PRUNING GALTON-WATSON TREES AND TREE-VALUED MARKOV PROCESSES

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ABSTRACT. We present a new pruning procedure on discrete trees by adding marks on the nodes of trees. This procedure allows us to construct and study a tree-valued Markov process $\{\mathcal{G}(u)\}$ by pruning Galton-Watson trees and an analogous process $\{\mathcal{G}^*(u)\}$ by pruning a critical or subcritical Galton-Watson tree conditioned to be infinite. Under a mild condition on offspring distributions, we show that the process $\{\mathcal{G}(u)\}$ run until its ascension time has a representation in terms of $\{\mathcal{G}^*(u)\}$. A similar result was obtained by Aldous and Pitman (1998) in the special case of Poisson offspring distributions where they considered uniform pruning of Galton-Watson trees by adding marks on the edges of trees.

1. INTRODUCTION

Using percolation on the branches of a Galton-Watson tree, Aldous and Pitman constructed by time-reversal in [4] an inhomogeneous tree-valued Markov process that starts from the trivial tree consisting only of the root and ends at time 1 at the initial Galton-Watson tree. When the final Galton-Watson tree is infinite, they define the ascension time A as the first time where the tree becomes infinite. They also define another process by pruning at branches the tree conditioned on non-extinction and they show that, in the special case of Poisson offspring distribution, some connections exist between the first process up to the ascension time and the second process.

Using the same kind of ideas, continuum trees valued Markov processes are constructed in [2] and an analogous relation is exhibited between the process obtained by pruning the tree and the other one obtained by pruning the tree conditioned on non-extinction. However, in that continuous framework, such results hold under very general assumptions on the branching mechanism.

Using the ideas of the pruning procedure [2], we propose here to prune a Galton-Watson tree on the nodes instead of the branches so that the connections pointed out in [4] hold for any offspring distribution.

Let us first explain the pruning procedure. Given a probability distribution $p = \{p_n, n = 0, 1, \cdots\}$, let $\mathcal{G}_p$ be a Galton Watson tree with offspring distribution p. Let $0 < u < 1$ be a constant. Then, if ν is an inner node of $\mathcal{G}_p$ that has n offsprings, we cut it (and discard all the sub-trees attached at this node) with probability u^{n-1} independently of the other nodes. The resulting tree is still a Galton-Watson tree with offspring distribution $p^{(u)}$ defined by:

$$p_n^{(u)} = u^{n-1} p_n, \quad \text{for } n \geq 1 \tag{1.1}$$

and

$$p_0^{(u)} = 1 - \sum_{n=1}^{\infty} p_n^{(u)}. \tag{1.2}$$

This particular pruning is motivated by the following lemma whose proof is postponed to Section 5.

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Lemma 1.1. *Let p and q be two offspring distributions. Let $\mathcal{G}_p$ and $\mathcal{G}_q$ be the associated Galton-Watson trees and let $\#\mathcal{L}_p$ and $\#\mathcal{L}_q$ denote the number of leaves of $\mathcal{G}_p$ and $\mathcal{G}_q$, respectively. Then we have that*

$$\forall N \geq 1, \quad \mathbb{P}(\mathcal{G}_p \in \cdot | \#\mathcal{L}_p = N) = \mathbb{P}(\mathcal{G}_q \in \cdot | \#\mathcal{L}_q = N) \tag{1.3}$$

if and only if

$$\exists u > 0, \ \forall n \geq 1, \ q_n = u^{n-1} p_n.$$

This lemma can be viewed has the discrete analogue of Lemma 1.6 of [1] that explains the choice of the pruning parameters for the continuous case. In [3], a similar result for Poisson Galton-Watson trees was obtained when conditioning by the total number of vertex, which explains why Poisson-Galton-Watson trees play a key role in [4].

Using the pruning at nodes procedure, given a critical offspring distribution p, we construct tree-valued (inhomogeneous) Markov processes $(\mathcal{G}(u), 0 \leq u < \bar{u})$ (the meaning of $\bar{u}$ will be clear later) such that

- the process is non-decreasing,
- for every u, $\mathcal{G}(u)$ is a Galton-Watson tree with offspring distribution $p^{(u)}$,
- the tree is critical for $u = 1$, sub-critical for $u < 1$ and super-critical for $u > 1$.

Let us state the main properties that we prove for that process and let us compare them with the results of [4]. We write $(\mathcal{G}^{AP}(u))$ for the tree-valued Markov process defined in [4].

In Section 3, we compute the forward transition probabilities and the forward transition rates for that process and exhibit a martingale that will appear several times (see Corollary 3.5). For a tree $\mathbf{t}$, we set

$$M(u, \mathbf{t}) = \frac{(1 - \mu(u))\#\mathcal{L}(\mathbf{t})}{p_0^{(u)}} \tag{1.4}$$

where $\#\mathcal{L}(\mathbf{t})$ denotes the number of leaves of $\mathbf{t}$ and $\mu(u)$ if the mean of offsprings in $\mathcal{G}(u)$. Then, the process

$$\big(M(u, \mathcal{G}(u)), 0 \leq u < 1\big)$$

is a martingale with respect to the filtration generated by $\mathcal{G}$. In [4], the martingale that appears (Corollary 23) for Poisson-Galton-Watson trees is $(1 - \mu(u))\#\mathcal{G}^{AP}(u)$.

When the tree $\mathcal{G}(u)$ is super-critical, it may be infinite. We define the ascension time A by:

$$A = \inf\{u \in [0, \bar{u}), \#\mathcal{G}(u) = \infty\}$$

with the convention $\inf \emptyset = \bar{u}$. We can then compute the joint law of A and $\mathcal{G}_{A-}$ (i.e. the tree just before it becomes infinite), see Proposition 4.2: we set $F(u)$ the extinction probability of $\mathcal{G}(u)$ and we have for $u \in [0, \bar{u})$

$$\mathbb{P}(A \leq u) = 1 - F(u)$$
$$\mathbb{P}(\mathcal{G}(A-) = \mathbf{t} \mid A = u) = M(\hat{u}, \mathbf{t})\mathbb{P}(\mathcal{G}(\hat{u}) = \mathbf{t})$$

with $\hat{u} = uF(u)$. These results are quite similar with those of Lemma 22 of [4], in particular the ascension time for $\mathcal{G}$ and $\mathcal{G}^{AP}$ for Poisson offspring distribution have suprisingly the same law. They must also be compared to the continuous framework, Theorem 6.5, Theorem 6.7 of [2].

When we have $p_0^{(\bar{u})} = 0$, then the final tree $\mathcal{G}(\bar{u})$ is a.s. infinite and the ascension time A is strictly less than $\bar{u}$. In that case, we consider the tree $\mathcal{G}^*(1)$ which is distributed as the tree $\mathcal{G}(1)$ conditioned on non-extinction. From this tree, by the same pruning procedure, we construct a non-decreasing tree-valued process $(\mathcal{G}^*(u), 0 \leq u \leq 1)$. We then prove the following representation formula (Proposition 4.3):

$$(\mathcal{G}(u), \ 0 \leq u < A) \stackrel{d}{=} (\mathcal{G}^*(u\gamma), \ 0 \leq u < \bar{F}^{-1}(1 - \gamma)),$$

where γ is a r.v, uniformly distributed on $(0,1)$, independent of $\{\mathcal{G}^*(\alpha) : 0 \leq \alpha \leq 1\}$. This result must also be compared to a similar result in [4], Proposition 26:

$$(\mathcal{G}^{AP}(u),\ 0 \leq u < A) \stackrel{d}{=} \left(\mathcal{G}^{AP*}(u\gamma),\ 0 \leq u < \frac{-\log\gamma}{(1-\gamma)}\right),$$

or to its continuous analogue, Corollary 8.2 of [2].

Let us stress again that, although the results are very similar, those in [4] only hold for Poisson-Galton-Watson trees whereas the results presented here hold for any offspring distribution.

The paper is organized as follows. In the next section, we recall some notations for trees and define the pruning procedure at nodes. In Section 3, we define the processes $\mathcal{G}$ and $\mathcal{G}^*$ and in Section 4 we state and prove the main results of the paper. Finally, we prove Lemma 1.1 in Section 5 .

2. Trees and Pruning

2.1. **Notation for Trees.** We present the framework developed in [8] for trees, see also [7] or [4] for more notations and terminology. Introduce the set of labels

$$\mathcal{W} = \bigcup_{n=0}^{\infty} (\mathbb{N}^*)^n,$$

where $\mathbb{N}^* = \{1, 2, \cdots\}$ and by convention $(\mathbb{N}^*)^0 = \{\emptyset\}$.

An element of $\mathcal{W}$ is thus a sequence $w = (w^1, \cdots, w^n)$ of elements of $\mathbb{N}$, and we set $|w| = n$, so that $|w|$ represents the generation of w or the height of w. If $w = (w^1, \cdots, w^m)$ and $v = (v^1, \cdots, v^n)$ belong to $\mathcal{W}$, we write $wv = (w^1, \cdots, w^m, v^1, \cdots, v^n)$ for the concatenation of w and v. In particular $w\emptyset = \emptyset w = w$. The mapping $\pi : \mathcal{W} \setminus \{\emptyset\} \longrightarrow \mathcal{W}$ is defined by $\pi((w^1, \cdots, w^n)) = (w^1, \cdots, w^{n-1})$ if $n \geq 1$ and $\pi((w^1)) = \emptyset$, and we say that $\pi(w)$ is the father of w. We set $\pi^0(w) = w$ and $\pi^n(w) = \pi^{n-1}(\pi(w))$ for $1 \leq n \leq |w|$. In particular, $\pi^{|w|}(w) = \emptyset$.

A (finite or infinite) rooted ordered tree $\mathbf{t}$ is a subset of $\mathcal{W}$ such that

1. $\emptyset \in \mathbf{t}$.
2. $w \in \mathbf{t} \setminus \{\emptyset\} \Longrightarrow \pi(w) \in \mathbf{t}$.
3. For every $w \in \mathbf{t}$, there exists a finite integer $k_w\mathbf{t} \geq 0$ such that, for every $j \in \mathbb{N}$, $wj \in \mathbf{t}$ if and only if $0 \leq j \leq k_w\mathbf{t}$ ($k_w\mathbf{t}$ is the number of children of $w \in \mathbf{t}$).

Let $\mathbf{T}^\infty$ denote the set of all such trees $\mathbf{t}$. Given a tree $\mathbf{t}$, we call an element in the set $\mathbf{t} \subset \mathcal{W}$ a node of $\mathbf{t}$. Denote the height of a tree $\mathbf{t}$ by $|\mathbf{t}| := \max\{|\nu| : \nu \in \mathbf{t}\}$. For $h \geq 0$, there exists natural restriction map $r_h : \mathbf{T}^\infty \to \mathbf{T}^h$ such that $r_h\mathbf{t} = \{\nu \in \mathbf{t} : |\nu| \leq h\}$, where $\mathbf{T}^h := \{\mathbf{t} \in T^\infty : |\mathbf{t}| \leq h\}$ ($\mathbf{T}^0 = \{\{\emptyset\}\}$).

We denote by $\#\mathbf{t}$ the number of nodes of $\mathbf{t}$. Let

$$\mathbf{T} := \{\mathbf{t} \in \mathbf{T}^\infty : \#\mathbf{t} < \infty\}$$

be the set of all finite trees. Then $\mathbf{T} = \cup_{h=1}^{\infty} \mathbf{T}^h$.

We define the shifted subtree of $\mathbf{t}$ above ν by

$$T_\nu\mathbf{t} := \{w : \nu w \in \mathbf{t}\}.$$

For $n \geq 0$, let $\mathrm{gen}(n, \mathbf{t})$ be the nth generation of individuals in $\mathbf{t}$. That is

$$\mathrm{gen}(n, \mathbf{t}) := \{v \in \mathbf{t} : |\nu| = n\}.$$

We say that $w \in \mathbf{t}$ is a leaf of $\mathbf{t}$ if $k_w\mathbf{t} = 0$ and set

$$\mathcal{L}(\mathbf{t}) := \{w \in \mathbf{t} : k_w\mathbf{t} = 0\}.$$

So $\mathcal{L}(\mathbf{t})$ denotes the set of leaves of $\mathbf{t}$ and $\#\mathcal{L}(\mathbf{t})$ is the number of leaves of $\mathbf{t}$.

We say that $w \in \mathbf{t}$ is an inner node of $\mathbf{t}$ if it is not a leaf (i.e. $k_w\mathbf{t} > 0$) and we denote by $\mathbf{t}^i$ the set of inner nodes of $\mathbf{t}$ i.e.

$$\mathbf{t}^i = \mathbf{t} \setminus \mathcal{L}(\mathbf{t}).$$

Given a probability distribution $p = \{p_n,\ n = 0, 1, \cdots\}$ with $p_1 < 1$, following [4], call a random tree $\mathcal{G}_p$ a Galton-Watson tree with offspring distribution p if the number of children of $\emptyset$ has distribution p:

$$\mathbb{P}(k_\emptyset \mathcal{G}_p = n) = p_n, \quad \forall\, n \geq 0$$

and for each $h = 1, 2, \cdots$, conditionally given $r_h\mathcal{G} = \mathbf{t}^h \in \mathbf{T}^h$, for $\nu \in \mathrm{gen}(h, \mathbf{t}^h)$, $k_\nu \mathcal{G}_p$ are i.i.d. random variables distributed according to p. That means

$$\mathbb{P}(r_{h+1}\mathcal{G} = \mathbf{t} \mid r_h\mathcal{G} = r_h\approx) = \prod_{\nu \in r_h\mathbf{t}} p_{k_\nu \mathbf{t}}, \quad \mathbf{t} \in \mathbf{T}^{h+1},$$

where the product is over all nodes ν of $\mathbf{t}$ of height h. We have then

$$\mathbb{P}(\mathcal{G} = \mathbf{t}) = \prod_{\nu \in \mathbf{t}} p_{k_\nu \mathbf{t}}, \quad \mathbf{t} \in T, \tag{2.1}$$

where the product is over all nodes ν of $\mathbf{t}$.

2.2. **Pruning at Nodes.** Let $\mathcal{T}$ be a tree in $\mathbf{T}^\infty$. For $0 \leq u \leq 1$, a random tree $\mathcal{T}(u)$ is called a *node pruning* of $\mathcal{T}$ with parameter u if it is constructed as follows: conditionnaly given $\mathcal{T} = \mathbf{t}, \mathbf{t} \in T^\infty$, for $0 \leq u \leq 1$, we consider a family of independent random variables $(\xi_\nu^u, \nu \in \mathbf{t})$ such that

$$P(\xi_\nu^u = 1) = 1 - P(\xi_\nu^u = 0) = \begin{cases} u^{k_\nu \mathbf{t} - 1}, & \text{if } k_\nu \mathbf{t} \geq 1, \\ 1, & \text{if } k_\nu \mathbf{t} = 0, \end{cases}$$

and define

$$\mathcal{T}(u) := \{\emptyset\} \bigcup \left\{ \nu \in \mathbf{t} \setminus \{\emptyset\} : \prod_{n=1}^{|\nu|} \xi^u_{\pi^n(\nu)} = 1 \right\}. \tag{2.2}$$

This means that if a node ν belongs to $\mathcal{T}(u)$ and $\xi_\nu^u = 1$, then $\nu j, j = 0, 1, \cdots, k_\nu(\mathbf{t})$ all belong to $\mathcal{T}(u)$ and if $\xi_\nu^u = 0$, then all subsequent offsprings of ν will be removed with the subtrees attached to these nodes. Thus $\mathcal{T}(u)$ is a random tree, $\mathcal{T}(u) \subset \mathcal{T}$ and we have for every $\nu \in \mathcal{W}$,

$$\mathbb{P}(k_\nu \mathcal{T}(u) = n \mid \nu \in \mathcal{T}(u), \mathcal{T} = \mathbf{t}) = \begin{cases} u^{n-1} 1_{\{k_\nu \mathbf{t} = n\}}, & n \geq 1, \\ 1_{\{k_\nu \mathbf{t} = 0\}} + (1 - u^{k_\nu \mathbf{t} - 1}) 1_{\{k_\nu \mathbf{t} \geq 1\}}, & n = 0. \end{cases} \tag{2.3}$$

We also have that for $h \geq 1$ and $\mathbf{t} \in \mathbf{T}^\infty$,

$$\mathbb{P}(r_h \mathcal{T}(u) = r_h \mathbf{t} \mid \mathcal{T} = \mathbf{t}) = \mathbb{P}\left(\prod_{\nu \in n(h, \mathbf{t})} \xi_\nu^u = 1 \right) = u^{\sum_{\nu \in n(h,\mathbf{t})} (k_\nu \mathbf{t} - 1)}, \tag{2.4}$$

where $n(h, \mathbf{t}) := \{\nu \in \mathbf{t} : k_\nu \mathbf{t} \geq 1 \text{ and } |\nu| < h\}$.

If $\mathcal{T}$ is a Galton Watson tree, we have the following proposition.

Proposition 2.1. *If $\mathcal{T}$ is a Galton Watson tree with offspring distribution $\{p_n, n \geq 0\}$, then $\mathcal{T}(u)$ is also a Galton Watson tree with offspring distribution $\{p_n^{(u)}, n \geq 0\}$ defined by (1.1) and (1.2).*

Proof. By (2.3),

$$\mathbb{P}(k_\emptyset \mathcal{T}(u) = 0) = \mathbb{P}(\mathcal{T} = \{\emptyset\}) + \sum_{n=1}^\infty \left(1 - u^{n-1}\right) \mathbb{P}(k_\emptyset \mathcal{T} = n) = p_0 + \sum_{n=1}^\infty \left(1 - u^{n-1}\right) p_n,$$

which is equal to $p_0^{(u)}$. For $n \geq 1$,

$$\mathbb{P}(k_\emptyset \mathcal{T}(u) = n) = u^{n-1} \mathbb{P}(k_\emptyset \mathcal{T} = n) = u^{n-1} p_n.$$

The fact that $\{\xi_\nu^u\}$ are, conditionally on $\mathcal{T}$ independent random variables, gives that for each $h = 1, 2, \cdots$, conditionally given $r_h\mathcal{T}(u) = \mathbf{t}^h \in \mathbf{T}^h$, for $\nu \in \mathbf{t}^h$ with $|\nu| = h$, $k_\nu\mathcal{T}(u)$ are independent. Meanwhile, again by (2.3),

$$\mathbb{P}(k_\nu\mathcal{T}(u) = n \mid r_h\mathcal{T}(u) = \mathbf{t}^h) = \begin{cases} u^{n-1}\mathbb{P}(k_\nu\mathcal{T} = n) = p_n^{(u)}, & \text{if } n \geq 1, \\ \mathbb{P}(k_\nu\mathcal{T} = 0) + \sum_{k\geq 1}(1 - u^{k-1})P(k_\nu\mathcal{T} = k) = p_0^{(u)}, & \text{if } n = 0. \end{cases}$$

Then the desired result follows readily. □

3. A Tree-valued Markov Process

3.1. **A tree-valued process given the terminal tree.** Let $\mathcal{T}$ be a tree in $\mathbf{T}^\infty$. We want to construct a $\mathbf{T}^\infty$-valued stochastic process $\{\mathcal{T}(u) : 0 \leq u \leq 1\}$ such that

- $\mathcal{T}(1) = \mathcal{T}$,
- for every $0 \leq u_1 < u_2 \leq 1$, $\mathcal{T}(u_1)$ is a node pruning of $\mathcal{T}(u_2)$ with pruning parameter u_1/u_2.

Recall that $\mathcal{T}^i$ is the set of the inner nodes of $\mathcal{T}$. Let $(\xi_\nu, \nu \in \mathcal{T}^i)$ be a family of independent random variables such that, for every $\nu \in \mathcal{G}^i$,

$$\mathbb{P}(\xi_\nu \leq u) = u^{k_\nu\mathcal{T}-1}.$$

Then, for very $u \in [0, 1]$, we set

$$\mathcal{T}(u) = \left\{\nu \in \mathcal{T},\ \forall 1 \leq n \leq |\nu|,\ \xi_{\pi^n(\nu)} \leq u\right\}.$$

We call the process $(\mathcal{T}(u), 0 \leq u \leq 1)$ a pruned process associated with $\mathcal{T}$. Let us remark that, contrary to the process of [4], the tree $\mathcal{T}(0)$ may not be reduced to the root as the nodes with one offspring are never pruned. More precisely, we have

$$\mathcal{T}(0) = \left\{\{1\}^n,\ n \leq \sup\{k,\ \forall l < k,\ k_{\{1\}^l}\mathcal{G} = 1\}\right\}$$

with the convention $\sup\emptyset = 0$ and $\{1\}^0 = \emptyset$.

We deduce from Formula (2.4) the following proposition:

Proposition 3.1. *We have that*

$$\lim_{u\to 1}\mathcal{T}(u) = \mathcal{T}, \quad a.s., \tag{3.1}$$

where the limit means that for almost every ω in the basic probability space, for each h there exists a $u(h, \omega) < 1$ such that $r_h\mathcal{T}(u, \omega) = r_h\mathcal{T}(\omega)$ for all $u(h, \omega) < u \leq 1$.

3.2. **Pruning Galton Watson trees.** Let $p = \{p_n,\ n = 0, 1, \ldots\}$ be an offspring distribution. Let $\mathcal{G}$ be a Galton-Watson tree with offspring distribution p. Then we consider the process $(\mathcal{G}(u), 0 \leq u \leq 1)$ such that, conditionally on $\mathcal{G}$, the process is a pruned process associated with $\mathcal{G}$.

Then for each $u \in [0, 1]$, $\mathcal{G}(u)$ is a Galton Watson tree with offspring distribution $p^{(u)}$. Let $g(s)$ denote the generating function of p. Then the distribution of $\mathcal{G}(u)$ is determined by the following generating function

$$g_u(s) = 1 - g(u)/u + g(us)/u, \quad 0 < u \leq 1. \tag{3.2}$$

3.3. **Forward transition probabilities.** Let $\mathcal{L}(u)$ be the set of leaves of $\mathcal{G}(u)$. Fix α and β with $0 \leq \alpha \leq \beta \leq 1$. Let us define

$$p_{\alpha,\beta}(k) = \frac{(1-(\alpha/\beta)^{k-1})p_k^{(\beta)}}{p_0^{(\alpha)}} \quad \text{for } k \geq 1 \quad \text{and} \quad p_{\alpha,\beta}(0) = \frac{p_0^{(\beta)}}{p_0^{(\alpha)}}.$$

and let $(\mathcal{G}^\nu_{\alpha,\beta},\ \nu \in \mathcal{L}(\alpha))$ be, conditionally given $\mathcal{G}(\alpha)$, i.i.d. random trees with distribution

$$\mathbb{P}(\mathcal{G}^\nu_{\alpha,\beta} = \mathbf{t} \mid \mathcal{G}(\alpha)) = \mathbb{P}(\mathcal{G}(\beta) = \mathbf{t} \mid k_\emptyset \mathcal{G}(\beta) = k_\emptyset \mathbf{t}) p_{\alpha,\beta}(k_\emptyset \mathbf{t}). \tag{3.3}$$

Remark 3.2. Equation (3.3) means that $\mathcal{G}^\nu_{\alpha,\beta}$ is a modified Galton Watson tree, in which the size of the first generation has distribution $p_{\alpha,\beta}$, while these and all subsequent individuals have offspring distribution $p^{(\beta)}(\cdot)$.

Set

$$\hat{\mathcal{G}}(\beta) = \mathcal{G}(\alpha) \cup \bigcup_{\nu \in \mathcal{L}(\alpha)} \{\nu w : w \in \mathcal{G}^\nu_{\alpha,\beta}\}. \tag{3.4}$$

That is $\hat{\mathcal{G}}(\beta)$ is a random tree obtained by adding a modified Galton Watson tree $\mathcal{G}^\nu_{\alpha,\beta}$ on each leaf ν of $\mathcal{G}(\alpha)$. The following proposition, which implies the Markov property of $\{\mathcal{G}(u),\ 0 \leq u \leq 1\}$, describes the transition probabilities of that tree-valued process.

Proposition 3.3. *For every $\leq \alpha \leq \beta \leq 1$, $(\mathcal{G}(\alpha), \mathcal{G}(\beta)) \overset{d}{=} (\mathcal{G}(\alpha), \hat{\mathcal{G}}(\beta))$.*

Proof. Let $\alpha < \beta$ and let $\mathbf{s}$ and $\mathbf{t}$ be two trees such that $\mathbf{s}$ can be obtained from $\mathbf{t}$ by a pruning at nodes. Then, by definition of the pruning procedure, we have

$$\begin{aligned}
\mathbb{P}(\mathcal{G}(\alpha) = \mathbf{s}, \mathcal{G}(\beta) = \mathbf{t}) &= \prod_{\nu \in \mathbf{t}} p^{(\beta)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathbf{s}^i} \left(\frac{\alpha}{\beta}\right)^{k_\nu \mathbf{t} - 1} \prod_{\nu \in \mathcal{L}(\mathbf{s}) \setminus \mathcal{L}(\mathbf{t})} \left(1 - \left(\frac{\alpha}{\beta}\right)^{k_\nu \mathbf{t} - 1}\right) \\
&= \prod_{\nu \in \mathbf{s}^i} p^{(\alpha)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathbf{t} \setminus \mathbf{s}^i} p^{(\beta)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathcal{L}(\mathbf{s})} \left(1 - \left(\frac{\alpha}{\beta}\right)^{k_\nu \mathbf{t} - 1} \mathbf{1}_{k_\nu \mathbf{t} > 0}\right) \\
&= \prod_{\nu \in \mathbf{s}^i} p^{(\alpha)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathbf{t} \setminus \mathbf{s}} p^{(\beta)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathcal{L}(\mathbf{s})} p^{(\beta)}_{k \nu \mathbf{t}} \left(1 - \left(\frac{\alpha}{\beta}\right)^{k_\nu \mathbf{t} - 1} \mathbf{1}_{k_\nu \mathbf{t} > 0}\right) \\
&= \prod_{\nu \in \mathbf{s}} p^{(\alpha)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathbf{t} \setminus \mathbf{s}} p^{(\beta)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathcal{L}(\mathbf{s})} \frac{p_{k \nu \mathbf{t}^{(\beta)}}}{p^{(\alpha)}_{k \nu \mathbf{t}}} \left(1 - \left(\frac{\alpha}{\beta}\right)^{k_\nu \mathbf{t} - 1} \mathbf{1}_{k_\nu \mathbf{t} > 0}\right) \\
&= \prod_{\nu \in \mathbf{s}} p^{(\alpha)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathbf{t} \setminus \mathbf{s}} p^{(\beta)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathcal{L}(\mathbf{s})} p_{\alpha,\beta}(k_\nu \mathbf{t}).
\end{aligned}$$

The definition of $p_{\alpha,\beta}$ and Remark 3.2 readily imply

$$\mathbb{P}(\mathcal{G}(\alpha) = \mathbf{s}, \hat{\mathcal{G}}(\beta) = \mathbf{t}) = \prod_{\nu \in \mathbf{s}} p^{(\alpha)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathbf{t} \setminus \mathbf{s}} p^{(\beta)}_{k_\nu \mathbf{t}} \prod_{\nu \in \mathcal{L}(\mathbf{s})} p_{\alpha,\beta}(k_\nu \mathbf{t})$$

which ends the proof. □

Let $\#\mathcal{L}(u)$ denote the number of leaves of $\mathcal{G}(u)$. The latter proposition together with the description of $\hat{\mathcal{G}}$ readily imply

$$(\mathcal{G}(\alpha), \#\mathcal{L}(\beta)) \overset{d}{=} \left(\mathcal{G}(\alpha), \sum_{\nu \in \mathcal{L}(\alpha)} \#\mathcal{L}(\mathcal{G}^\nu_{\alpha,\beta})\right). \tag{3.5}$$

We can also describe the forward transition rates when the trees are finite. If $\mathbf{s}$ and $\mathbf{t}$ are two trees and if $\nu \in \mathcal{L}(\mathbf{s})$, we define the tree obtained by grafting $\mathbf{t}$ on ν by

$$\mathbf{r}(\mathbf{s},\nu;\mathbf{t}) := \mathbf{s} \cup \{\nu w, w \in \mathbf{t}\}.$$

Corollary 3.4. *Let $\mathbf{s} \in \mathbf{T}$, $\mathbf{t} \in \mathbf{T}^\infty$, $\mathbf{t} \neq \{\emptyset\}$ and let $\nu \in \mathcal{L}(\mathbf{s})$. Then the transition rate at time u from $\mathbf{s}$ to $\mathbf{r}(\mathbf{s},\nu;\mathbf{t})$ is given by*

$$q_u(\mathbf{s} \to \mathbf{r}(\mathbf{s},\nu;\mathbf{t})) := \frac{k_\emptyset \mathbf{t} - 1}{u} \frac{\mathbb{P}(\mathcal{G}(u) = \mathbf{t})}{p_0^{(u)}}, \tag{3.6}$$

and no other transitions are allowed.

Proof. By Proposition 3.3, we have

$$\begin{aligned}\mathbb{P}\left(\mathcal{G}(u) = \mathbf{s}, \mathcal{G}(u+du) = \mathbf{r}(\mathbf{s},\nu;\mathbf{t})\right) &= \mathbb{P}\left(\mathcal{G}(u) = \mathbf{s}, \hat{\mathcal{G}}(u+du) = \mathbf{r}(\mathbf{s},\nu;\mathbf{t})\right) \\ &= \mathbb{P}\left(\mathcal{G}(u) = \mathbf{s}\right) \mathbb{P}(\mathcal{G}^{\nu}_{u,u+du} = \mathbf{t}) \prod_{\tilde{\nu} \in \mathcal{L}(\mathbf{s}) \setminus \{\nu\}} \mathbb{P}(\mathcal{G}^{\tilde{\nu}}_{u,u+du} = \{\emptyset\}) \\ &= \mathbb{P}\left(\mathcal{G}(u) = \mathbf{s}\right) \mathbb{P}(\mathcal{G}^{\nu}_{u,u+du} = \mathbf{t}) p_{u,u+du}(0)^{\#\mathcal{L}(\mathbf{s})-1}.\end{aligned}$$

Using (3.3), we get

$$\begin{aligned}&\mathbb{P}\Big(\mathcal{G}(u+du) = \mathbf{r}(\mathbf{s},\nu;\mathbf{t}) \;\Big|\; \mathcal{G}(u) = \mathbf{s}\Big) \\ &\qquad = \mathbb{P}\left(\mathcal{G}(u+du) = \mathbf{t}\right) \frac{p_{u,u+du}(k_\emptyset \mathbf{t})}{p^{(u+du)}_{k_\emptyset \mathbf{t}}} p_{u,u+du}(0)^{\#\mathcal{L}(\mathbf{s})-1} \\ &\qquad = \mathbb{P}\left(\mathcal{G}(u+du) = \mathbf{t}\right) \frac{1}{p_0^{(u)}} \left(1 - \left(\frac{u}{u+du}\right)^{k_\emptyset \mathbf{t} - 1}\right) \left(\frac{p_0^{(u+du)}}{p_0^{(u)}}\right)^{\#\mathcal{L}(\mathbf{s})-1} \\ &\qquad \underset{du \to 0}{\sim} \mathbb{P}\left(\mathcal{G}(u) = \mathbf{t}\right) \frac{1}{p_0^{(u)}} (k_\emptyset \mathbf{t} - 1) \frac{du}{u}.\end{aligned}$$

This gives Formula (3.6). A similar computation gives that, if $\mathbf{t}$ is obtained by grafting two trees (or more) on the leaves of $\mathbf{s}$,

$$\mathbb{P}\left(\mathcal{G}(u+du) = \mathbf{t} \;\middle|\; \mathcal{G}(u) = \mathbf{s}\right) = o(du)$$

and it is clear by construction that, in all the other cases,

$$\mathbb{P}\left(\mathcal{G}(u+du) = \mathbf{t} \;\middle|\; \mathcal{G}(u) = \mathbf{s}\right) = 0.$$

□

Let us define

$$\mu(u) := \sum_{k=1}^{\infty} k p_k^{(u)} \tag{3.7}$$

if it exits the mean of $p^{(u)}$. We set $u_1 = \sup\{u \in [0,1],\ \mu(u) \leq 1\}$. Recall the definition of function M in (1.4).

Corollary 3.5. *The process*

$$(M(u, \mathcal{G}(u)), \quad 0 \leq u < u_1), \tag{3.8}$$

is a martingale with respect to the filtration generated by $\{\mathcal{G}(u),\ 0 \leq u < u_1\}$.

Proof. First, by the branching property of Galton Watson process, for each $n \geq 1$, $0 \leq u < u_1$ and $\ell \geq n$,

$$\mathbb{P}\left(\#\mathcal{L}(u) = \ell \,\middle|\, k_\emptyset \mathcal{G}(u) = n\right) = \mathbb{P}\left(\sum_{i=1}^{n} L_i = \ell\right),$$

where $L_1, L_2, \cdots$ are i.i.d. copies of $\#\mathcal{L}(u)$. This gives

$$\mathbb{E}\left[\#\mathcal{L}(u)\right] = p_0^{(u)} + \mathbb{E}\left[k_\emptyset \mathcal{G}(u)\right] \mathbb{E}\left[\#\mathcal{L}(u)\right]$$

which implies

$$\mathbb{E}\left[\#\mathcal{L}(u)\right] = \frac{p_0^{(u)}}{1-\mu(u)}. \tag{3.9}$$

A straightforward computation gives that the mean of the offspring distribution $p_{\alpha,\beta}$ is

$$\mu_{\alpha,\beta} := \frac{\mu(\beta) - \mu(\alpha)}{p_0^{(\alpha)}}.$$

By the same reasoning, (3.5) and (3.3) imply, for $0 \leq \alpha \leq \beta < u_1$,

$$\begin{aligned}\mathbb{E}\left[\#\mathcal{L}(\beta) \,\middle|\, \mathcal{G}(\alpha)\right] &= \#\mathcal{L}(\alpha)\mathbb{E}\left[\#\mathcal{L}(\mathcal{G}^{\nu}_{\alpha,\beta})\right] = \#\mathcal{L}(\alpha)\left(p_{\alpha,\beta}(0) + \mu_{\alpha,\beta}\mathbb{E}\left[\#\mathcal{L}(\beta)\right]\right) \\ &= \#\mathcal{L}(\alpha)\left(p_{\alpha,\beta}(0) + \mu_{\alpha,\beta}\frac{p_0^{(\beta)}}{1-\mu(\beta)}\right)\end{aligned}$$

by (3.9). Then the martingale property of (3.8) follows from a simple calculation. □

3.4. **Pruning a Galton Watson Tree Conditioned on Non-Extinction.** Let p be a critical or sub-critical offspring distribution with mean μ such that $p_0 < 1$. We define the size-biased probability distribution p^* of p by

$$p_k^* = \frac{kp_k}{\mu}, \quad k \geq 0.$$

Let $\mathcal{G}$ be a Galton Watson tree with offspring distribution p. For a tree $\mathbf{t}$, we denote by $Z_n\mathbf{t} = \#\text{gen}(n, \mathbf{t})$ the number of individuals in the nth generation of $\mathbf{t}$. We first recall a result in [6].

Proposition 3.6. *(Kesten* [6]*, Aldous and Pitman* [4]*)*

(i) The conditional distribution of $\mathcal{G}$ given $\{Z_n\mathcal{G} > 0\}$ converges, as n tends to $+\infty$, toward the law of a random family tree $\mathcal{G}^\infty$ specified by

$$\mathbb{P}(r_h\mathcal{G}^\infty = \mathbf{t}) = \mu^{-h}(Z_h\mathbf{t})\mathbb{P}(r_h\mathcal{G} = \mathbf{t}) \quad \forall\, \mathbf{t} \in \mathbf{T}^{(h)}, h \geq 0.$$

(ii) Almost surely $\mathcal{G}^\infty$ contains a unique infinite path ($\emptyset = V_0, V_1, V_2, \cdots$) such that $\pi(V_{h+1}) = V_h$ for every $h = 0, 1, 2, \cdots$.

(iii) The joint distribution of $(V_0, V_1, V_2, \cdots)$ and $\mathcal{G}^\infty$ is determined recursively as follows: for each $h = 0, 1, 2, \cdots$, given $(V_0, V_1, V_2, \cdots, V_h)$ and $r_h\mathcal{G}^\infty$, the numbers of children $(k_\nu\mathcal{G}^\infty, \nu \in gen(h, \mathcal{G}^\infty))$ are independent with distribution p for $\nu \neq V_h$, and with the size-biased distribution p^* for $\nu = V_h$; given also the numbers of children $k_\nu\mathcal{G}^\infty$ for $\nu \in gen(h, \mathcal{G}^\infty)$, the vertex V_{h+1} has uniform distribution on the set $\{(V_h, i), 1 \leq i \leq k_{V_h}\mathcal{G}^\infty\}$.

We say that $\mathcal{G}^\infty$ is the Galton Watson tree associated with p conditioned on non-extinction. We then define the process $(\mathcal{G}^*(u), 0 \leq u \leq 1)$ as a pruned process associated with $\mathcal{G}^\infty$. By Proposition 3.1, $\mathcal{G}^*(1-) = \mathcal{G}^*(1) = \mathcal{G}^\infty(1)$ almost surely. And since there exists a unique infinite path, we get that $\mathcal{G}^*(u)$ is finite almost surely for all $0 \leq u < 1$.

The distribution of $\mathcal{G}^*(u)$ for fixed u is given in the following proposition. Let us recall that $\mu(u)$ is the mean of $p^{(u)}$ defined in (3.7).

Proposition 3.7. *For each* $0 \le u < 1$,

$$\mathbb{P}(\mathcal{G}^*(u) = \mathbf{t}) = \left(\sum_{\nu \in \mathcal{L}(\mathbf{t})} \frac{1}{\mu(1)^{|\nu|+1}} \right) \frac{\mu(1) - \mu(u)}{p_0^{(u)}} \mathbb{P}(\mathcal{G}(u) = \mathbf{t}), \quad \mathbf{t} \in T. \tag{3.10}$$

Proof. We prove (3.10) inductively. First, note that

$$\mathbb{P}(k_\emptyset \mathcal{G}^\infty = n) = p_n^* = np_n/\mu(1).$$

Then

$$\mathbb{P}(\mathcal{G}^*(u) = \{\emptyset\}) = \sum_{n \ge 1} (1 - u^{n-1}) \mathbb{P}(k_\emptyset \mathcal{G}^\infty(1) = n) = (\mu(1) - \mu(u))/\mu(1).$$

Since $\mathbb{P}(\mathcal{G}(u) = \{\emptyset\}) = p_0^{(u)}$, (3.10) holds for $\mathbf{t} = \{\emptyset\}$.

On the other hand, by Proposition 3.6, we have

$$\mathbb{P}\left(T_\nu \mathcal{G}^\infty = \mathbf{t} \;\middle|\; \nu = V_{|\nu|}\right) = \mathbb{P}\left(\mathcal{G}^\infty = \mathbf{t}\right)$$

and

$$\mathbb{P}\left(T_\nu \mathcal{G}^\infty = \mathbf{t} \;\middle|\; \nu \ne V_{|\nu|}\right) = \mathbb{P}\left(\mathcal{G} = \mathbf{t}\right)$$

which gives

$$\mathbb{P}\left(T_\nu \mathcal{G}^*(u) = \mathbf{t} \;\middle|\; \nu \in \mathcal{G}^*(u), \nu = V_{|\nu|}\right) = \mathbb{P}\left(\mathcal{G}^*(u) = \mathbf{t}\right) \tag{3.11}$$

and

$$\mathbb{P}\left(T_\nu \mathcal{G}^*(u) = \mathbf{t} \;\middle|\; \nu \in \mathcal{G}^*(u), \nu \ne V_{|\nu|}\right) = \mathbb{P}\left(\mathcal{G}(u) = \mathbf{t}\right), \tag{3.12}$$

respectively. Meanwhile, since $\mathcal{G}(u)$ is Galton-Watson tree,

$$\mathbb{P}\left(T_\nu \mathcal{G}(u) = \mathbf{t} \;\middle|\; \nu \in \mathcal{G}(u)\right) = \mathbb{P}(\mathcal{G}(u) = \mathbf{t}),$$

which implies

$$\mathbb{P}(\mathcal{G}(u) = \mathbf{t}) = p_{k_\emptyset \mathbf{t}}^{(u)} \prod_{1 \le j \le k_\emptyset \mathbf{t}} \mathbb{P}(\mathcal{G}(u) = T_{(j)}\mathbf{t}) \tag{3.13}$$

For some $h \ge 0$, assume that (3.10) holds for all trees in $\mathbf{T}^h$. By (3.11) and (3.12), we have for $\mathbf{t} \in \mathbf{T}^{h+1} \backslash \mathbf{T}^h$,

$$\begin{aligned}
\mathbb{P}(\mathcal{G}^*(u) = \mathbf{t}) &= \mathbb{P}(k_\emptyset \mathcal{G}^*(u) = k_\emptyset \mathbf{t}) \sum_{i=1}^{k_\emptyset \mathbf{t}} \mathbb{P}(\forall 1 \le j \le k_\emptyset \mathbf{t},\ T_{(j)}\mathcal{G}^*(u) = T_{(j)}\mathbf{t} \;\big|\; V_1 = i) \mathbb{P}(V_1 = i) \\
&= u^{k_\emptyset \mathbf{t} - 1} p_{k_\emptyset \mathbf{t}}^* \frac{1}{k_\emptyset \mathbf{t}} \sum_{i=1}^{k_\emptyset \mathbf{t}} \left(\mathbb{P}(\mathcal{G}^*(u) = T_{(i)}\mathbf{t}) \cdot \prod_{j \ne i, 1 \le j \le k_\emptyset \mathbf{t}} \mathbb{P}\left(\mathcal{G}(u) = T_{(j)}\mathbf{t}\right) \right) \\
&= \frac{p_{k_\emptyset \mathbf{t}}^{(u)}}{\mu(1)} \sum_{i=1}^{k_\emptyset \mathbf{t}} \sum_{\nu \in \mathcal{L}(T_{(i)}\mathbf{t})} \left(\frac{1}{\mu(1)^{|\nu|}} \frac{\mu(1) - \mu(u)}{p_0^{(u)}} \cdot \prod_{1 \le j \le k_\emptyset \mathbf{t}} \mathbb{P}\left(\mathcal{G}(u) = T_{(j)}\mathbf{t}\right) \right) \\
&= \left(\sum_{\nu' \in \mathcal{L}(\mathbf{t})} \frac{1}{\mu(1)^{|\nu'|+1}} \right) \frac{\mu(1) - \mu(u)}{p_0^{(u)}} \mathbb{P}(\mathcal{G}(u) = \mathbf{t}),
\end{aligned}$$

where the last equality follows from (3.13) and the facts

$$\bigcup_{1 \le i \le k_\emptyset \mathbf{t}} \{i\nu : \nu \in \mathcal{L}(T_{(i)}\mathbf{t})\} = \mathcal{L}(\mathbf{t})$$

and $|i\nu| = |\nu| + 1$. Since $\mathbf{T} = \cup_{h=1}^\infty \mathbf{T}^h$, (3.10) follows inductively. □

Remark 3.8. If $\mathcal{G}$ is a critical Galton Watson tree ($\mu(1) = 1$) with $p_1 < 1$, Formula (3.10) becomes

$$\mathbb{P}(\mathcal{G}^*(u) = \mathbf{t}) = \frac{\#\mathcal{L}(\mathbf{t})(1-\mu(u))}{p_0^{(u)}}\mathbb{P}(\mathcal{G}(u) = \mathbf{t}) = M(u, \mathbf{t})\mathbb{P}(\mathcal{G}(u) = \mathbf{t}).$$

In other words, the law of $\mathcal{G}^*(u)$ is absolutely continuous with respect to the law of $\mathcal{G}(u)$ with density the martingale of Corollary 3.5.

4. The Ascension Process and Its Representation

In this section, we consider a critical offspring distribution p with $p_1 < 1$ and set

$$I = \left\{ u \geq 0,\ \sum_{k=1}^{+\infty} u^{k-1} p_k \leq 1 \right\}.$$

Let us remark that $I = [0, \bar{u}]$ or $I = [0, \bar{u})$ with $\bar{u} \geq 1$. Let us give some exemples:

Example 1: The binary case.
We consider the offspring distribution p defined by $p_0 = p_2 = 1/2$ (each individual dies out or gives birth to two children with equal probability). In that case, we have

$$\sum_{k=1}^{\infty} u^{k-1} p_k = \frac{1}{2} u$$

and hence we have $\bar{u} = 2$, $I = [0, 2]$ and $p_0^{(2)} = 0$.

Example 2: The geometric case.
We consider the offspring distribution p defined by

$$\begin{cases} p_k = \alpha\beta^{k-1} & \text{for } k \geq 1 \\ p_0 = 1 - \frac{\alpha}{1-\beta} \end{cases}$$

Then, for every u, $p^{(u)}$ is still of that form. As the offspring distribution p is critical, we must have $\alpha = (1-\beta)^2$, $0 < \beta < 1$. In that case, we have

$$\sum_{k=1}^{+\infty} p_k u^{k-1} = \frac{(1-\beta)^2}{1-\beta u}$$

and hence we have $\bar{u} = 2 - \beta$, $I = [0, \bar{u}]$, $p_0^{\bar{u}} = 0$.

Example 3: We suppose now that the offspring distribution is

$$\begin{cases} p_k = \dfrac{1}{k2^k} & \text{for } k \geq 1 \\ p_0 = 1 - \displaystyle\sum_{k\geq 1} \frac{1}{k2^k} \end{cases}$$

In that case, $\bar{u} = 2$ and $I = [0, 2)$.

For $u \in I$, let us define

$$\begin{cases} p_k^{(u)} = u^{k-1} p_k, \quad k \geq 1, \\ p_0^{(u)} = 1 - \displaystyle\sum_{k=1}^{\infty} p_k^{(u)}. \end{cases} \tag{4.1}$$

Then, for $u \in I$, $p^{(u)}$ is still an offspring distribution, it is sub-critical for $u < 1$, critical for $u = 1$ and super-critical for $u > 1$.

We construct a tree-valued process $(\mathcal{G}(u),\ u \in I)$ such that

- for every $u \in I$, $\mathcal{G}(u)$ is a Galton-Watson tree with offspring distribution $p^{(u)}$,
- for every $\alpha, \beta \in I$, $\alpha < \beta$, $\mathcal{G}(\alpha)$ is a pruning of $\mathcal{G}(\beta)$,

by the following method:

1st case: $I = [0, \bar{u}]$.
The process $(\mathcal{G}(t\bar{u}),\ t \in [0,1])$ is a pruned process associated with $\mathcal{G}(\bar{u})$.

2nd case: $I = [0, \bar{u})$.
We can define for every $v < \bar{u}$ a process $(\mathcal{G}(u),\ 0 \le u \le v)$ by the first case method. The distributions of these processes satisfy a straightforward compatible condition as v varies. Hence there exists a projective limit that gives the distribution of the process $(\mathcal{G}(u),\ 0 \le u < \bar{u})$.

We now consider $\{\mathcal{G}(u),\ u \in I\}$ as an *ascension process* with the *ascension time*

$$A := \inf\{u \in I,\ \#\mathcal{G}(u) = \infty\}$$

with the convention $\inf \emptyset = \bar{u}$.

The state in the ascension process at time u is $\mathcal{G}(u)$ if $0 \le u < A$ and $\mathbf{t}(\infty)$ if $A \le u$ where $\mathbf{t}(\infty)$ is a state representing any infinite tree. Then the ascension process is still a Markov process with countable state-space $\mathbf{T} \cup \mathbf{t}(\infty)$, where $\mathbf{t}(\infty)$ is an absorbing state.

Denote by $F(u)$ the extinction probability of a Galton Watson process with offspring distribution $p^{(u)}$, which is the least non-negative root of the following equation with respect to s

$$s = g_u(s) = 1 - g_1(u)/u + g_1(us)/u \tag{4.2}$$

. where g_1 is the generating function associated with the offspring distribution p.

We set

$$\bar{F}(u) = 1 - F(u) \tag{4.3}$$

Thus

$$u\bar{F}(u) + g_1(u - u\bar{F}(u)) = g_1(u). \tag{4.4}$$

The distribution of the ascension process is determined by the transition rates (3.6) and

$$q_u(\mathbf{s} \to \mathbf{t}(\infty)) = \frac{\#\mathcal{L}(\mathbf{s})}{up_0^{(u)}} \sum_{k=2}^{\infty} (k-1) p_k^{(u)} (1 - F(u)^k). \tag{4.5}$$

Define the conjugate $\hat{u}$ by

$$\hat{u} = uF(u) \qquad \text{for } u \in I. \tag{4.6}$$

We can restate Equation (4.4) into

$$g_1(\hat{u}) - g_1(u) = \hat{u} - u. \tag{4.7}$$

We first prove the following result which is already well-known, see for instance [5], p52. We just restate this property in terms of our pruning parameter.

Proposition 4.1. *For any $u \in I$, $u \ge 1$*

$$\mathbb{P}(\mathcal{G}(u) = \mathbf{t}) = F(u)\mathbb{P}(\mathcal{G}(\hat{u}) = \mathbf{t}), \quad \mathbf{t} \in \mathbf{T}. \tag{4.8}$$

Proof. By (2.1), we have

$$\mathbb{P}(\mathcal{G}(u) = \mathbf{t}) = \prod_{\nu \in \mathbf{t} \setminus \mathcal{L}(\mathbf{t})} p_{k_\nu \mathbf{t}}^{(u)} \cdot \prod_{\nu \in \mathcal{L}(\mathbf{t})} p_0^{(u)}$$

and by (4.6),

$$\mathbb{P}(\mathcal{G}(u) = \mathbf{t}) = F(u)^{-(\sum_{\nu \in \mathbf{t} \setminus \mathcal{L}(\mathbf{t})} (k_\nu \mathbf{t} - 1))} \left(\frac{p_0^{(u)}}{p_0^{(\hat{u})}} \right)^{\#\mathcal{L}(\mathbf{t})} \mathbb{P}(\mathcal{G}(\hat{u}) = \mathbf{t}).$$

We also have

$$p_0^{(\hat{u})} = 1 - \sum_{k=1}^{\infty} F(u)^{k-1} p_k^{(u)} = 1 + p_0^{(u)}/F(u) - g_u(F(u))/F(u) = p_0^{(u)}/F(u). \tag{4.9}$$

Then the desired result follows from the fact that given a tree $\mathbf{t} \in \mathbf{T}$,

$$\#\mathcal{L}(\mathbf{t}) = 1 + \sum_{\nu \in \mathbf{t} \setminus \mathcal{L}(\mathbf{t})} (k_\nu \mathbf{t} - 1).$$

□

In what follows, we will often suppose that

$$I = [0, \bar{u}] \quad \text{and} \quad p_0^{(\bar{u})} = 0, \tag{4.10}$$

which is equivalent to the condition

$$\sum_{k=1}^{+\infty} \bar{u}^{k-1} p_k = 1$$

and which implies that $\mathcal{G}(\bar{u})$ is infinite and $A < \bar{u}$ a.s. We can however give the law of A and of the tree $\mathcal{G}(A-)$ just before the ascension time in general, this is the purpose of the next proposition.

Proposition 4.2. *For $u \in [1, \bar{u})$ and $\mathbf{t} \in \mathbf{T}$,*

$$\mathbb{P}(A \leq u) = \bar{F}(u). \tag{4.11}$$

$$\mathbb{P}(\mathcal{G}(A-) = \mathbf{t} | A = u) = M(\hat{u}, \mathbf{t}) \mathbb{P}(\mathcal{G}(\hat{u}) = \mathbf{t}). \tag{4.12}$$

$$\mathbb{P}(\#\mathcal{G}(A-) < +\infty \mid A = u) = 1. \tag{4.13}$$

Furthermore, under assumption (4.10),

$$\left(A, \frac{\hat{A}}{A}\right) = (A, F(A)) \stackrel{d}{=} \left(\bar{F}^{-1}(1-\gamma), \gamma\right), \tag{4.14}$$

where $\bar{F}^{-1} : [0,1] \to [1, \bar{u}]$ is the inverse function of $\bar{F}$ and γ is a r.v uniformly distributed on $(0, 1)$.

Proof. We have $\mathbb{P}(A \leq u) = \mathbb{P}(\#\mathcal{G}(u) = \infty) = \bar{F}(u)$ which gives (4.11).
This gives

$$\mathbb{P}(A \in du) = -F'(u) du$$

and derivating (4.2) gives

$$u\big(1 - g_1'(\hat{u})\big) F'(u) = 1 - g'(u) - F(u)\big(1 - g'(\hat{u})\big). \tag{4.15}$$

By (4.5), we have for $\mathbf{t} \in \mathbf{T}$

$$\mathbb{P}(\mathcal{G}(A-) = \mathbf{t}, A \in du) = \frac{\#\mathcal{L}(\mathbf{t}) \mathbb{P}(\mathcal{G}(u) = \mathbf{t})}{u p_0^{(u)}} \sum_{k=2}^{\infty} (k-1) p_k^{(u)} (1 - F(u)^k) du. \tag{4.16}$$

Now, using (4.8), we have

$$\mathbb{P}(\mathcal{G}(A-) = \mathbf{t}, A \in du) = \frac{\#\mathcal{L}(\mathbf{t}) \mathbb{P}(\mathcal{G}(\hat{u}) = \mathbf{t}) F(u)}{u p_0^{(u)}} \sum_{k=2}^{\infty} (k-1) p_k^{(u)} (1 - F(u)^k) du.$$

Easy computations give

$$\sum_{k=2}^{\infty}(k-1)p_k^{(u)}(1-F(u)^k) = g_1'(u) - \frac{g_1(u)}{u} - F(u)g_1'(\hat{u}) + \frac{g(\hat{u})}{u}$$
$$= -F'(u)u\big(1-g_1'(\hat{u})\big) + 1 - F(u) + \frac{g_1(\hat{u})-g_1(u)}{u}$$
$$= -F'(u)u\big(1-g_1'(\hat{u})\big)$$

using first Equation (4.15) and then Equation (4.7).

This finally gives

$$\mathbb{P}(\mathcal{G}(A-)=\mathbf{t}, A\in du) - \frac{\#\mathcal{L}(\mathbf{t})\mathbb{P}(\mathcal{G}(\hat{u})=\mathbf{t})F(u)}{p_0^{(u)}}F'(u)\big(1-g_1'(\hat{u})\big)du$$
$$= -\frac{\#\mathcal{L}(\mathbf{t})\mathbb{P}(\mathcal{G}(\hat{u})=\mathbf{t})}{p_0^{(\hat{u})}}F'(u)\big(1-\mu(\hat{u})\big)du$$

by Equation (4.9), which yields (4.12).

Summing (4.12) over all finite trees $\mathbf{t}$ gives

$$\mathbb{P}(\#\mathcal{G}(A-)<+\infty \mid A=u) = \mathbb{E}\left[M(\hat{u},\mathcal{G}(\hat{u}))\right] = 1$$

by the martingale property, which is (4.13).

Finally, under assumption (4.10), (4.11) gives

$$A \stackrel{d}{=} \bar{F}^{-1}(\gamma) \stackrel{d}{=} \bar{F}^{-1}(1-\gamma).$$

Thus $(A,\bar{F}(A)) \stackrel{d}{=} (\bar{F}^{-1}(1-\gamma), 1-\gamma)$. So we have

$$(A,F(A)) = (A, 1-\bar{F}(A)) \stackrel{d}{=} (\bar{F}^{-1}(1-\gamma),\gamma).$$

which is just (4.14). □

With Remark 3.8 and Proposition 4.2 in hand, we have the following representation of the ascension process $\{\mathcal{G}(\alpha) : 0\le\alpha<A\}$ under assumption (4.10).

Proposition 4.3. *Under assumption (4.10),*

$$\{\mathcal{G}(u),\ 0\le u<A\} \stackrel{d}{=} \{\mathcal{G}^*(u\gamma) : 0\le u<\bar{F}^{-1}(1-\gamma)\}, \tag{4.17}$$

where γ is a r.v with uniform distribution on $(0,1)$, independent of $\{\mathcal{G}^*(u) : 0\le u\le 1\}$.

Proof. Let $\{\mathcal{G}^*(u) : 0\le u\le 1\}$ be independent of A. Then by Remark 3.8,

$$\mathbb{P}(\mathcal{G}^*(\hat{A})=\mathbf{t} \mid A=a) = \mathbb{P}(\mathcal{G}^*(\hat{a})=\mathbf{t}) = \frac{\#\mathcal{L}(\mathbf{t})(1-\mu(\hat{a}))}{p_0^{(\hat{a})}}\mathbb{P}(\mathcal{G}(\hat{a})=\mathbf{t}).$$

Thus it follows from (4.12) that $(A,\mathcal{G}(A-)) = (A,\mathcal{G}^*(\hat{A}))$. On the other hand, by the definition of node-pruning, for every $\mathbf{t}\in\mathbf{T}$,

$$\mathbb{P}\big((\mathcal{G}(s\beta), 0\le s\le 1)\in\cdot \mid \mathcal{G}(\beta)=\mathbf{t}\big) = \mathbb{P}\big((\mathcal{G}^*(s\alpha), 0\le s\le 1)\in\cdot \mid \mathcal{G}^*(\alpha)=\mathbf{t}\big).$$

Thus conditioning on the terminal value implies

$$\{\mathcal{G}(u),\ 0\le u<A\} \stackrel{d}{=} \{\mathcal{G}^*(\hat{A}u/A) : 0\le u<A\}.$$

Then (4.17) follows from (4.14). □

Example 4.4. (Binary case)
If $p = \frac{1}{2}\delta_2 + \frac{1}{2}\delta_0$, then $\bar{u} = 2$ and $\mathcal{G}(u)$ is a Galton Watson tree with binary offspring distribution $\frac{u}{2}\delta_2 + (1 - \frac{u}{2})\delta_0$ for $0 \leq u \leq 2$. In this case, we have

$$F(u) = \frac{2}{u} - 1$$

and the ascension time A is distributed as

$$\bar{F}^{-1}(1-\gamma) = \frac{2}{1+\gamma}$$

where γ is a uniform random variable on $(0, 1)$. Its density is given by

$$f(t) = -F'(t) = \frac{2}{t^2}1_{[1,2]}(t).$$

Example 4.5. (Geometric case)
We suppose that the offspring distribution p is of the form

$$p_k = (1-\beta)^2\beta^{k-1} \text{ for } k \geq 1, \qquad p_0 = \beta.$$

In that case, we have

$$\begin{cases} p_k^{(u)} = (1-\beta)^2(u\beta)^{k-1} & \text{for } k \geq 1, \\ p_0^{(u)} = 1 - \frac{(1-\beta)^2}{1-u\beta}, \end{cases}$$

$\bar{u} = 2 - \beta$, and assumption (4.10) is statisfied.

We then get

$$F(u) = \frac{2-u-\beta}{1-u\beta}\frac{1}{u}$$

and the ascension time A has density

$$\left(\frac{2-\beta}{u^2} + \frac{(1-\beta)^2\beta}{(1-u\beta)^2}\right)\mathbf{1}_{[1,2-\beta]}(u).$$

5. Proof of Lemma 1.1

Let $\mathcal{G}_p$ be a Galton-Watson tree with offspring distribution p such that $p_1 < 1$.

If $\mathbf{t}$ is a tree, we denote by $(a_1, a_2, \dots, a_m)$ the numbers of offsprings of the inner nodes. Its number of leaves is then

$$a_1 + \cdots + a_m - m + 1.$$

If $\mathbf{t}$ is a tree with n leaves, we have

$$\mathbb{P}(\mathcal{G}_p = \mathbf{t}) = p_{a_1} \cdots p_{a_m} p_0^n$$

and therefore

$$\mathbb{P}(\#\mathcal{L}_p = n) = C_p(n)p_0^n$$

with

$$C_p(n) = \sum_{\mathbf{t},\ \#\mathcal{L}(\mathbf{t})=n} p_{a_1} \cdots p_{a_m}.$$

Then we have, for every n such that $C_p(n) \neq 0$,

$$\mathbb{P}(\mathcal{G}_p = \mathbf{t} | \#\mathcal{L}_p = n) = \mathbb{P}(\mathcal{G}_q = \mathbf{t} | \#\mathcal{L}_q = n) \iff \frac{p_{a_1} \cdots p_{a_m}}{C_p(n)} = \frac{q_{a_1} \cdots q_{a_m}}{C_q(m)}. \tag{5.1}$$

First, let us suppose that

$$\mathbb{P}(\mathcal{G}_p = \mathbf{t} | \#\mathcal{L}_p = n) = \mathbb{P}(\mathcal{G}_q = \mathbf{t} | \#\mathcal{L}_q = n).$$

For $n = 1$, all the trees with one leaf are those with one offspring at each generation until the last individual dies. Therefore,

$$\begin{aligned}\mathbb{P}(\mathcal{G}_p = \mathbf{t} | \#\mathcal{L}_p = 1) = \mathbb{P}(\mathcal{G}_q = \mathbf{t} | \#\mathcal{L}_q = 1) &\iff \forall k \geq 0,\ p_1^k(1-p_1) = q_1^k(1-q_1)\\ &\iff p_1 = q_1.\end{aligned}$$

We set $n_0 = \inf\{n \geq 2, p_n > 0\}$. We then set $u = (q_{n_0}/p_{n_0})^{1/(n_0-1)}$.

If the only nonzero terms of p are p_0, p_1 and p_{n_0}, the relation

$$q_n = u^{n-1} p_n$$

is trivally true for every $n \geq 1$.

In the other cases, let $n > n_0$ such that $p_n > 0$ and let N be the integer defined by:

$$N = 2(n-1)(n_0-1).$$

Let us consider first a tree $\mathbf{t}$ that has $N+1$ leaves, $n-1$ inner nodes with n_0 offsprings and $n_0 - 1$ inner nodes with n offsprings. Applying (5.1) to that tree gives

$$\frac{p_{n_0}^{n-1} p_n^{n_0-1}}{C_p(N+1)} = \frac{q_{n_0}^{n-1} q_n^{n_0-1}}{C_q(N+1)}.$$

Then, let us consider another tree with $N+1$ leaves composed of $2(n-1)$ inner nodes with n_0 offsprings. For that new tree, (5.1) gives

$$\frac{p_{n_0}^{2(n-1)}}{C_p(N+1)} = \frac{q_{n_0}^{2(n-1)}}{C_q(N+1)}.$$

Dividing the two latter equations gives

$$q_n = u^{n-1} p_n.$$

It remains to remark that this identity also holds when $n = n_0$ and when $p_n = 0$.

Conversely, let us suppose that $q_n = u^{n-1} p_n$ for every $n \geq 1$. Let n such that $C_p(n) \neq 0$. Then, for every $\mathbf{t}$ with n leaves, we have

$$\begin{aligned} q_{a_1} \dots q_{a_m} &= u^{a_1 - 1} p_{a_1} \dots u^{a_m - 1} p_{a_m}\\ &= u^{a_1 + \cdots a_m - m} p_{a_1} \dots p_{a_m}\\ &= u^{n-1} p_{a_1} \dots p_{a_m}.\end{aligned}$$

We then have $C_q(n) = u^{n-1} C_p(n)$ and

$$\frac{q_{a_1} \dots q_{a_m}}{C_q(n)} = \frac{u^{n-1} p_{a_1} \dots p_{a_m}}{u^{n-1} C_p(n)} = \frac{p_{a_1} \dots p_{a_m}}{C_p(n)},$$

that is

$$P(\mathcal{G}_p = \mathbf{t} | \#\mathcal{L}_p = n) = P(\mathcal{G}_q = \mathbf{t} | \#\mathcal{L}_q = n).$$

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