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Genetic analysis of thoroughbred racing performance in Spain

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Summary — Variance components of Spanish thoroughbred racing performance are estimated by REML using 2 models. The performance characteristics studied are speed, earnings, and rank for each race. The first model considered (model 1) includes the following fixed effects: jockey; sex; weight; age; distance; race course; and race year. A single race effect is substituted for the last 3 effects of model 1 in the second model 2. Heritability and repeatability estimates are close to zero for speed while they range from 0.07 to 0.17 and from 0.19 to 0.26 respectively for rank and earnings.

genetic analysis / performance / thoroughbred / heritability / repeatability

INTRODUCTION

Traditionally, a horse's breeding value is estimated through a subjective combination of race performances such as ranking, earnings, and racing time.

The first step in any genetic improvement scheme is to estimate parameters (heritability, repeatability, and genetic and phenotypic correlation) of the studied performances. In order to calculate these parameters, we need to estimate the variance of the random factors included in a chosen model. The values obtained determine the choice of an optimal breeding program for genetic improvement.
The most used methods for estimating variance components include Henderson's methods I, II and III (Henderson, 1953), MIVQUE (Rao, 1971), maximum likelihood (Hartley and Rao, 1967), and restricted maximum likelihood (Patterson and Thompson, 1971). Brief descriptions of these methods and their properties are given by Schaeffer (1983), Henderson (1984), and Van Raden (1986). Restricted maximum likelihood (REML) has the most suitable statistical properties for most animal breeding studies, but until recently its use has been severely limited by its computational requirements.

The thoroughbred population in Spain has some interesting characteristics from a genetic standpoint. First, performance data are detailed and complete (cross-data for each race and each horse). Second, its environmental parameters are well known and easily controlled. Third, genetic evaluations have been regularly undertaken using a mixed model since 1986. Despite its small size, this population allows reliable comparisons of various models.

The objective of this study is to compare the variance component estimations obtained with 2 different models. The traits studied are racing speed, earnings, and rank. It also compares the different alternatives to genetic evaluation in the case of thoroughbreds and attempts to clarify certain aspects of the Hill-Cunningham polemic in *Nature* (Gaffney and Cunningham, 1988; Hill, 1988) in complement to Chico (1990).

**MATERIALS AND METHODS**

**Materials**

The racing records used in this study are the same ones used in a precedent study (Chico, 1990). Only Grand-Prix races from 1980 to 1989 are used. Grand-Prix are high-priced adult horse races where all participants carry approximately the same weight. Genealogical data is limited to 3 generations, and missing or irrelevant genealogical information is discarded; the relationship matrix totals 2,705 animals. The number of observations (number of races by number of participants) amounts to 2,596 pertaining to 776 race horses, which were progeny of 315 sires and 597 mares.

**Method**

The fixed effects included in the mixed model are (in parentheses: number of levels or classes): race (260); jockey (27); sex (2); weight (5); age (4); distance (14); race course (3); and race year (10). Jockeys that have run less than 14 times are grouped in a single class. The 3 variables studied are speed, earnings, and rank in each race. Speed is estimated based on the distance of the race, the winner's run-time, and the distance between horses at arrival (Chico, 1988).

Horses 'out of the money' (those that rank below the fifth position) are given artificial earnings equal to half the earnings of the one that ranked immediately ahead of it. Normal distribution of the earnings is then obtained by applying a logarithmic transformation (Langlois, 1975; Chico, 1990).

Ranking is not a normal variable either. To get closer to this requirement, ranking is replaced by a normal score which is the expected value of the rank of a standardized normal variable (Pearson and Hartley, 1972).

Two different methods for the estimation of variance components are used: an EM-type algorithm of Dempster *et al.* (1977) to estimate repeatability and a DF-algorithm to estimate the additive genetic variance and the permanent environment variance separately. Two models for fixed effects are applied:

Model 1 includes the effects of jockey, sex, weight, age, distance, race course, and race year. Model 2 uses a single race effect to replace distance, race course, and year effects.

This single race effect represents the interaction of race distance, race course and year, as well as the effect of the environment in the considered race. The second model should yield smaller estimates of residual variance because it better accounts for race environment.

The REML method developed by Patterson and Thompson (1971) associated with the EM-type algorithm of Dempster *et al.* (1977) is
used to estimate performance repeatability, only considering the horse effect.

The analytical model is defined as follows:

\[ y = Xb + Zu + e \]

where:
- \( y \) is the vector of observations for speed, earning, or rank;
- \( b \) is the vector of unknown fixed effects;
- \( u \) is the vector of random horses effects representing both genetic and permanent environmental effects associated with a horse;
- \( e \) is the vector of random residual effects;
- \( X \) is the known incidence matrix for fixed effects; and
- \( Z \) is the known incidence matrix for random effects.

Henderson's mixed-model equations (MME) were used to predict animals' value.

where if \( u = a + p \):

\[ a_1 = \frac{(1 - r)/r}{s^2_y} = \frac{s^2_y}{s^2_a + s^2_p}; \]
\[ r = \frac{s^2_y}{s^2_a + s^2_p} = \text{repeatability}; \]
\[ s^2_y = s^2_a + s^2_p + s^2_e; \]
\[ a = \text{genetic value}; \]
\[ p = \text{permanent effect}; \]
\[ s^2_a = \text{variance of genetic value}; \]
\[ s^2_p = \text{variance of error}; \]
\[ s^2_p = \text{variance of permanent effect}; \]
\[ s^2_y = \text{total variance}. \]

REML method with a derivative-free algorithm (DFREML)

The use of a derivative-free REML algorithm was suggested by Grasser et al. (1987) to estimate variance components for a single trait analyzed with an animal model. The basic principle is the maximization of the log likelihood of a series of contrasts of the data, without calculating the 1st and 2nd derivatives of this log likelihood.

The linear model of analysis, an individual animal model (IAM), is used to obtain an additive genetic value not only for animals with records but also for any parent included in the relationship matrix. It also gives a BLU estimate for the fixed effect included in the model (Meyer, 1988). The programs of Meyer (1988) are used for this algorithm.

The analytical model is now:

\[ y = Xb + Zu + Zp + e \]

where:
- \( y, X, b, e \) as before;
- \( Zu \) and \( Zp \) are known incidences matrices;
- \( u \) is the vector of random additive genetic value effects; and
- \( p \) is the vector of random permanent environment effect.

It was assumed that the random vectors have zero mean and that

\[ \text{var}(Y) = \begin{bmatrix} V & 0 & 0 & 0 \\ 0 & G & 0 & 0 \\ 0 & 0 & P & 0 \\ 0 & 0 & 0 & R \end{bmatrix} \]

where:
- \( V(Y) = ZGZ' + ZPZ' + R \);
- \( G \) and \( R \) are non-singular matrices;
- \( G = \text{var}(u) = A \ast s^2_u. \)
- \( A \) is the full relationship matrix with all animals from the pedigree;
- \( P = I \ast s^2_p; s^2_p \) is the permanent environment variance;
- \( R = I \ast s^2_e; \) and \( s^2_e \) is the error variance.

The MME (Henderson, 1973) are:

\[ \begin{bmatrix} X'X & X'Z & X'Z + a_1A^{-1} \\ Z'Z + a_1A^{-1} & Z'Z + a_2A^{-1} \end{bmatrix} \begin{bmatrix} b \\ u \end{bmatrix} = \begin{bmatrix} X'Y \\ Z'Y \end{bmatrix} \]

where:
- \( a_1 = (1 - r)/h^2; a_2 = (1 - r)/(r - h^2); \)
- \( h^2 = s^2_a/s^2_y = \text{heritability}. \)

RESULTS AND ANALYSIS

Table I shows the repeatability estimated by REML-EM and DFREML methods using
the 2 models proposed. The iterative REML method always converges for all variables in both algorithms.

This table clearly shows that repeatability estimates for all 3 traits are slightly higher in model 2 than in model 1. This indicates a better data correction due to considering race effects; in addition to race course, distance and year, it includes racing characteristics. The repeatability of speed is very low in all methods and models. The repeatability of ranking and logarithm of earnings reach the same levels in all 3 methods with a small advantage in the second model (similar results were obtained by Chico et al (1990)).

The difference between heritability and repeatability (ignoring dominance effects and epistasis) is the proportion of total variance attributable to permanent environmental effects. Permanent environmental effects on racing performance may include such factors as owner, trainer, early nutrition, and injury.

The speed heritability estimates obtained are close to zero. Model 1 values are comparable to those obtained by Chico (1990). This is also confirmed by the work of Langlois (1975, 1978) who obtains heritability values below 0.10 as Buttram et al (1988).

Heritability estimates are in the order of magnitude of 0.1 for the logarithm of earnings and slightly larger for rank, whatever model is used. These results confirm those previously obtained by Chico (1988, 1990) with smaller values and smaller standard error.

The estimates of residual variances for each trait by model 2 are smaller than those from model 1 indicating a better ability of model 2 to account for variation in these racing traits. The earnings and rank heritability estimates valuation in this study are also equivalent to those obtained by Field and Cunningham (1976), Langlois and Chico (1989), Katona and Distl (1988), Klemestad (1990), and Arnason et al (1988).

This analysis allows us to conclude that speed is not the proper criterion for selection program for Spanish thoroughbred population. Heritability and repeatability are close to zero. However, earnings and rank are similar traits because estimates of the genetic correlation between these 2 traits are close to 1 in both models and their heritability and repeatability values are similar. Ranking appears to be the most important thoroughbred performance evaluation criterion. This fully justifies Tavernier's proposition (1990, 1991).

**CONCLUSION**

Thoroughbred race performances are moderately heritable and repeatable in terms of earning and rank and are independent of chronometric measurement. It is therefore not surprising that chronometric performances show little improvement over time (Gaffney and Cunningham, 1988; Hill, 1988). Speed is not a good selection criterion for thoroughbreds in Spain. On the other hand, earning and rank are good criteria for genetic improvement programs thanks to heritability, repeatability as well as genetic and phenotypic correlations.
Model 2 shows a better ability to account for variation in racing traits thanks to smaller estimates of residual variances. Heritability and repeatability are also slightly higher in model 2. Model 2 would be preferable to apply in future work.

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