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Nathanael Enriquez

► **To cite this version:**

| Nathanael Enriquez. An invariance principle for Azéma martingales. 2005. <hal-00013964>

**HAL Id: hal-00013964**

**<https://hal.archives-ouvertes.fr/hal-00013964>**

Submitted on 15 Nov 2005

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# AN INVARIANCE PRINCIPLE FOR AZÉMA MARTINGALES.

NATHANAËL ENRIQUEZ

**Abstract:** An invariance principle for Azéma martingales is presented as well as a new device to construct solutions of Emery's structure equations.

## 1. INTRODUCTION

In [9], Meyer raised the problem of finding “normal” martingales, i.e. martingales  $M_t$  on a filtered probability space  $(\Omega, \mathcal{F}, \mathcal{F}_t, P)$  satisfying  $\langle M, M \rangle_t = t$ , which would enjoy the chaotic representation property. This property, which is stronger than the previsible representation property, requires the direct sum of the chaoses of  $M$  to be equal to the whole space  $L^2(\Omega)$ . The only known examples of such martingales were the Brownian motion and the compensated Poisson process until Emery found in [5] a way to construct a whole family of examples, by introducing the beautiful device of structure equations. Among these examples, some of them, which are self-similar, appeared to be the generalization of two martingales considered earlier, for different purposes. One had been considered by Azéma in [1], and the other by Parthasarathy in an unpublished paper [13]. The whole class of these self-similar martingales is now usually called Azéma martingales. However, their behavior remains a little mysterious, especially, in the neighbourhood of 0. We refer the reader to the two lecture notes of Yor [16], and of Mansuy and Yor [8], where a chapter is devoted to the subject.

Inspired by some elementary remarks on renewal processes, we present an invariance principle for Azéma martingales, and propose a new device to construct solutions of Emery's structure equations. The principle of these approximations turns out to be quite unusual and differs from the truncation method one can find in dealing with Lévy processes, or from the discretization scheme already used in the subject by Meyer in [10] to construct solutions of Emery's structure equations. It is mainly based on the introduction of some randomness in the size of the jumps together with keeping the normal martingale property.

## 2. THE BASIC EXAMPLE: THE FIRST AZÉMA MARTINGALE

The aim of this section is to introduce, in a natural way, the device we use in the next section to approximate general Azéma martingales.

Let us start with the first Azéma martingale. It was obtained by Azéma in [1], by projecting a Brownian motion  $B_t$  starting from 0, on the filtration of  $\text{sign}(B)$ . This projection yields the following normal martingale

$$X_t = \sqrt{\frac{\pi}{2}} \text{sign}(B_t) \sqrt{(t - g_t)}$$

where  $g_t = \sup\{s : s \leq t, B_s = 0\}$ . We refer the reader to Chapter IV of Protter's book [14] for a comprehensive discussion about the martingale property of  $X_t$ . A clear justification of the projection property is given in an article of Azéma and Yor [2]: they use an elegant path decomposition of the Brownian motion on  $[0, t]$  involving the Brownian meander, seen as the renormalisation of the last incomplete excursion.

Now, it is well known, from P. Lévy, that  $2X_t^2/\pi t$  follows the arcsine law. On the other hand, there is an old result of Dynkin [4] presented also in Feller's book ([7] Chap. XIV.3), obtaining arcsine laws from asymptotic waiting times of renewal processes having non-integrable inter-arrival times. Now, the question is: can we associate a martingale to any given renewal process, exactly like the Azéma martingale is associated to the zeroes of the Brownian motion?

The following elementary proposition gives a positive answer to this question:

**Proposition 1.** *Let  $X_1, X_2, \dots$  be an iid sequence of positive random variables admitting a density function. We denote by  $S_n := \sum_{i=1}^n X_i$ , and  $N_t = \sup\{n : S_n \leq t\}$ . Now, we denote by  $\bar{F}(t) = P(X_1 > t)$ , the tail distribution function of the  $X_i$ 's.*

*Let us now introduce a sequence  $\varepsilon_1, \varepsilon_2, \dots$  of independent symmetric Bernoulli variables i.e. satisfying  $P(\varepsilon_i = 1) = \frac{1}{2}$  and  $P(\varepsilon_i = -1) = \frac{1}{2}$ .*

*The two following assertions hold:*

(i) *For every law of  $X_1$ , the process  $Z_t := \frac{\varepsilon_{N_t}}{\bar{F}(t - S_{N_t})}$  is a martingale with respect to its natural filtration.*

(ii) *The martingale  $Z_t$  is normal if and only if  $X_1$  follows the distribution of*

$$-\ln U + \frac{1}{2U^2} - \frac{1}{2}$$

*where  $U$  denotes a uniform variable on  $[0, 1]$ .*

Proof: The proof of (i) is based on the fact that, conditionally on the event  $(t - S_{N_t} = x)$ , the probability for  $Z_t$  to jump between  $t$  and  $t + h$  is equal to  $-\frac{\bar{F}'(x)}{\bar{F}(x)}h + o(h)$ , and if a jump occurs, its expectation is equal to  $-\frac{\varepsilon_{N_t}}{\bar{F}}$ . On the other hand, when there is no jump, the increment of the process, between  $t$  and  $t + h$ , is equal to  $-\varepsilon_{N_t} \frac{\bar{F}'(x)}{(\bar{F}(x))^2}h + o(h)$ . As a result, the expectations of the increments in both situations are balanced.

Let us now prove (ii). The martingale is normal if and only if the conditional expectation of the square of the increment of  $Z_t$  between  $t$  and  $t + h$  is always equal to  $h + o(h)$ . Again, conditionally on the event  $(t - S_{N_t} = x)$ , the probability for  $Z_t$  to jump between  $t$  and  $t + h$  is equal to  $-\frac{\bar{F}'(x)}{\bar{F}(x)}h + o(h)$ , and if a jump occurs, the expectation of its square is equal to  $\frac{1}{2}((1 + \frac{1}{\bar{F}})^2 + (1 - \frac{1}{\bar{F}})^2)$  which is  $1 + \frac{1}{\bar{F}^2}$ . Now, the absence of a jump contributes only in a  $o(h)$  in the expectation of the square of the increment of the process.

Therefore, the martingale is normal if and only if

$$-(\frac{1}{\bar{F}(x)} + \frac{1}{\bar{F}^3(x)})\bar{F}'(x) = 1.$$

This yields  $-\ln \bar{F}(x) + \frac{1}{2\bar{F}^2(x)} - \frac{1}{2} = x$ , so that the distribution function of  $X_1$  is the inverse of  $x \mapsto -\ln(1 - x) + \frac{1}{2(1 - x)^2} - \frac{1}{2}$ . This gives the result.  $\square$

*Remark:* We might have included in Proposition 1, analogous results concerning the sometimes called "second Azéma martingale" i.e.  $\sqrt{\frac{\pi}{2}}\sqrt{t - g_t} - l_t$ , where  $l_t$  denotes the

local time of zero at time  $t$ . Indeed, for every law of  $X_1$ ,  $\tilde{Z}_t := \frac{1}{\bar{F}(t - S_{N_t})} - N_t$  is a martingale which is normal if and only if  $X_1$  follows the distribution of  $\frac{1}{2U^2} - \frac{1}{2}$ .

But we shall not insist on that fact since this second Azéma martingale is neither markovian nor enjoys the chaotic representation property, and does not enter the class of Azéma martingales solving structure equations.

Let us denote by  $Z_t^{(-1)}$  the normal martingale  $Z_t$  of (ii). We note that  $Z_t^{(-1)}$  is a Markov process, and more precisely,

**Proposition 2.** *The process  $Z_t^{(-1)}$  is a Markov process with generator:*

$$L^{(-1)}g(x) = \frac{\frac{1}{2}(g(-1) + g(1)) - g(x) + xg'(x)}{1 + x^2}$$

Proof: From Proposition 1 (ii), we deduce that

$$-\ln \bar{F}(x) + \frac{1}{2\bar{F}^2(x)} - \frac{1}{2} = x$$

which implies, by differentiation:

$$-\frac{\bar{F}'(x)}{\bar{F}(x)} = \frac{1}{1 + \frac{1}{\bar{F}^2(x)}}$$

Hence, the rate of jump of the process at time  $t$ , which is equal to the value of the function  $-\frac{\bar{F}'}{\bar{F}}$  at  $(t - S_{N_t})$ , is precisely  $\frac{1}{1 + (Z_t^{(-1)})^2}$ , and the speed of the trajectory at a time  $t$

between two jumps is equal to  $-\varepsilon_{N_t} \frac{\bar{F}'(t - S_{N_t})}{\bar{F}^2(t - S_{N_t})}$  which is precisely  $\frac{Z_t^{(-1)}}{1 + (Z_t^{(-1)})^2}$ .  $\square$

**Corollary 1.** *The process  $Z_t^{(-1,n)} := \frac{Z_{nt}^{(-1)}}{\sqrt{n}}$  is a Markov process with generator:*

$$L^{(-1,n)}g(x) = \frac{\frac{1}{2}(g(-\frac{1}{\sqrt{n}}) + g(\frac{1}{\sqrt{n}})) - g(x) + xg'(x)}{\frac{1}{n} + x^2}$$

Proof: Consider a differentiable function  $g$ . Introduce the function  $h(x) = g(\frac{x}{\sqrt{n}})$ .

The image by  $L^{(-1,n)}$  of the function  $g$  is given by:

$$L^{(-1,n)}g = n \cdot (L^{(-1)}h)(\sqrt{n} \cdot x)$$

A direct computation gives the announced expression.  $\square$

This last corollary already gives an idea of the proximity between  $Z_t^n$  and the first Azéma martingale. Indeed, as writes Emery in [5] about Azéma martingales: “formally (this means: informally !) it should be Markov, with generator

$$Lg(x) = \begin{cases} \frac{g((1 + \beta)x) - g(x) - \beta xg'(x)}{\beta^2 x^2} & \text{if } x \neq 0 \\ \frac{1}{2}g''(x) & \text{if } x = 0 \end{cases}$$

(in the case of the first Azéma martingale, the parameter  $\beta$  is equal to  $-1$ )

This is the aim of the next section which is to prove, among other things, the convergence in distribution in the sense of the Skorohod topology of  $\frac{Z_{nt}^{(-1)}}{\sqrt{n}}$  towards the first Azéma martingale.

## 3. AN INVARIANCE PRINCIPLE

Let us begin with the definition of the Azéma martingales.

Following Emery, let us introduce a real parameter  $\beta$ . The structure equation

$$(1) \quad d[X, X]_t = dt + \beta X_{t-} dX_t$$

has a unique solution.

The existence was proven in a note of Meyer [10], and the uniqueness of the solution was proven by Emery in [5] (for  $\beta < 0$ ) and [6] (for  $\beta > 0$ ). The solution of this structure equation is called the Azéma martingale with parameter  $\beta$ .

When a jump occurs at time  $t$ , the value of the process changes by some prescribed factor. More precisely,  $X_t = (1 + \beta)X_{t-}$ .

- The case  $\beta = 0$  corresponds to Brownian motion: the continuous normal martingale.
- The case  $\beta = -1$  is solved by the martingale  $(\text{sign } B_t)\sqrt{2(t - g_t)}$  mentioned in the previous section.
- The case  $\beta = -2$  is solved by Parthasarathy's martingale satisfying  $|X_t| = \sqrt{t}$ , and who changes of sign according to a Poisson point process with intensity  $dt/4t$ .
- In the case  $\beta \leq -2$ , the solution, at least starting at a non-null point, is rather easy to define, since  $|X_t|$  always goes away from 0. In proving the uniqueness of the solution, Emery in [5] shows how to define the process starting from 0 by using a self-similarity argument and a representation of the process by using a time-changed Poisson process.

For  $\beta > -2$ , things go in a more complicated way since the process  $X_t$  reaches 0 in finite time, and the above formal generator cannot be of much help after that time.

Let us mention finally that the problem of the chaotic representation property still remains a challenging open question for parameters  $\beta$  which do not belong to  $[-2, 0]$ .

Now, let us turn to our main purpose concerning the statement of an invariance principle for these processes. This question is quite natural, since Azéma martingales are self-similar in the sense that, for all  $\lambda > 0$  the processes  $X_t$  and  $X_{\lambda^2 t}/\lambda$  are equal in distribution (one can indeed easily check that the process  $X_{\lambda^2 t}/\lambda$  solves the right structure equation, and conclude by uniqueness in law of the solution).

In order to define the process which will be, in our statement, at the origin of such a limit theorem, we generalize the definition of  $Z_t^{(-1)}$  introduced in the previous section, and define for all real parameters  $\beta$  the Markov process  $Z_t^{(\beta)}$  having generator:

$$L^{(\beta)} = \frac{\frac{1}{2}(g((1 + \beta)x - 1) + g((1 + \beta)x + 1)) - g(x) - \beta x g'(x)}{1 + \beta^2 x^2}$$

Let us denote the process  $Z_t^{(\beta, n)} := \frac{Z_{nt}^{(\beta)}}{\sqrt{n}}$ . It has generator:

$$L^{(\beta, n)} = \frac{\frac{1}{2}(g((1 + \beta)x - \frac{1}{\sqrt{n}}) + g((1 + \beta)x + \frac{1}{\sqrt{n}})) - g(x) - \beta x g'(x)}{\frac{1}{n} + \beta^2 x^2}$$

We can now state our main result:

**Theorem 1.** *The sequence of processes  $Z_t^{(\beta, n)} := \frac{Z_{nt}^{(\beta)}}{\sqrt{n}}$  converges to the Azéma martingale with parameter  $\beta$ , in the sense of the weak convergence for the Skorohod topology.*

Proof: The image by  $L^{(\beta, n)}$ , of  $x$  and  $x^2$ , being respectively 0 and 1, we deduce that  $Z_t^{(f, n)}$  are normal martingales. Therefore, by Rebolledo's theorem (see [15], II.3.1), this sequence is tight for the weak convergence in the Skorohod topology. All we have to prove is that all

limiting points of the sequence of processes  $Z_t^{(\beta,n)}$  satisfy the structure equation (1). The uniqueness in law of the solution of (1) allows to conclude.

Let us define at a jump time  $t$  of the process  $Z_t^{(\beta,n)}$ , the symmetric Bernoulli variable  $\varepsilon_t^{(\beta,n)}$  defined by

$$Z_t^{(\beta,n)} - Z_{t^-}^{(\beta,n)} = \beta Z_t^{(\beta,n)} + \frac{\varepsilon_t^{(\beta,n)}}{\sqrt{n}}$$

Let us consider  $[Z^{(\beta,n)}, Z^{(\beta,n)}]_t$ . (We replace, in the sequel, the superscript  $(\beta, n)$  by  $n$ .)

$$\begin{aligned} [Z^n, Z^n]_t &= \sum_{s \leq t} (Z_s^n - Z_{s^-}^n)^2 = \sum_{s \leq t} (\beta Z_{s^-}^n + \frac{\varepsilon_s^n}{\sqrt{n}})(Z_s^n - Z_{s^-}^n) \\ &= \int_0^t \beta Z_{s^-}^n dZ_s^n - \int_0^t \beta Z_{s^-}^n 1_{\Delta Z_s^n \neq 0} dZ_s^n + \sum_{s \leq t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} (\beta Z_{s^-}^n + \frac{\varepsilon_s^n}{\sqrt{n}}) \\ &= \int_0^t \beta Z_{s^-}^n dZ_s^n + t - \int_0^t \frac{ds}{1 + n(\beta Z_{s^-}^n)^2} + \sum_{s \leq t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} (\beta Z_{s^-}^n + \frac{\varepsilon_s^n}{\sqrt{n}}) \end{aligned}$$

Our task, now, is to show that the two last terms vanish when  $n$  goes to infinity, whereas the other terms converge to their analogous quantity for the limiting process, so that every limiting process will solve the structure equation satisfied by an Azéma martingale, and will therefore be identified with the Azéma martingale with parameter  $\beta$ .

Now we proceed, like in Meyer's proof of the existence of solutions of structure equations, and introduce successively:

- a probability space  $(\Omega, \mathcal{G}, Q)$ , obtained from the Skorohod representation theorem, on which a subsequence of processes  $X_s^{\sigma(n)}$ , following the law of an extracted sequence of  $Z_s^n$  converges almost surely in  $\lambda \otimes Q$ . For ease of reading, we keep writing  $Z_s^n$  instead of  $X_s^{\sigma(n)}$ . In other words, we suppose that, for almost every time  $t$ ,  $Z_t^n$  converges almost surely.

- a sequence of stopping times  $T_N^n = \inf\{t : |Z_t^n| > N\}$  where  $N$  is large enough, so that  $Q(T_N^n \leq t)$  is smaller than any prescribed  $\epsilon$ , uniformly in  $n$  (Doob's inequality makes it possible to do so, since the processes  $Z_t^n$  have the same variance).

Let us compare, in a first step,  $[Z^n, Z^n]_{T_N^n \wedge t}$  with the sum of the square of its increments along some subdivision  $(t_i)$  of the interval  $[0, t]$ :

$$[Z^n, Z^n]_{T_N^n \wedge t} - \sum_i (Z_{T_N^n \wedge t_{i+1}}^n - Z_{T_N^n \wedge t_i}^n)^2 = 2 \sum_i \int_{t_i}^{t_{i+1}} (Z_{T_N^n \wedge s}^n - Z_{T_N^n \wedge t_i}^n) dZ_{T_N^n \wedge s}^n$$

The compensator of  $Z_{T_N^n \wedge s}^n$  being equal to  $T_N^n \wedge s$ , we deduce,

$$E[(Z^n, Z^n)_{T_N^n \wedge t} - \sum_i (Z_{T_N^n \wedge t_{i+1}}^n - Z_{T_N^n \wedge t_i}^n)^2] \leq 4E[\sum_i \int_{t_i}^{t_{i+1}} (Z_{T_N^n \wedge s}^n - Z_{T_N^n \wedge t_i}^n)^2 ds]$$

which in turn is smaller than  $4E[\sum_i \int_{t_i}^{t_{i+1}} (Z_{T_N^n \wedge s}^n - Z_{T_N^n \wedge t_i}^n)^2 ds] \leq \frac{4}{3} \sum_i (t_{i+1} - t_i)^3$ . (In this first step, stopping the martingales at  $T_N^n$  just ensures the integrability we need. In the further steps, this stopping argument will be more widely used.)

As a conclusion, we deduce that the "discrete quadratic Riemann sums" of  $Z^n$  approximate  $[Z^n, Z^n]$  uniformly, in probability. Now, since the corresponding sums for  $Z$  approximate  $[Z, Z]$  in probability, it just remains to notice that we can choose the time of our subdivisions among the times  $t$  for which  $Z_t^n$  converges almost surely to  $Z_t$  so that the quadratic sums of  $Z^n$  will converge almost surely and therefore in probability to their corresponding analog for the process  $Z$ .

In a second step, we deal with the second term  $\int_0^t Z_{s-}^n dZ_s^n$ . We prove its convergence to the corresponding quantity for  $Z$ , exactly the same way as in Meyer's proof.

Now, we have to prove the convergence to 0 in probability of both remaining terms  $\int_0^t \frac{ds}{1+n(\beta Z_{s-}^n)^2}$  and  $\sum_{s \leq t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} (\beta Z_{s-}^n + \frac{\varepsilon_s^n}{\sqrt{n}})$ .

Again, it suffices to prove it for these quantities when  $t$  is replaced by  $T_N^n \wedge t$ .

This result will be a consequence of the following key lemma, which gives some uniform control (in  $n$ ) on the time spent by the processes  $Z_t^n$  near the origin, and on the number of their jumps:

**Lemma 1.** *The two following convergence results hold: for all  $N > 0$ , and all  $t > 0$ ,*

(i)  $\forall \epsilon > 0, \exists \delta > 0, \exists n_0 \in \mathbb{N}$ , such that  $\forall n \geq n_0$ ,

$$P\left(\int_0^{T_N^n \wedge t} 1_{Z_s^n \in [-\delta, \delta]} ds > \epsilon\right) < \epsilon$$

(ii)  $\frac{N_{T_N^n \wedge t}^n}{n}$  converges to 0 in probability.

Proof: We start with (i). We will use the following notation for a constant which will appear repeatedly:  $C = \max(1, 1 + |\beta|)$ .

Let us fix  $\varepsilon > 0$ , and, for some  $\delta \in ]0, \frac{N}{2C}[$ , consider the successive stopping times:

- $\tau_1^n = \inf\{t : |Z_t^n| > 2\delta\}$
  - $\tau_2^n = \inf\{t > \tau_1^n : |Z_t^n| < \delta \text{ or } |Z_t^n| > N\}$
  - If  $\tau_2^n \neq T_N^n$ , we define  $\tau_3^n := \inf\{t > \tau_2^n : |Z_t^n| > 2\delta\}$
- and more generally for  $i \geq 2$ ,

- $\tau_{2i}^n := \inf\{t > \tau_{2i-1}^n : |Z_t^n| < \delta \text{ or } |Z_t^n| > N\}$
- If  $\tau_{2i}^n \neq T_N^n$ , then  $\tau_{2i+1}^n := \inf\{t > \tau_{2i}^n : |Z_t^n| > 2\delta\}$

Let us denote by  $K$ , the *random* integer  $K := \inf\{i \geq 1 : \tau_{2i}^n = T_N^n\}$ .

From these definitions, we get obviously:

$$\int_0^{T_N^n \wedge t} 1_{Z_s^n \in [-\delta, \delta]} ds \leq \tau_1^n + (\tau_3^n - \tau_2^n) + \dots + (\tau_{2K-1}^n - \tau_{2K-2}^n)$$

Now, we can bound the expectation of the right hand term, by first noticing that, by the optional sampling theorem applied to the martingale  $(Z_t^n)^2 - t$ ,

$$\begin{aligned} \forall i \geq 1, \quad E[\tau_{2i+1}^n - \tau_{2i}^n \mid K > i] &= E[(Z_{\tau_{2i+1}^n}^n)^2 - (Z_{\tau_{2i}^n}^n)^2 \mid K > i] \\ &\leq E[(Z_{\tau_{2i+1}^n}^n)^2 \mid K > i] \leq 4 \max(1, (1 + \beta)^2) \delta^2 = 4C^2 \delta^2 \end{aligned}$$

And secondly, we can estimate  $P(K = i + 1 \mid K > i)$ , by exploiting that

$$E[Z_{\tau_{2i+2}^n}^n \mid K > i] = E[Z_{\tau_{2i+1}^n}^n \mid K > i].$$

But, on one hand,  $|Z_{\tau_{2i+2}^n}^n| \geq 2\delta$  and, on the other hand,  $E[Z_{\tau_{2i+2}^n}^n \mid K > i]$  decomposes into  $E[1_{K=i+1} Z_{\tau_{2i+2}^n}^n] + E[1_{K>i+1} Z_{\tau_{2i+2}^n}^n]$ , with

$$|E[1_{K=i+1} Z_{\tau_{2i+2}^n}^n]| \leq NC.P(K = i + 1 \mid K > i)$$

and

$$|E[1_{K>i+1} Z_{\tau_{2i+2}^n}^n]| \leq \delta$$

Therefore,

$$P(K = i + 1 \mid K > i) \geq \frac{\delta}{NC}$$

which implies that  $P(K \geq i) \leq (1 - \frac{\delta}{NC})^{i-1}$ .

Now, writing

$$E[\tau_1^n + (\tau_3^n - \tau_2^n) + \dots + (\tau_{2K-1}^n - \tau_{2K-2}^n)] = \sum_{i \geq 1} E[(\tau_{2i-1}^n - \tau_{2i-2}^n) 1_{K \geq i}]$$

and, using the strong Markov property at time  $\tau_{2i-1}^n$ , we get

$$E[\tau_1^n + (\tau_3^n - \tau_2^n) + \dots + (\tau_{2K-1}^n - \tau_{2K-2}^n)] \leq 4C^2 \delta^2 \sum_{i \geq 1} (1 - \frac{\delta}{NC})^{i-1} = 4NC^3 \delta$$

We just have to choose  $\delta$  such that  $4NC^3 \delta = \varepsilon^2$  and use Markov inequality, to get (i).

To prove (ii), we will strongly exploit (i). To make the reading easier, we will fix  $t$  equal to 1. Let us divide the interval  $[0, t] = [0, 1]$  into  $n$  intervals of the type  $I_k^n = [\frac{k}{n}, \frac{k+1}{n}]$ , where  $0 \leq k \leq n-1$ .

Now,  $\varepsilon$  being fixed, we introduce  $\delta$  like in (i), and define the stopping times:

$$\sigma_k^n = \inf\{s \in I_k^n, |Z_s^n| > \delta\}$$

with  $\sigma_k^n = \infty$  if  $\forall s \in I_k^n, |Z_s^n| \leq \delta$ .

From (i), we get that, for  $n$  large enough, with probability bigger than  $1 - \varepsilon$ , for more than  $(1 - \varepsilon)n$  integers  $k$ , between 0 and  $n-1$ , we have  $\sigma_k^n < \infty$ .

Let us introduce now the stopping time  $\tau_k^n := \inf\{s > \sigma_k^n \mid |Z_s^n - Z_{s-}^n| \neq 0\}$ , with  $\tau_k^n = \infty$  whenever  $\sigma_k^n = \infty$ .

Now, for  $n$  large enough, the quantity  $P(|\tau_k^n - \sigma_k^n| < \frac{2}{n} \mid \sigma_k^n < \infty)$  can be bounded uniformly in  $k$  by some sequence depending on  $n$ , which converges to 0. Indeed, if at some time  $s$ ,  $|Z_s^n| > \delta$ , and if moreover there is no jump between  $s$  and  $s + \frac{2}{n}$ , then  $|Z_s^n|$  remains bigger than  $\delta/2$  for  $n$  large enough. Now, on the time interval  $[s, s + \frac{2}{n}]$ , the intensity of the jump remains smaller than  $\frac{1}{\beta^2(\delta/2)^2}$ .

We deduce that, conditionally on the event  $(\sigma_k^n < \infty)$ ,  $|\tau_k^n - \sigma_k^n| \wedge \frac{2}{n}$  is stochastically bigger than  $\mathbf{e}_{\frac{4}{\beta^2 \delta^2}} \wedge \frac{2}{n}$ , where  $\mathbf{e}_{\frac{4}{\beta^2 \delta^2}}$  is an exponential variable having parameter  $\frac{4}{\beta^2 \delta^2}$ .

So that, for  $n$  large enough,

$$P(|\tau_k^n - \sigma_k^n| > \frac{2}{n} \mid \sigma_k^n < \infty) > \exp(-\frac{8}{\beta^2 \delta^2 n})$$

Now,  $|\tau_k^n - \sigma_k^n| > \frac{2}{n}$  implies that there is no jump inside the time interval  $I_{k+1}^n$ .

So, we deduce that, for  $n$  large enough, with probability bigger than  $1 - 2\varepsilon$ , the process  $Z_s^n$  will not jump on more than  $(1 - 2\varepsilon)n$  intervals among the  $I_k^n$ 's.

Finally, since the rate jump of the process is permanently less than  $n$  and the length of each  $I_k^n$  is equal to  $\frac{1}{n}$ , for  $n$  large enough, we get that, with an arbitrarily high probability, there are not more than  $4\varepsilon n$  jumps on the remaining  $2\varepsilon n$  intervals  $I_k^n$  which may have some jumps.  $\square$

*End of the proof of Theorem 1:* The (i) part of Lemma 1 gives immediately the convergence in probability to 0, of  $\int_0^{T_N^n \wedge t} \frac{ds}{1 + n(\beta Z_{s-}^n)^2}$ .



$$\text{Now, } \sum_{s \leq T_N^n \wedge t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} (\beta Z_{s^-}^n + \frac{\varepsilon_s^n}{\sqrt{n}}) = \frac{N_{T_N^n \wedge t}^n}{n} + \beta \sum_{s \leq T_N^n \wedge t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} Z_{s^-}^n$$

Part (i) of Lemma 1 implies the convergence to 0 of the first term of this sum, and we bound the  $L^2$ -norm of the second term:

$$E\left[\left(\sum_{s \leq T_N^n \wedge t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} Z_{s^-}^n\right)^2\right] = \frac{1}{n} E\left[\sum_{s \leq T_N^n \wedge t} 1_{\Delta Z_s^n \neq 0} (Z_{s^-}^n)^2\right] \leq N^2 \cdot E\left[\frac{N_{T_N^n \wedge t}^n}{n}\right]$$

Now, for all  $\varepsilon > 0$ ,

$$E\left[\frac{N_{T_N^n \wedge t}^n}{n}\right] \leq \varepsilon + E\left[\frac{N_{T_N^n \wedge t}^n}{n} \cdot 1_{\left(\frac{N_{T_N^n \wedge t}^n}{n} > \varepsilon\right)}\right] \leq \varepsilon + E\left[\left(\frac{N_{T_N^n \wedge t}^n}{n}\right)^2\right]^{\frac{1}{2}} P\left(\frac{N_{T_N^n \wedge t}^n}{n} > \varepsilon\right)^{\frac{1}{2}}$$

Now, coupling the jump times of the process  $Z_t^n$  with the points of a Poisson process with intensity  $n$ , we are able to bound  $E\left[\left(\frac{N_{T_N^n \wedge t}^n}{n}\right)^2\right]$  uniformly in  $n$ , by  $t^2$ .

We conclude that  $\sum_{s \leq T_N^n \wedge t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} (\beta Z_{s^-}^n + \frac{\varepsilon_s^n}{\sqrt{n}})$  converges to 0 in probability, and obtain

Theorem 1. □

*Remark:* We might have replaced the randomness of the jumps, having distribution  $\frac{1}{2}(\delta_{-1} + \delta_1)$  in the case of the process  $Z_t^{(\beta,1)}$ , by any other one, having compact supported distribution with mean 0 and variance 1. In other words, the result of Theorem 1 remains valid, if we take, instead of  $Z_t^{(\beta,1)}$ , any process having generator

$$\frac{\left(\int_{\mathbb{R}} g((1+\beta)x+y)\mu(dy)\right) - g(x) - \beta x g'(x)}{1 + \beta^2 x^2}$$

where  $\mu$  is a compact supported measure, with mean 0 and variance 1. Indeed, one can easily check that we still deal with normal martingales.

#### 4. AN EXTENSION TO STRUCTURE EQUATIONS

Actually, Meyer is able to construct solutions of the more general structure equation

$$(2) \quad d[X, X]_t = dt + f(X_{t^-})dX_t$$

where  $f$  is an arbitrary continuous function.

The strategy of Meyer consists in solving a discretized structure equation, which leads to a cascade of second degree polynomial equations satisfied by the discretized increments, having always exactly two solutions of opposite signs, which are chosen with respect to the (unique) probability preserving the martingale property of the constructed discrete process.

We want to show in this section how our approach can be used to exhibit a solution of the structure equation (which may be different from Meyer's in the non-uniqueness cases). Inspired by the processes of previous section, we introduce a sequence of normal martingales  $Z_t^{(f,n)}$  whose limiting points in the sense of the weak convergence for the Skorohod topology will satisfy the structure equation (2).

However, although our approach has the advantage to give an invariance principle in the self-similar case of previous section, nonetheless the construction of structure equations for general continuous functions induces delicate discussions on the time spent near the zeroes of the function  $f$ , and we did not work them out for a general continuous function  $f$ . So, our aim, in this section, is more to exemplify things, than to arrive to the most

general statement, which might be valid for general continuous functions, but we will let this question open here.

We define  $Z_t^{(f,n)}$ , as the Markov process starting at 0, and having generator:

$$L^{(f,n)}g(x) = \frac{\frac{1}{2} \left( g(x + f(x) + \frac{1}{\sqrt{n}}) + g(x + f(x) - \frac{1}{\sqrt{n}}) \right) - g(x) - f(x)g'(x)}{\frac{1}{n} + f(x)^2}$$

**Proposition 3.** *Suppose  $f$  is a continuous function with isolated zeroes. Suppose, in addition, that, in the neighbourhood of each zero  $x_j$  of  $f$ ,  $f(x_j + h) = o(\sqrt{h})$ .*

*Then, the limiting points of the sequence of processes  $Z_t^{(f,n)}$ , in the sense of the weak convergence for the Skorohod topology, satisfy the structure equation (2).*

Proof: Like in the previous section, we start by noticing that the image by  $L^{(f,n)}$ , of  $x$  and  $x^2$ , is respectively 0 and 1. Indeed,

$$\frac{1}{2} \left( \left( x + f(x) + \frac{1}{\sqrt{n}} \right) + \left( x + f(x) - \frac{1}{\sqrt{n}} \right) \right) - x - f(x) = 0$$

and

$$\frac{1}{2} \left( \left( x + f(x) + \frac{1}{\sqrt{n}} \right)^2 + \left( x + f(x) - \frac{1}{\sqrt{n}} \right)^2 \right) - x^2 - 2xf(x) = \frac{1}{n} + f(x)^2$$

We deduce that  $Z_t^{(f,n)}$  are normal martingales. Therefore, by Rebolledo's theorem (see [15], II.3.1), this sequence is tight for the weak convergence in the Skorohod topology.

Now, we follow the proof of Theorem 1, and write:

$$[Z^n, Z^n]_t = \int_0^t f(Z_{s-}^n) dZ_s^n + t - \int_0^t \frac{ds}{1 + nf(Z_{s-}^n)^2} + \sum_{s \leq t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} (f(Z_{s-}^n) + \frac{\varepsilon_s^n}{\sqrt{n}})$$

Applying again Skorohod representation theorem, we are able to prove the convergence of the terms  $[Z^n, Z^n]_t$  and  $\int_0^t f(Z_{s-}^n) dZ_s^n$  to the analogous quantities for the limiting process.

Now, the key point is to check the analogous property to Part (i) of Lemma 1.

In other words, we have to prove, that, for every  $x_j \in [-N, N]$  (there are a finite number of them),  $\forall \epsilon > 0, \exists \delta_j > 0, \exists n_0 \in \mathbb{N}$ , such that  $\forall n \geq n_0$ ,

$$P\left(\int_0^{T_N^n \wedge t} 1_{Z_s^n \in [x_j - \delta_j, x_j + \delta_j]} ds > \epsilon\right) < \epsilon$$

We then introduce the same stopping times as in Lemma 1, and the differences with the proof of Lemma 1 come out when we estimate  $E[\tau_{2i+1}^n - \tau_{2i}^n \mid K > i]$ :

$$\begin{aligned} \forall i \geq 1, E[\tau_{2i+1}^n - \tau_{2i}^n \mid K > i] &= E[(Z_{\tau_{2i+1}^n}^n)^2 - (Z_{\tau_{2i}^n}^n)^2 \mid K > i] \\ &\leq E[(Z_{\tau_{2i+1}^n}^n)^2 \mid K > i] \leq (2\delta_j + \sup_{[-2\delta_j, 2\delta_j]} |f(x)|)^2 = o(\delta_j) \end{aligned}$$

whereas  $P(K = i + 1 \mid K > i)$  remains bigger than a constant times  $\delta_j$ .

As a result,  $E[\tau_1^n + (\tau_3^n - \tau_2^n) + \dots + (\tau_{2K-1}^n - \tau_{2K-2}^n)]$  is bounded by a constant times  $\frac{(2\delta_j + \sup_{[-2\delta_j, 2\delta_j]} |f(x)|)^2}{\delta_j} = o(1)$ .

As in Lemma 1, we choose  $\delta_i$  such that this fraction is equal to  $\varepsilon^2$ .

We finish the proof the same way as for Theorem 1.  $\square$

*Remark 1:* The framework of Proposition 3 includes the case of asymmetric Azéma martingales corresponding to a function  $f$  of the form  $ax1_{x \geq 0} + bx1_{x < 0}$ . In this case, Phan proved in [12] that the solution of the structure equation is unique. From this result, we

deduce that these processes are self-similar and Proposition 3 provides, in this case, an invariance principle.

*Remark 2:* Our assumptions in Proposition 3 are obviously not optimal: for instance, if  $f$  is null, we are back to the case of a simple random walk, and the convergence to the corresponding solution of the structure equation which is, in this case, the Brownian motion is obviously valid. It would not be hard to check that one can allow  $f$  to be null on a locally finite number of intervals, and satisfy the assumptions of Proposition 3 on the complement of the union of these intervals. In this case, the term  $\sum_{s \leq t} 1_{\Delta Z_s^n \neq 0} \frac{\varepsilon_s^n}{\sqrt{n}} (f(Z_{s^-}^n) + \frac{\varepsilon_s^n}{\sqrt{n}})$  will not vanish in the limit and will “alternatively” contribute with  $\int_0^t \frac{f(Z_{s^-}^n)^2}{\frac{1}{n} + f(Z_{s^-}^n)^2} ds$  to build a compensator equal to  $t$ .

It is also very likely that the condition on  $f$  to be  $o(\sqrt{h})$  near its zeroes is not optimal (even if our attempts to relax it failed).

## 5. CONSTRUCTION OF THE PROCESS $Z_t^{(\beta, n)}$

We might just present the construction of  $Z_t^{(\beta, 1)}$  and rescale it, but we preferred to keep the parameter  $n$  in evidence, in order to point out clearly where the randomization of the jumps appears.

Let us introduce the functions  $\bar{F}$  and  $H$  defined on  $\mathbb{R} \times \mathbb{R}_+$ , in the following way: for any  $(x_0, t) \in \mathbb{R} \times \mathbb{R}_+$ ,

-  $\bar{F}(x_0, t)$  denotes the probability that the first jump of the process  $Z_t^{(\beta, n)}$ , starting at  $x_0$ , takes place after time  $t$ .

-  $H(x_0, t)$  is the value of  $Z_t^{(\beta, n)}$ , starting at  $x_0$ , conditioned on the event that the first jump of the process takes place after time  $t$ .

From the expression of the generator, we deduce the following system of differential equations:

$$\begin{cases} \frac{1}{\bar{F}} \frac{\partial \bar{F}}{\partial t}(x_0, t) &= -\frac{1}{\frac{1}{n} + \beta^2 H^2(x_0, t)} \\ \frac{\partial H}{\partial t}(x_0, t) &= -\frac{\beta H(x_0, t)}{\frac{1}{n} + \beta^2 H^2(x_0, t)} \end{cases}$$

with the boundary conditions  $\bar{F}(x_0, 0) = 1$  and  $H(x_0, 0) = x_0$ .

After noticing that  $\frac{1}{\beta} \frac{\partial \ln H}{\partial t} = \frac{\partial \ln \bar{F}}{\partial t}$ , we deduce that

$$H(x_0, t) = x_0 (\bar{F}(x_0, t))^\beta$$

From this fact, we deduce the following autonomous equation for  $\bar{F}$ :

$$-\frac{1}{n} \frac{1}{\bar{F}} \frac{\partial \bar{F}}{\partial t} - \beta^2 x_0^2 \bar{F}^{2\beta-1} \frac{\partial \bar{F}}{\partial t} = 1$$

which yields

$$-\frac{1}{n} \ln \bar{F} - \frac{\beta x_0^2}{2} (\bar{F}^{2\beta} - 1) = t$$

Therefore, the first jump of the process  $Z_t^{(\beta, n)}$ , starting at  $x_0$ , follows the law of

$$-\frac{1}{n} \ln U - \frac{\beta x_0^2}{2} (U^{2\beta} - 1)$$

where  $U$  is a uniform random variable on  $[0, 1]$ . We notice here that the special case  $n = 1$ ,  $\beta = -1$  and  $|x_0| = 1$  allows us to recover the law of the random variable we discussed in Proposition 1 (ii).

Denote  $T_0 = 0$  and  $T_i$  the  $i$ -th time of jump of the process  $Z_t^{(\beta, n)}$ . We can describe the process  $Z_t^{(\beta, n)}$  as follows: there exists a sequence  $(U_i)_{i \geq 1}$  of independent uniform variables on  $[0, 1]$ , and a sequence  $(\varepsilon_i)_{i \geq 1}$  of independent symmetric Bernoulli variables, such that, for all  $i \geq 1$ , (we omit here the superscript  $(\beta, n)$ .)

$$\begin{cases} T_i &= T_{i-1} - \frac{1}{n} \ln U_i - \beta \frac{Z_{T_{i-1}}^2}{2} (U_i^{2\beta} - 1) \\ Z_{T_i^-} &= Z_{T_{i-1}} U_i^\beta \\ Z_{T_i} &= Z_{T_i^-} (1 + \beta) + \frac{1}{\sqrt{n}} \varepsilon_i \end{cases}$$

On the time interval  $[T_{i-1}, T_i[$ , the trajectory  $(t, Z_t)$  can be parametrized by  $(t(s), Z_{t(s)})_{0 \leq s \leq 1 - U_i}$  in the following way: for all  $s \in [0, 1 - U_i[$ ,

$$\begin{cases} t(s) &= T_{i-1} - \frac{1}{n} \ln(1 - s) - \beta \frac{Z_{T_{i-1}}^2}{2} ((1 - s)^{2\beta} - 1) \\ Z_{t(s)} &= Z_{T_{i-1}} (1 - s)^\beta \end{cases}$$

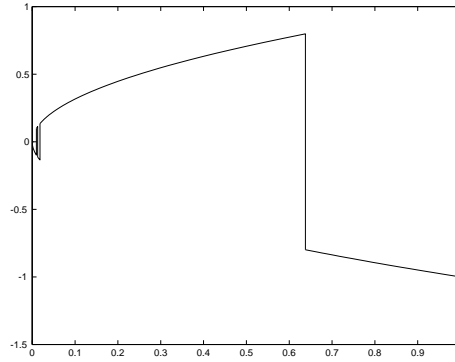
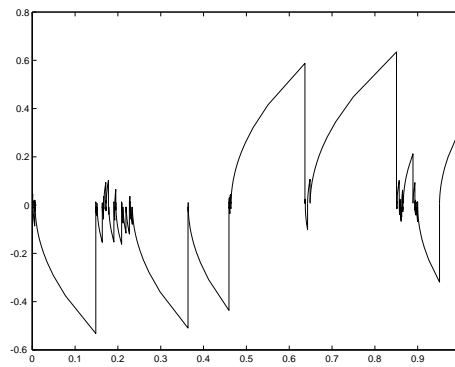
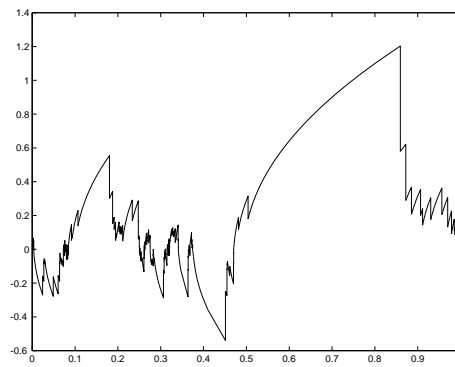
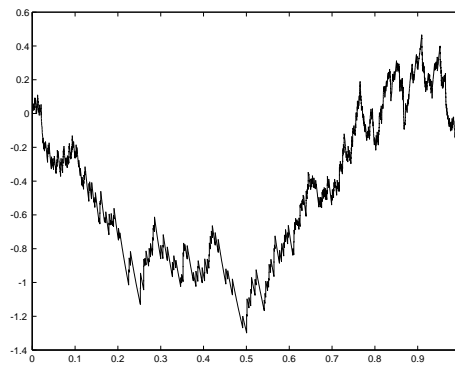
*Remark:* From this description of  $Z_t^{(\beta, n)}$ , we can easily deduce a coupling between  $Z_t^{(-1, n)}$  and  $Z_t^{(\beta, n)}$ , for all  $n$  and all parameters  $\beta$ . But, unfortunately, this coupling cannot be transported at the level of the corresponding Azéma martingales by merely making  $n$  tend to infinity.

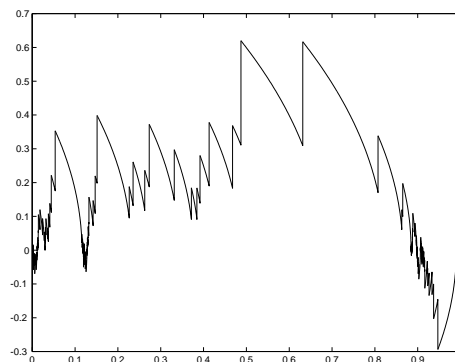
This description allows us to draw some pictures of the process  $Z_t^{(\beta, n)}$  for large values of  $n$  and different values of the parameter  $\beta$ , having therefore some good approximations of trajectories of Azéma martingales. However, in this paper, we do not intend to quantify any error term.

We finish this paper with some simulations of Azéma martingales, which were obtained by writing a code in MATLAB. It basically draws  $Z_t^{(\beta, 1)}$ , asks for the number of arches one wants to appear in the graph, and rescale the obtained function in order to get a function defined on  $[0, 1]$ . In the figures below, we asked for 1000 arches to appear.

Figure 1 represents Parthasarathy's martingale, whose graph is included inside the parabola of equation  $|y| = \sqrt{x}$ . Figure 2 represents the first Azéma martingale. We can see on Figure 4 that, for very small values of the parameter  $\beta$ , the Azéma martingale is a perturbation of the Brownian motion. It was considered as a candidate in [3] to replace the Brownian motion in the Black-Scholes model. We chose  $\beta = -0.07$  which was the parameter formerly estimated by Dritschel and Protter to modelise the variations of some asset prices.

We finally couple figures 3 and 5, which correspond respectively to the parameters  $-0.5$  and  $1$ , and notice similarities in their graphs up to some symmetry and horizontal affinity. Parthasarathy conjectured, after noticing the duality between the two formal generators of these processes that, up to a time change, the first process was the time reversal of the other (the conjecture was formulated, of course, for general positive values of the parameter). This conjecture was actually proven in an unpublished part of Phan's thesis [11] by a nice and unusual argument involving polynomial test functions.

Fig. 1:  $\beta = -2$ Fig. 2:  $\beta = -1$ Fig. 3:  $\beta = -0.5$ Fig. 4:  $\beta = -0.07$

Fig. 5:  $\beta = 1$ 

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