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Behavior of the plastic deformation of an elasto-perfectly-plastic oscillator with noise.¹

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Abstract

Earlier works in engineering, partly experimental, partly computational have revealed that asymptotically, when the excitation is a white noise, plastic deformation and total deformation for an elasto-perfectly-plastic oscillator have a variance which increases linearly with time with the same coefficient. In this work, we prove this result and we characterize the corresponding drift coefficient. Our study relies on a stochastic variational inequality governing the evolution between the velocity of the oscillator and the non-linear restoring force. We then define long cycles behavior of the Markov process solution of the stochastic variational inequality which is the key concept to obtain the result. An important question in engineering is to compute this coefficient. Also, we provide numerical simulations which show successful agreement with our theoretical prediction and previous empirical studies made by engineers.

Résumé

Le comportement de la déformation plastique pour un oscillateur élastique-parfaitement-plastique excité par un bruit blanc. Des résultats expérimentaux en sciences de l'ingénieur ont montré que, pour un oscillateur élasto-plastique-parfait excité par un bruit blanc, la déformation plastique et la déformation totale ont une variance, qui asymptotiquement, croît linéairement avec le temps avec le même coefficient. Dans ce travail, nous prouvons ce résultat et nous caractérisons le coefficient de dérive. Notre étude repose sur une inéquation variationnelle stochastique gouvernant l'évolution entre la vitesse de l'oscillateur et la force de rappel non-linéaire. Nous définissons alors le comportement en cycles longs du processus de Markov solution de l'inéquation variationnelle stochastique qui est le concept clé pour obtenir le résultat. Une question importante en sciences de l'ingénieur est de calculer ce coefficient. Les résultats numériques confirment avec succès notre prédiction théorique et les études empiriques faites par les ingénieurs.

Version française abrégée

Dans cet article, nous étudions la variance de l'oscillateur élastique-parfaitement-plastique (EPP) excité par un bruit blanc. La dynamique de l'oscillateur s'exprime à l'aide d'une équation à mémoire. (voir

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(1)-(2)). A.Bensoussan and J.Turi ont montré que la relation entre la vitesse et la composante élastique satisfait une inéquation variationnelle stochastique (voir (SVI)). Dans ce cadre, nous introduisons les cycles long indépendants (définis plus loin) et nous justifions qu'ils permettent de caractériser la variance de la déformation totale et de la déformation plastique (voir (4)).

1. Introduction

In civil engineering, an elasto-perfectly-plastic (EPP) oscillator with one single degree of freedom is employed to estimate prediction of failure of mechanical structures under random vibrations. This elasto-plastic oscillator consists in a one dimensional model simple and representative of the elasto-plastic behavior of a class of structure dominated by their first mode of vibration. The main difficulty to study these systems comes from a frequent occurrence of plastic phases on small intervals of time. A plastic deformation is produced when the stress of the structure crosses over an elastic limit. The dynamics of the EPP-oscillator has memory, so it has been formulated in the engineering literature as a process with hysteresis $x(t)$, which stands for the displacement of the oscillator. We study the problem

$$\ddot{x} + c_0 \dot{x} + \mathbf{F}_t = \dot{w} \quad (1)$$

with initial conditions of displacement and velocity $x(0) = 0$, $\dot{x}(0) = 0$. Here $c_0 > 0$ is the viscous damping coefficient, $k > 0$ the stiffness, w is a Wiener process; $\mathbf{F}_t := \mathbf{F}(x(s), 0 \leq s \leq t)$ is a non-linear functional which depends on the entire trajectory $\{x(s), 0 \leq s \leq t\}$ up to time t . The plastic deformation denoted by $\Delta(t)$ at time t can be recovered from the pair $(x(t), \mathbf{F}_t)$ by the following relationship:

$$\mathbf{F}_t = \begin{cases} kY & \text{if } x(t) = Y + \Delta(t), \\ k(x(t) - \Delta(t)) & \text{if } x(t) \in]-Y + \Delta(t), Y + \Delta(t)[, \\ -kY & \text{if } x(t) = -Y + \Delta(t). \end{cases} \quad (2)$$

where $\Delta(t) = \int_0^t y(s) \mathbf{1}_{\{|\mathbf{F}_s| = kY\}} ds$ and Y is an elasto-plastic bound. Karnopp & Scharton [5] proposed a separation between elastic states and plastic states. They introduced a fictitious variable $z(t) := x(t) - \Delta(t)$ and noticed the simple fact that between two plastic phases $z(t)$ behaves like a linear oscillator. In addition, other works made by engineers [6], partly experimental, partly computational, have revealed that total deformation has a variance which increases linearly with time:

$$\lim_{t \rightarrow \infty} \frac{\sigma^2(x(t))}{t} = \sigma^2. \quad (3)$$

where σ^2 is a positive real number.

Recently, the right mathematical framework of stochastic variational inequalities (SVI) modeling an elasto-plastic oscillator with noise (presented below) has been introduced by the first author with J. Turi in [3]. The inequality governs the relationship between the velocity $y(t)$ and the elastic component $z(t)$:

$$dy(t) = -(c_0 y(t) + k z(t)) dt + dw(t), \quad (dz(t) - y(t) dt)(\phi - z(t)) \geq 0, \quad \forall |\phi| \leq Y, \quad |z(t)| \leq Y. \quad (\text{SVI})$$

The plastic deformation $\Delta(t)$ can be recovered by $\int_0^t y(s) \mathbf{1}_{\{|z(s)| = Y\}} ds$.

Throughout the paper, the objective is to prove (3) and to provide a exact characterization of (3) based on (SVI). Now, we introduce *long cycles* in view to the fact that we are interested in identifying independent sequences in the trajectory.

1.1. Long cycles

Denote $\tau_0 := \inf\{t > 0, \ y(t) = 0 \text{ and } |z(t)| = Y\}$ and $s := \text{sign}(z(\tau_0))$ which labels the first boundary hit by the process $(y(t), z(t))$. Define $\theta_0 := \inf\{t > \tau_0, \ y(t) = 0 \text{ and } z(t) = -sY\}$. In a recurrent manner, knowing θ_n , we can define for $n \geq 0$

$$\begin{aligned}\tau_{n+1} &:= \inf\{t > \theta_n, \ y(t) = 0 \text{ and } z(t) = sY\}, \\ \theta_{n+1} &:= \inf\{t > \tau_{n+1}, \ y(t) = 0 \text{ and } z(t) = -sY\}.\end{aligned}$$

Now, according to the previous setting we can define the n -th long cycle (resp. first part of the cycle, second part of the cycle) as the piece of trajectory delimited by the interval $[\tau_n, \tau_{n+1})$, (resp. $[\tau_n, \theta_{n+1})$ and $[\theta_{n+1}, \tau_{n+1})$). Indeed, at the instants $\{\tau_n, n \geq 1\}$, the process $(y(t), z(t))$ is in the same state at the instant τ_0 . In addition, there are two types of cycles depending on the sign of s . The set of stopping times $\{\tau_n, n \geq 0\}$ represents the times of occurrence of long cycles. Let us remark that the plastic deformation and the total deformation are the same on a long cycle since $\int_{\tau_0}^{\tau_1} y(t)dt = \int_{\tau_0}^{\tau_1} y(t)\mathbf{1}_{\{|z(s)|=Y\}}dt + \int_{\tau_0}^{\tau_1} y(t)\mathbf{1}_{\{|z(s)|<Y\}}dt$ and that $\int_{\tau_0}^{\tau_1} y(t)\mathbf{1}_{\{|z(s)|<Y\}}dt = z(\tau_1) - z(\tau_0) = 0$. As main result, we have obtained the following theorem.

Theorem 1.1 (Characterization of the variance related to the plastic/total deformation). *In the previously defined context, we have shown*

$$\lim_{t \rightarrow \infty} \frac{\sigma^2(x(t))}{t} = \frac{\mathbb{E} \left(\int_{\tau_0}^{\tau_1} y(t)dt \right)^2}{\mathbb{E}(\tau_1 - \tau_0)}. \quad (4)$$

Our proofs are based on solving nonlocal partial differential equations related to long cycles. Simpler formula will be given below at equation (13).

2. The issue of long cycles and plastic deformations

Let us introduce notations.

Notation 1. $D := \mathbb{R} \times (-Y, +Y)$, $D^+ := (0, \infty) \times \{Y\}$, $D^- := (-\infty, 0) \times \{-Y\}$, and the differential operators $A\zeta := -\frac{1}{2}\zeta_{yy} + (c_0y + kz)\zeta_y - y\zeta_z$, $B_+\zeta := -\frac{1}{2}\zeta_{yy} + (c_0y + kY)\zeta_y$, $B_-\zeta := -\frac{1}{2}\zeta_{yy} + (c_0y - kY)\zeta_y$, where ζ is a regular function on D .

Let f be a bounded measurable function, we want to solve

$$Av^+ = f(y, z) \text{ in } D, \quad B_+v^+ = f(y, Y) \text{ in } D^+, \quad B_-v^+ = f(y, -Y) \text{ in } D^- \quad (P_{v^+})$$

with the nonlocal boundary conditions $v^+(y, Y)$ continuous and $v^+(0^-, -Y) = 0$, and

$$Av^- = f(y, z) \text{ in } D, \quad B_+v^- = f(y, Y) \text{ in } D^+, \quad B_-v^- = f(y, -Y) \text{ in } D^- \quad (P_{v^-})$$

with the nonlocal boundary conditions $v^-(0^+, Y) = 0$ and $v^-(y, -Y)$ continuous. The functions $v^+(y, z; f)$ and $v^-(y, z; f)$ are called long cycles. In addition, we define $\pi^+(y, z)$ and $\pi^-(y, z)$ by

$$A\pi^+ = 0 \text{ in } D, \quad \pi^+(y, Y) = 1 \text{ in } D^+, \quad \pi^+(y, -Y) = 0 \text{ in } D^- \quad (P_{\pi^+})$$

and

$$A\pi^- = 0 \text{ in } D, \quad \pi^-(y, Y) = 0 \text{ in } D^+, \quad \pi^-(y, -Y) = 1 \text{ in } D^- \quad (P_{\pi^-})$$

Note that $\pi^+(y, z) + \pi^-(y, z) = 1$, so the existence and uniqueness of a bounded solution (P_{π^+}) and (P_{π^-}) is clear.

Proposition 2.1. *We have the properties $\pi^-(0^+, -Y) > 0$ and $\pi^+(0^-, Y) > 0$.*

PROOF. We check only the first property. We consider the elastic process $(y_{yz}(t), z_{yz}(t))$:

$$z_{yz}(t) = e^{-\frac{c_0 t}{2}} \left\{ z \cos(\omega t) + \frac{1}{\omega} \left(y + \frac{c_0}{2} z \right) \sin(\omega t) \right\} + \frac{1}{\omega} \int_0^t e^{-\frac{c_0}{2}(t-s)} \sin(\omega(t-s)) dw(s),$$

$$y_{yz}(t) = -\frac{c_0}{2} z_{yz}(t) + e^{-\frac{c_0 t}{2}} \left\{ -\omega z \sin(\omega t) + \left(y + \frac{c_0}{2} z \right) \cos(\omega t) \right\} + \int_0^t e^{-\frac{c_0}{2}(t-s)} \cos(\omega(t-s)) dw(s)$$

where $\omega := \frac{\sqrt{4k-c_0^2}}{2}$ (we assume $4k > c_0^2$). Note that the condition $4k > c_0^2$ is needed so that $(y_{yz}(t), z_{yz}(t))$ have real valued solutions. Set $\tau_{yz} := \inf\{t > 0, |z_{yz}(t)| \geq Y\}$ then we have the probabilistic interpretation $\pi^+(y, z) = \mathbb{P}(z_{yz}(\tau_{yz}) = Y)$, $\pi^-(y, z) = \mathbb{P}(z_{yz}(\tau_{yz}) = -Y)$. We can state

$$\pi^-(y, z) \rightarrow 1 \quad \text{as } y \rightarrow -\infty, \quad z \in [-Y, Y]. \quad (5)$$

Indeed, $\forall t$ with $0 < t < \frac{\pi}{\omega}$ we have $z_{yz}(t) \rightarrow -\infty$, as $y \rightarrow -\infty$ a.s. Therefore $\forall t$ with $0 < t < \frac{\pi}{\omega}$, a.s. $z_{yz}(t) < -Y$ for y sufficiently large. Hence, a.s. $\tau_{yz} < t$ for y sufficiently large. Therefore a.s. $\limsup_{y \rightarrow -\infty} \tau_{yz} < t$. Since t is arbitrary, necessarily a.s. $\tau_{yz} \rightarrow 0$, as $y \rightarrow -\infty$ which implies (5). Moreover the function $\pi^+(y, Y)$ cannot have a minimum or a maximum at any finite $y < 0$. It is then monotone and since $\pi^-(\infty, Y) = 1$, it is monotone decreasing. It follows that $\pi^-(0^-, Y) < 1$. It cannot be 0. Otherwise, $\pi^-(y, Y)$ is continuous at $y = 0$, and $(0, Y)$ is a point of minimum of $\pi^-(y, z)$. Since for $y < 0$, $\pi_y^-(y, Y) < 0$ from the equation of π^- we get $\limsup_{y \rightarrow 0} \pi_{yy}^-(y, Y) \leq 0$ which is not possible since $(0, Y)$ is a minimum.

We next define $\eta(y, z)$ by

$$-\frac{1}{2}\eta_{yy} + (c_0 y + k z)\eta_y - y\eta_z = f(y, z) \quad \text{in } D, \quad \eta(y, Y) = 0 \quad \text{in } D^+, \quad \eta(y, -Y) = 0 \quad \text{in } D^- \quad (P_\eta)$$

with the local boundary conditions $\eta(0^+, Y) = 0$ and $\eta(0^-, -Y) = 0$. For f bounded (P_η) has a unique bounded solution. Then define $\varphi_+(y; f)$ by solving

$$-\frac{1}{2}\varphi_{+,yy} + (c_0 y + k Y)\varphi_{+,y} = f(y, Y), \quad y > 0, \quad \varphi_+(0; f) = 0. \quad (6)$$

We check easily the formula

$$\varphi_+(y; f) = 2 \int_0^\infty d\xi \exp(-(c_0 \xi^2 + 2kY\xi)) \int_\xi^{\xi+y} f(\zeta; Y) \exp(-2c_0 \xi(\zeta - \xi)) d\zeta, \quad y \geq 0. \quad (7)$$

We consider $\psi_+(y, z; f)$ defined by

$$\begin{cases} -\frac{1}{2}\psi_{+,yy} + (c_0 y + k z)\psi_{+,y} - y\psi_{+,z} = 0 & \text{in } D, \\ \psi_+(y, Y) = \varphi_+(y; f) & \text{in } D^+, \\ \psi_+(y, -Y) = 0 & \text{in } D^- \end{cases} \quad (P_{\psi_+})$$

and similarly $\varphi_-(y; f), \psi_-(y, z; f)$ defined by

$$-\frac{1}{2}\varphi_{-,yy} + (c_0 y - k Y)\varphi_{-,y} = f(y, -Y), \quad y < 0, \quad \varphi_-(0; f) = 0. \quad (8)$$

which leads to

$$\varphi_-(y; f) = 2 \int_0^\infty d\xi \exp(-(c_0 \xi^2 - 2kY\xi)) \int_{y-\xi}^{-\xi} f(\zeta; -Y) \exp(-2c_0 \xi(\zeta - \xi)) d\zeta, \quad y \leq 0.$$

and

$$\begin{cases} -\frac{1}{2}\psi_{-,yy} + (c_0 y + k z)\psi_{-,y} - y\psi_{-,z} = 0 & \text{in } D, \\ \psi_-(y, Y) = 0 & \text{in } D^+, \\ \psi_-(y, -Y) = \varphi_-(y; f) & \text{in } D^- \end{cases} \quad (P_{\psi_-})$$

We can state

Proposition 2.2. *The solution of (P_{v+}) is given by*

$$\begin{aligned} v^+(y, z; f) = & \eta(y, z; f) + \psi_+(y, z; f) + \psi_-(y, z; f) \\ & + \frac{\eta(0^-, Y; f) + \psi_+(0^-, Y; f) + \psi_-(0^-, Y; f)}{\pi^-(0^-, Y)} \pi^+(y, z). \end{aligned} \quad (9)$$

and the solution of (P_{v-}) is given by

$$\begin{aligned} v^-(y, z; f) = & \eta(y, z; f) + \psi_+(y, z; f) + \psi_-(y, z; f) \\ & + \frac{\eta(0^+, -Y; f) + \psi_+(0^+, -Y; f) + \psi_-(0^+, -Y; f)}{\pi^+(0^+, Y)} \pi^-(y, z). \end{aligned} \quad (10)$$

PROOF. Direct checking.

3. Complete cycle

First, let us check that $\mathbb{E}[x(t)] = 0$ and then $\sigma^2(x(t)) = \mathbb{E}[x^2(t)]$. Indeed, by symmetry of the inequality (SVI) and by the choice of initial conditions $y(0) = 0, z(0) = 0$, the processes $(y(t), z(t))$ and $-(y(t), z(t))$ have same law. Then, $\mathbb{E}[x(t)] = \mathbb{E}\left[\int_0^t y(s)ds\right] = 0$. In addition, $(y(\tau_1), z(\tau_1))$ is equal to $(0, -Y)$ or $(0, Y)$ with probability $1/2$ for both, therefore with no loss of generality we can suppose that $(y(0), z(0)) = (0, -Y)$ or $(0, Y)$ with probability $1/2$ for both. Let us treat the case $y(0) = 0, z(0) = Y$. So, $\tau_0 = 0$ and $\theta_1 = \inf\{t > 0, z(t) = -Y, y(t) = 0\}$. We can assert that $\mathbb{E}\theta_1 = v^+(0, Y; 1)$ hence $\theta_1 < \infty$ a.s.. Next we define $\tau_1 = \inf\{t > \theta_1, z(t) = Y, y(t) = 0\}$ then $\mathbb{E}\tau_1 = v^+(0, Y; 1) + v^-(0, -Y; 1) = 2v^+(0, Y; 1)$. At time τ_1 the state of the system is again $(0, Y)$. So the sequence $\{\tau_n, n \geq 0\}$ is such that $\tau_n < \tau_{n+1}$ and in the interval (τ_n, τ_{n+1}) we have a cycle identical to $(0, \tau_1)$. Consider the variable $\int_0^{\tau_1} y(t)dt$. We have $\mathbb{E} \int_0^{\tau_1} y(t)dt = \mathbb{E} \int_0^{\theta_1} y(t)dt + \mathbb{E} \int_{\theta_1}^{\tau_1} y(t)dt = v^+(0, Y; y) + v^-(0, -Y; y)$. However if f is antisymmetric $f(-y, -z) = -f(y, z)$ we have $v^+(0, Y; f) = -v^-(0, -Y; f)$, therefore $\mathbb{E} \int_0^{\tau_1} y(t)dt = 0$. Hence, $\mathbb{E} \left(\int_0^{\tau_n} y(t)dt \right)^2 = \mathbb{E} \left(\sum_{j=0}^{n-1} \int_{\tau_j}^{\tau_{j+1}} y(t)dt \right)^2 = n \mathbb{E} \left(\int_0^{\tau_1} y(t)dt \right)^2$. Then

$$\frac{\mathbb{E} \left(\int_0^{\tau_n} y(t)dt \right)^2}{\mathbb{E}\tau_n} = \frac{\mathbb{E} \left(\int_0^{\tau_1} y(t)dt \right)^2}{\mathbb{E}\tau_1}.$$

Let $T > 0$ and N_T with $\tau_{N_T} \leq T < \tau_{N_T+1}$, $N_T = 0$ if $\tau_1 > T$, $\tau_0 = 0$ and then calculations lead to the following

$$\frac{\mathbb{E} \left(\int_0^{\tau_{N_T+1}} y(t)dt \right)^2}{\mathbb{E}\tau_{N_T+1}} = \frac{\mathbb{E} \left(\int_0^{\tau_1} y(t)dt \right)^2}{\mathbb{E}\tau_1}.$$

Next, we can justify that

$$\lim_{T \rightarrow \infty} \frac{1}{T} \mathbb{E} \left(\int_0^T y(t)dt \right)^2 = \lim_{T \rightarrow \infty} \frac{1}{T} \mathbb{E} \left(\int_0^{\tau_{N_T+1}} y(t)dt \right)^2 \quad (11)$$

and that we have a lower bound and an upper bound for (11), that is

$$\frac{\mathbb{E} \left(\int_0^{\tau_1} y(t)dt \right)^2}{\mathbb{E}\tau_1} \leq \frac{1}{T} \mathbb{E} \left(\int_0^{\tau_{N_T+1}} y(t)dt \right)^2 \leq \left(1 + \frac{\mathbb{E}\tau_1}{T} \right) \frac{\mathbb{E} \left(\int_0^{\tau_1} y(t)dt \right)^2}{\mathbb{E}\tau_1}.$$

Therefore

$$\lim_{T \rightarrow \infty} \frac{1}{T} \mathbb{E} \left(\int_0^T y(t)dt \right)^2 = \frac{\mathbb{E} \left(\int_0^{\tau_1} y(t)dt \right)^2}{\mathbb{E}\tau_1}. \quad (12)$$

Moreover, we can simplify (12), indeed we have

$$\mathbb{E} \left(\int_0^{\tau_1} y(t) dt \right)^2 = \mathbb{E} \left[\left(\int_0^{\theta_1} y(t) dt \right)^2 + \left(\int_{\theta_1}^{\tau_1} y(t) dt \right)^2 + 2 \int_0^{\theta_1} y(t) dt \int_{\theta_1}^{\tau_1} y(t) dt \right]$$

and therefore, we can justify that

$$\mathbb{E} \left(\int_0^{\tau_1} y(t) dt \right)^2 = 2 \left[\mathbb{E} \left(\int_0^{\theta_1} y(t) dt \right)^2 - \left(\int_0^{\theta_1} y(t) dt \right)^2 \right] = 2 \left[\mathbb{E} \left(\int_0^{\theta_1} y(t) dt \right)^2 - (v^+(0, Y; y))^2 \right].$$

Since $\mathbb{E}\tau_1 = 2v^+(0, Y; 1)$ we obtain

$$\lim_{T \rightarrow \infty} \frac{1}{T} \mathbb{E} \left(\int_0^T y(t) dt \right)^2 = \frac{\mathbb{E} \left(\int_0^{\theta_1} y(t) dt \right)^2 - (v^+(0, Y; y))^2}{v^+(0, Y; 1)}. \quad (13)$$

4. Numerical evidence in support of our result

In this section, we provide computational results which confirm our theoretical results. A C code has been written to simulate $(y(t), z(t))$. See [2] for the numerical scheme considered to do direct simulation. Let $T > 0$, $N \in \mathbb{N}$ and $\{t_n = n\delta t, 0 \leq n \leq N\}$ where $\delta t = \frac{T}{N}$. Then, to compute the left hand side of (4), we consider $MC \in \mathbb{N}$ and we generate MC numerical solutions of (SVI) $\{y^i(t), 0 \leq t \leq T, 1 \leq i \leq MC\}$ up to the time T . By the law of large numbers, we can approximate $\frac{1}{T} \mathbb{E} \left(\int_0^T y(s) ds \right)^2$ by $X_{MC} := \frac{1}{T} \frac{1}{MC} \sum_{i=1}^{MC} \left(\sum_{n=1}^N y^i(t_n) \delta t \right)^2$ and also $\frac{1}{T^2} \mathbb{E} \left(\int_0^T y(s) ds \right)^4$ by $X_{MC}^2 := \frac{1}{T^2} \frac{1}{MC} \sum_{i=1}^{MC} \left(\sum_{n=1}^N y^i(t_n) \delta t \right)^4$. Denote $s_X := \sqrt{X_{MC}^2 - (X_{MC})^2}$, we also know by the central limit theorem that $\frac{1}{T} \mathbb{E} \left(\int_0^T y(s) ds \right)^2 \in [X_{MC} - \frac{1.96s_X}{\sqrt{MC}}, X_{MC} + \frac{1.96s_X}{\sqrt{MC}}]$ with 95% of confidence. Similarly, to compute the right hand side of (4), we generate MC numerical long cycles. For each trajectory $\{y^i(t), t \geq 0\}$, we consider N_c^i the required number of time step to obtain a completed cycle. Denote $\delta_{MC} := \frac{1}{MC} \sum_{i=1}^{MC} \left(\sum_{n=0}^{N_c^i} y^i(t_n) \delta t \right)^2$, $\delta_{MC}^2 := \frac{1}{MC} \sum_{i=1}^{MC} \left(\sum_{n=0}^{N_c^i} y^i(t_n) \delta t \right)^4$, $\tau_{MC} := \frac{1}{MC} \sum_{i=1}^{MC} N_c^i \delta t$, $\tau_{MC}^2 := \frac{1}{MC} \sum_{i=1}^{MC} (N_c^i \delta t)^2$, $s_\delta := \sqrt{\delta_{MC}^2 - (\delta_{MC})^2}$ and $s_\tau := \sqrt{\tau_{MC}^2 - (\tau_{MC})^2}$. We also know that $\frac{\delta_{MC}}{\tau_{MC}} \in [\frac{\delta_{MC} - \frac{1.96s_\delta}{\sqrt{MC}}}{\tau_{MC} + \frac{1.96s_\tau}{\sqrt{MC}}}, \frac{\delta_{MC} + \frac{1.96s_\delta}{\sqrt{MC}}}{\tau_{MC} - \frac{1.96s_\tau}{\sqrt{MC}}}]$ with 95% of confidence. In table 1, is shown a comparison of the results obtained for X_{MC} and $\frac{\delta_{MC}}{\tau_{MC}}$ where $T = 500$, $\delta t = 10^{-4}$ and $MC = 5000$.

Remark 4.1. *From a numerical point of view, the behavior in long cycles is very relevant. Indeed, computing the left hand side of (4) is much more expensive in time compared to the right hand side (see the τ_{MC} -column of table 1).*

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$c_0 = 1, k = 1$				
Y	$X_{MC}, T = 500$	$\frac{\delta_{MC}}{\tau_{MC}}$	τ_{MC}	Relative error %
0.1	0.807 ± 0.031	0.834 ± 0.069	6.61 ± 0.11	3.2
0.2	0.649 ± 0.026	0.624 ± 0.047	8.74 ± 0.13	3.8
0.3	0.493 ± 0.020	0.464 ± 0.034	10.45 ± 0.16	5.8
0.4	0.361 ± 0.014	0.355 ± 0.026	12.12 ± 0.18	1.7
0.5	0.266 ± 0.011	0.257 ± 0.019	13.80 ± 0.21	3.3
0.6	0.195 ± 0.008	0.198 ± 0.014	16.15 ± 0.26	1.5
0.7	0.137 ± 0.005	0.149 ± 0.011	18.84 ± 0.31	8
0.8	0.103 ± 0.004	0.112 ± 0.008	22.80 ± 0.39	8
0.9	0.071 ± 0.003	0.086 ± 0.006	26.79 ± 0.47	15

Table 1: Monte-Carlo simulations to compare numerical solution of the left and right hand sides of (4), $T = 500$, $\delta t = 10^{-4}$ and $MC = 5000$.