N$  is diagonal. Because  $\mathbf{G}_e$  is diagonal (Fernando and Grossman, 1989), we can write :

$$\mathbf{W}_N \mathbf{G}_e \mathbf{W}'_N = \sum_{j=1}^{2(n-s)} \mathbf{w}_j \mathbf{w}'_j \mathbf{g}_j$$

where  $\mathbf{w}_j$  is the column  $j$  of  $\mathbf{W}_N$  and  $g_j$  is diagonal element  $j$  of  $\mathbf{G}_e$ . Now,  $\mathbf{w}_j$  has all its elements equal to zero except for a 1 in position  $j$ . Therefore, the matrix  $\mathbf{w}_j \mathbf{w}'_j \mathbf{g}_j$  has all elements equal to zero except for element  $j$ ,  $j$  which is equal to  $\mathbf{g}_j$ . The paternal and maternal MQTL additive effects of an animal are in consecutive columns of the matrix  $\mathbf{W}$  (and  $\mathbf{W}_N$ ),  $\mathbf{w}_j$  and  $\mathbf{w}_{j+1}$  say, and these are equal. We then have :

$$\sum_{j=1}^{2(n-s)} \mathbf{w}_j \mathbf{w}'_j \mathbf{g}_j = \sum_{j=1,3,5\dots}^{2(n-s)} \mathbf{w}_j \mathbf{w}'_j (\mathbf{g}_j + \mathbf{g}_{j+1})$$

and  $\mathbf{W}_N \mathbf{G}_e \mathbf{W}'_N$  is diagonal with non-zero elements equal to  $\mathbf{g}_j + \mathbf{g}_{j+1}$ .

Now,  $(g_j + g_{j+1})\sigma_v^2 = \text{Var}(\boldsymbol{\epsilon}_o^p) + \text{Var}(\boldsymbol{\epsilon}_o^m) = 2r(1-r)(1-f_s) + 2r(1-r)(1-f_d)$ , where  $f_s$  and  $f_d$  are the inbreeding coefficients of sire and dam for the MQTL, respectively. The last equality follows from expressions (12a) and (12b) in Fernando and Grossman (1989). After some rearranging, the diagonal element of  $\mathbf{W}_N \mathbf{G}_e \mathbf{W}'_N$  is :

$$2r(1-r)(2 - f_s - f_d)$$

when both parents of the individual are known. If the sire is unknown,  $\boldsymbol{\epsilon}_o^p = \mathbf{v}_o^p$  and the diagonal element is  $2r(1-r)(1-f_d) + 1$ . If the dam is unknown,  $\boldsymbol{\epsilon}_o^m = \mathbf{v}_o^m$  and the diagonal element is  $2r(1-r)(1-f_s) + 1$ . If both parents are unknown the diagonal element of  $\mathbf{W}_N \mathbf{G}_\boldsymbol{\epsilon} \mathbf{W}'_N$  is 2.

### APPENDIX B

#### Solutions of equations (10) when non-parents have one record

When non-parents have one record ( $\mathbf{Z}_N = \mathbf{I}$ ), equations (10) reduce to :

$$\begin{bmatrix} \mathbf{I} + \mathbf{D}_A^{-1} \boldsymbol{\alpha}_A & \mathbf{W}_N \\ \mathbf{W}'_N & \mathbf{W}'_N \mathbf{W}_N + \mathbf{G}_\boldsymbol{\epsilon}^{-1} \boldsymbol{\alpha}_v \end{bmatrix} \begin{bmatrix} \hat{\boldsymbol{\phi}} \\ \hat{\boldsymbol{\epsilon}} \end{bmatrix} = \begin{bmatrix} \mathbf{I} \\ \mathbf{W}'_N \end{bmatrix} (\mathbf{y}_N - \mathbf{X}_N \hat{\boldsymbol{\beta}} - \mathbf{P} \hat{\mathbf{u}}_P - \mathbf{W}_N \mathbf{F} \hat{\mathbf{v}}_P) \quad (\text{B.1})$$

On absorbing the equations for  $\hat{\boldsymbol{\phi}}$ , the solution for  $\hat{\boldsymbol{\epsilon}}$  is :

$$\begin{aligned} & (\mathbf{W}'_N [\mathbf{I} - (\mathbf{I} + \mathbf{D}_A^{-1} \boldsymbol{\alpha}_A)^{-1}] \mathbf{W}_N + \mathbf{G}_\boldsymbol{\epsilon}^{-1} \boldsymbol{\alpha}_v) \hat{\boldsymbol{\epsilon}} \\ & = \mathbf{W}'_N [\mathbf{I} - (\mathbf{I} + \mathbf{D}_A^{-1} \boldsymbol{\alpha}_A)^{-1}] (\mathbf{y}_N - \mathbf{X}_N \hat{\boldsymbol{\beta}} - \mathbf{P} \hat{\mathbf{u}}_P - \mathbf{W}_N \mathbf{F} \hat{\mathbf{v}}_P) \end{aligned}$$

The matrix  $\mathbf{D}_o = \mathbf{I} - (\mathbf{I} + \mathbf{D}_A^{-1} \boldsymbol{\alpha}_A)^{-1}$  is diagonal with element  $d_{oii}$  being equal to :

$$1 - \left( 1 + \frac{\boldsymbol{\alpha}_A}{d_{Aii}} \right)^{-1} = \frac{\boldsymbol{\alpha}_A}{\boldsymbol{\alpha}_A + d_{Aii}}$$

and  $d_{Aii}$  is diagonal element  $i$  of  $\mathbf{D}_A$ . Since  $\mathbf{W}$  (and  $\mathbf{W}_N$ ) has rows with 2 consecutive elements equal to 1 and the rest equal to 0,  $\mathbf{W}'_N \mathbf{D}_o \mathbf{W}_N$  is block diagonal, each block being of order  $2 \times 2$  with all elements equal to  $d_{oii}$ . Adding  $\mathbf{G}_\boldsymbol{\epsilon}^{-1} \boldsymbol{\alpha}_v$  gives the coefficient matrix on the left-hand side of (B.1) and solutions for  $\hat{\boldsymbol{\epsilon}}$  can be obtained by solving  $(n - s)$  systems of order 2. The system for animal  $i$  is equal to :

$$\begin{bmatrix} d_{oii} + \frac{\boldsymbol{\alpha}_v}{2r(1-r)(1-f_s)} & d_{oii} \\ d_{oii} & d_{oii} + \frac{\boldsymbol{\alpha}_v}{2r(1-r)(1-f_d)} \end{bmatrix} \begin{bmatrix} \hat{\boldsymbol{\epsilon}}_i^p \\ \hat{\boldsymbol{\epsilon}}_i^m \end{bmatrix} = \begin{bmatrix} 1 \\ 1 \end{bmatrix} d_{oii} \mathbf{y}_i^* \quad (\text{B.2})$$

and  $\mathbf{y}_i^*$  is element  $i$  of the vector  $\mathbf{y}_N - \mathbf{X}'\boldsymbol{\beta} - \mathbf{P}\hat{\mathbf{u}}_P - \mathbf{W}_N \mathbf{F} \hat{\mathbf{v}}_P$ .

After solving for  $\hat{\boldsymbol{\epsilon}}$ , the first equation in (B.1) can be solved as follows :

$$(\mathbf{I} + \mathbf{D}_A^{-1} \boldsymbol{\alpha}_A) \hat{\boldsymbol{\phi}} = \mathbf{y}^* - \mathbf{W}_N \hat{\boldsymbol{\epsilon}}$$

Since the coefficient matrix of this system is diagonal, BLUP ( $\phi_i$ ) is :

$$\hat{\phi} = \left( \frac{d_{Aii}}{d_{Aii} + \alpha_A} \right) (y_i^* - \hat{\epsilon}_i^p - \hat{\epsilon}_i^m) \tag{B.3}$$

If there is more than one MQTL the matrix of system (10) becomes poorly conditioned. The reason is that all off-diagonal elements are equal to 1 and the diagonals are relatively large (may be in the order of hundreds, depending on the  $\alpha$ 's). An exact solution can be obtained by writing the matrix of system (10) for each animal as  $\mathbf{jj}' + \mathbf{S}$ , where  $\mathbf{j}$  is a  $1 + 2m$  vector with all elements equal to one and  $\mathbf{S}$  is a diagonal matrix. Using the inverse of the sum of matrices formula (Henderson and Searle, 1981), we have that :

$$(\mathbf{jj}' + \mathbf{S})^{-1} = \mathbf{S}^{-1} - \mathbf{S}^{-1}\mathbf{j}(\mathbf{j}'\mathbf{S}^{-1}\mathbf{j} + 1)^{-1}\mathbf{j}'\mathbf{S}^{-1} \tag{B.4}$$

Notice that the expression in parenthesis on the right-hand side of (B.4) is a scalar. Using (B.4) the inverse of the matrix of system (10) when there are  $m$  MQTL's for a non-parent is :

$$(\mathbf{jj}' + \mathbf{S})^{-1} = \begin{bmatrix} \mathbf{S}_1(1 - \frac{\mathbf{S}_1}{\mathbf{g}}) & -\frac{\mathbf{S}_1\mathbf{S}_2}{\mathbf{g}} & -\frac{\mathbf{S}_1\mathbf{S}_3}{\mathbf{g}} & \dots \\ -\frac{\mathbf{S}_2\mathbf{S}_1}{\mathbf{g}} & \mathbf{S}_2(1 - \frac{\mathbf{S}_2}{\mathbf{g}}) & -\frac{\mathbf{S}_2\mathbf{S}_3}{\mathbf{g}} & \dots \\ -\frac{\mathbf{S}_3\mathbf{S}_1}{\mathbf{g}} & -\frac{\mathbf{S}_3\mathbf{S}_2}{\mathbf{g}} & \mathbf{S}_3(1 - \frac{\mathbf{S}_3}{\mathbf{g}}) & \dots \\ \dots & \dots & \dots & \dots \end{bmatrix} \tag{B.5}$$

and  $\mathbf{g} = 1 + \mathbf{S}_1 + \mathbf{S}_2 + \dots + \mathbf{S}_{2m}$ . The  $\mathbf{S}$ 's are such that  $\mathbf{S}_1 = \mathbf{D}_{Aii}\alpha_A^{-1}$ ,  $\mathbf{S}_2 = \mathbf{G}_{\epsilon_s}\alpha_{v1}^{-1} = 2r_1(1 - r_1)(1 - f_{1s})\alpha_{v1}^{-1}$ ,  $\mathbf{S}_3 = \mathbf{G}_{ed}\alpha_{v1}^{-1} = 2r_1(1 - r_1)(1 - f_{1d})\alpha_{v1}^{-1}$ , and so on for MQTL's 2 to  $m$ . Expression (B.5) is easy to program, does not require iteration and, more importantly, it is not subject to the numerical problems that occur when solving such a system of equations.

**ACKNOWLEDGMENTS**

This work was supported by the Natural Sciences and Engineering Research Council of Canada. LR Schaeffer provided many helpful comments. G Banos and N Caron translated the summary into French. Thanks are also extended to the editor for his comments and to a referee for pointing out a problem with the definition of the inbreeding coefficient at the MQTL.

**REFERENCES**

- Bulmer MG (1985) *The Mathematical Theory of Quantitative Genetics*. Clarendon Press, Oxford
- Dekkers JCM, Dentine MR (1991) Quantitative genetic variance associated with chromosomal markers in segregating populations. *Theor Appl Gen* 81, 212-220
- Fernando RL (1990) Statistical problems in marker assisted selection for QTL. In : *4th World Congr, Genetics Applied to Livestock Production, Edinburgh, 23-27 July 1990*, vol 13 (Hill WG, Thompson R, Wooliams JA, eds) 433-436
- Fernando RL, Grossman M (1989) Marker assisted selection using best linear unbiased prediction. *Gen Sel Evol* 21, 467-477
- Goddard ME (1991) A mixed model for analyses of data on multiple genetic markers. *Theor Appl Gen* (submitted)
- Henderson CR (1984) *Applications of Linear Models in Animal Breeding*. University of Guelph, Guelph, Ontario
- Henderson CR (1985) Equivalent linear models to reduce computations. *J Dairy Sci* 68, 2267-2277
- Henderson HV, Searle SR (1981) On deriving the inverse of a sum of matrices. *SIAM Rev* 23, 53-59
- Quaas RL, Pollak EJ (1980) Mixed model methodology for farm and ranch beef cattle testing programs. *J Anim Sci* 51, 1277-1287
- Schaeffer LR, Wilton JW (1987) RAM computing strategies and multiple traits. In : *Prediction of Genetic Values for Beef Cattle : Proc Workshop II, Kansas City, 10-11 March 1987*. Winrock Int Inst Agric Dev, Morrilton, Arkansas, USA, 25-52